# Prediction from Partial Information and Hindsight, an Alternative Proof 

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#### Abstract

Let $X$ be a random variable distributed over $n$-bit strings with $H(X) \geq n-k$, where $k \ll n$. Using subadditivity we know that the average coordinate has high entropy. Meir and Wigderson [1] showed that a random coordinate looks random to an adversary who is allowed to query around $n / k$ other coordinates non-deterministically. They used this result to obtain top-down arguments in depth-3 circuit lower bounds. In this note we give an alternative proof of their main result which tightens their parameters. Our proof is inspired by a paper of Paturi, Pudlák and Zane [3] who gave a non-trivial $k$-SAT algorithm and tight depth-3 circuit lower bounds for parity.


## 1 Introduction

Motivated by developing top-down arguments in circuit complexity, Meir and Wigderson [1] recently studied the following question: let $X=\left(X_{1}, \ldots, X_{n}\right) \in\{0,1\}^{n}$ be a random variable with entropy at least $n-k$. An adversary who knows the distribution of $X$ and a uniformly chosen coordinate $i \in[n]$ needs to predict the value of $X_{i}$. He is allowed to query $q$ coordinates of $X$ other than $i$. How large should $q$ be so that the adversary has non-negligible advantage? The answer is $\omega(n / k)$ and it holds even if the adversary can make his queries non-deterministically. This is formally captured as follows.

Definition 1. A witness for a coordinate $i \in[n]$ is a pair $(Q, a)$ where $Q \subseteq[n] \backslash\{i\}$ and $a \in\{0,1\}^{|Q|}$. The witness appears in a string $x \in\{0,1\}^{n}$ if $\left.x\right|_{Q}=a$. The length of the witness is $|Q|$.

Definition 2. A $q$-family of witnesses $\mathcal{F}$ for a coordinate $i \in[n]$ is a set of witnesses for $i$ of length at most $q$. We say that a string $X \in\{0,1\}^{n}$ satisfies $\mathcal{F}$, denoted by $X \models \mathcal{F}$, if at least one of the witnesses in $\mathcal{F}$ appears in $X$. For a random string $X \in\{0,1\}^{n}$, a bit $b \in\{0,1\}$ and $0 \leq \epsilon \leq 1$, we say that $\mathcal{F} \epsilon$-predicts $X_{i}=b$ if

$$
\operatorname{Pr}\left[X_{i}=b \mid X \text { satisfies } \mathcal{F}\right] \geq \frac{1}{2}+\frac{\epsilon}{2} .
$$

Theorem 1 ([1]). Let $X$ be a random variable over $\{0,1\}^{n}$ such that $H(X) \geq n-k$, and $q \leq n$. For $\epsilon \in(0,1]$, every $i \in[n]$ and $b \in\{0,1\}$ let $\mathcal{F}_{i}^{b}$ be a $q$-family of witnesses that $\epsilon$-predicts $X_{i}=b$ and let $\sigma_{i}$ be the probability that either $X \models \mathcal{F}_{i}^{0}$ or $X \models \mathcal{F}_{i}^{1}$. Then $\bar{\sigma}:=\mathbb{E}_{i}\left[\sigma_{i}\right]$ is at most $\frac{300 \cdot k \cdot q}{\epsilon^{3} \cdot n}$.

[^0]The application of this theorem in circuit lower bounds uses $\epsilon=1$ (see [1]). The focus of this note is an improvement of Theorem 1 for this case and the background which gave rise to the proof, interestingly coming from similar questions in circuit lower bounds.

Definition 3. Given a random string $X \in\{0,1\}^{n}$, a coordinate $i \in[n]$ and a bit $b \in\{0,1\}$, a $b$-certificate for $i$ is a witness $(Q, a)$ such that

$$
\operatorname{Pr}\left[X_{i}=b|X|_{Q}=a\right]=1
$$

Theorem 2. Let $X$ be a random variable over $\{0,1\}^{n}$ such that $H(X) \geq n-k$, and $q \leq n$. For every $i \in[n]$ and $b \in\{0,1\}$ let $\sigma_{i}$ be the probability that $X$ contains any certificate for $X_{i}$. Then $\bar{\sigma}:=\mathbb{E}_{i}\left[\sigma_{i}\right]$ is at most $\frac{k \cdot(q+1)}{n}$.

It is easy to compare this theorem with Theorem 1 for case $\epsilon=1$. Note that the bound in Theorem 2 is tight. For any $\sigma \in[0,1]$ we can consider a random variable $X=\left(X_{1}, X_{2}, \ldots, X_{n}\right)$ defined by the following process: group the first $\sigma n$ coordinates in blocks of size $q+1$. Then for every block assign all coordinates except the last one uniformly at random and set the last one to be the xor of the previous ones. Finally assign all remaining coordinates $X_{\sigma n+1}, X_{\sigma n+2}, \ldots, X_{n}$ uniformly at random. It is easy to see that

$$
H(X)=\frac{\sigma n}{q+1} \cdot q+(1-\sigma) \cdot n=n-\frac{\sigma n}{q+1} .
$$

Applying Theorem 2 with $k=\sigma n /(q+1)$ we get $\bar{\sigma} \leq \sigma$. On the other hand each of the first $\sigma n$ coordinates has certificate (with probability 1 ) of size $q$ that predicts it, hence $\bar{\sigma} \geq \sigma$.

## 2 Background for the proof

Let $\psi$ be a $k$-CNF formula in $n$ variables. We say that a satisfying assignment $\alpha$ is isolated if flipping any single bit of $\alpha$ falsifies some clause. Paturi, Pudlák and Zane [3] showed that the number of isolated solutions is upper bounded by $2^{(1-1 / k) n}$. This immediately implies a $2^{n / k-1}$ lower bound for parity against $\Sigma_{k}^{3}$ circuits (depth-3 OR-AND-OR circuits with bottom fan-in bounded by $k$ ). Furthermore since they proved this via an efficient encoding, they also managed to obtain an algorithm for $k$-SAT. Following a similar line of reasoning Paturi, Pudlák, Saks and Zane [2] showed that if every pair of solutions of a $k$-CNF disagree on a super-constant number of bits, the number of satisfying assignments is upper bounded by $2^{\left(1-\pi^{2} / 6 k+o(1)\right) n}$. This subsequently gave an improved lower bound for error-correcting codes with proper parameters and an improved $k$-SAT algorithm which still remains the best known. The best $\Sigma_{k}^{3}$ lower bounds are $2^{\Omega(n / k)}$ and the best upper bound bound for $k$-SAT is $2^{(1-\Omega(1 / k)) n}$. Improving the dependence on $k$ remains a major challenge in complexity theory. One way to make progress in this direction would be to make the arguments of [3] and [2] more flexible through replacing the efficient encoding by direct entropy arguments. This brings us to the result of Meir and Wigderson.

Let $X$ be a uniformly chosen isolated solution of $\psi$. We can generate a ( $k-1$ )-family of certificates $\mathcal{F}$ for $X$ as follows. For every isolated solution $\alpha$ and every $i \in[n]$, flip the $i$ th bit. This falsifies some clause $C_{\alpha, i}$ which means that under $\alpha, C_{\alpha, i}$ is satisfied only by $x_{i}$. We then add the set of variables other than $x_{i}$ in $C_{\alpha, i}$ and their values under $\alpha$ to $\mathcal{F}_{i}$. By the construction of $\mathcal{F}$ for every $i \in[n]$ we have that $\operatorname{Pr}\left[X \models \mathcal{F}_{i}\right]=1$ and thus $\mathbb{E}_{i}\left[\sigma_{i}\right]=1$. Theorem 1 then implies
that $H(X) \leq(1-1 / 300 k) n$, since otherwise we would have $\mathbb{E}_{i}\left[\sigma_{i}\right]<1$. However [3] gives the tight bound $H(X) \leq(1-1 / k) n$. This suggests that it might be possible to improve Theorem 1 and this is what we achieved.

Our proof is inspired by the argument of [3] and it bounds the entropy by considering random permutations of the bits, whereas the proof of [1] considers random splittings of $X$ in two parts. Unfortunately our improvement does not imply any new result. However we would like to stress our original motivation which led to the current proof and pose the question of reproving or improving the circuit lower bound of [2] as follows.

Question 1. Let $\psi$ be a $k$-CNF whose satisfying assignments are pairwise of super-constant Hamming distance. Let $X$ be uniformly distributed over the set of satisfying assignments. Can we upper bound $H(X)$ directly without adhering to the explicit encoding of [2]?

## 3 Proof of Theorem 2

Let $X=\left(X_{1}, \ldots, X_{n}\right)$ be a random variable over $\{0,1\}^{n}$. For $i \in[n]$ we will use $X_{<i}$ as a shorthand for $\left(X_{1}, \ldots, X_{i-1}\right)$ and let $\mathcal{F}_{i}$ be the set of all certificates for $X_{i}$.

Using the chain rule one can derive the following equality

$$
H(X)=H\left(X_{1}\right)+H\left(X_{2} \mid X_{1}\right)+\cdots+H\left(X_{n} \mid X_{<n}\right)
$$

The same equality holds if we permute the coordinates of $X$ :

$$
H(X)=H\left(X^{\pi}\right)=H\left(X_{1}^{\pi}\right)+H\left(X_{2}^{\pi} \mid X_{1}^{\pi}\right)+\cdots+H\left(X_{n}^{\pi} \mid X_{<n}^{\pi}\right) .
$$

where $\pi \in S_{n}$ is a permutation of $[n]$ and $X^{\pi}$ is a random variable equal to $X$ permuted with $\pi$, i.e. $X^{\pi}=\left(X_{1}^{\pi}, X_{2}^{\pi}, \ldots, X_{n}^{\pi}\right)=\left(X_{\pi(1)}, X_{\pi(2)}, \ldots, X_{\pi(n)}\right)$.

Lemma 1. For any $i \in[n]$ and $\alpha \in\{0,1\}^{n}$ such that $\alpha$ contains a certificate $(Q, a)$ for $X_{i}$ of size at most $q$,

$$
\underset{\pi}{\mathbb{E}}\left[H\left(X_{i} \mid X_{<\pi^{-1}(i)}^{\pi}=\alpha_{<\pi^{-1}(i)}^{\pi}\right)\right] \leq 1-\frac{1}{q+1} .
$$

Proof. If for a fixed permutation $\pi$ all indices in $Q$ precede $\pi^{-1}(i)$, i.e., $\forall j \in Q, \pi^{-1}(j)<\pi^{-1}(i)$, then $H\left(X_{i} \mid X_{<\pi^{-1}(i)}^{\pi}=\alpha_{<\pi^{-1}(i)}^{\pi}\right)=H\left(X_{i}|X|_{Q}=a\right)=0$ by the definition of certificate. We have

$$
\begin{aligned}
\underset{\pi}{\mathbb{E}}\left[H\left(X_{i} \mid X_{<\pi^{-1}(i)}^{\pi}=\alpha_{<\pi^{-1}(i)}^{\pi}\right)\right] & \leq \underset{\pi}{\operatorname{Pr}}\left[\forall j \in Q, \pi^{-1}(j)<\pi^{-1}(i)\right] \cdot H\left(X_{i}|X|_{Q}=a\right) \\
& +\underset{\pi}{\operatorname{Pr}}\left[\exists j \in Q, \pi^{-1}(j)>\pi^{-1}(i)\right] \cdot H\left(X_{i}\right) \\
& =\left(1-\frac{1}{|Q|+1}\right) \cdot H\left(X_{i}\right) \\
& \leq 1-\frac{1}{|Q|+1} \leq 1-\frac{1}{q+1} .
\end{aligned}
$$

Hence we get the statement of the lemma.

Let us bound the expectation of $H(X)$ over all permutations of $[n]$.

$$
\begin{aligned}
\underset{\pi}{\mathbb{E}}\left[H\left(X^{\pi}\right)\right]= & \underset{\pi}{\mathbb{E}}\left[\sum_{i=1}^{n} H\left(X_{i}^{\pi} \mid X_{<i}^{\pi}\right)\right] \\
= & \underset{\pi}{\mathbb{E}}\left[\sum_{i=1}^{n} \sum_{\beta \in\{0,1\}^{i-1}} \operatorname{Pr}\left[X_{<i}^{\pi}=\beta\right] \cdot H\left(X_{i}^{\pi} \mid X_{<i}^{\pi}=\beta\right)\right] \quad \text { (chain rule) } \\
= & \underset{\pi}{\mathbb{E}}\left[\sum_{i=1}^{n} \sum_{\beta \in\{0,1\}^{i-1}} \sum_{\substack{\alpha \in\{0,1\}^{n} \\
\alpha_{<i}=\beta}} \operatorname{Pr}[X=\alpha] \cdot H\left(X_{i}^{\pi} \mid X_{<i}^{\pi}=\beta\right)\right] \quad \text { (expansion) } \\
= & \underset{\pi}{\mathbb{E}}\left[\sum_{i=1}^{n} \sum_{\alpha \in\{0,1\}^{n}} \operatorname{Pr}[X=\alpha] \cdot H\left(X_{i}^{\pi} \mid X_{<i}^{\pi}=\alpha_{<i}^{\pi}\right)\right] \quad \text { (expansion) } \\
= & \underset{\pi}{\mathbb{E}}\left[\sum_{j=1}^{n} \sum_{\alpha \in\{0,1\}^{n}} \operatorname{Pr}[X=\alpha] \cdot H\left(X_{j} \mid X_{<\pi^{-1}(j)}^{\pi}=\alpha_{<\pi^{-1}(j)}^{\pi}\right)\right] \quad(j=\pi(i)) \\
= & \underset{\pi}{\mathbb{E}}\left[\sum_{j=1}^{n} \sum_{\substack{\alpha \in\{0,1\}^{n} \\
\alpha \not \mathcal{F}_{j}}} \operatorname{Pr}[X=\alpha] \cdot H\left(X_{j} \mid X_{<\pi^{-1}(j)}^{\pi}=\alpha_{<\pi^{-1}(j)}^{\pi}\right)\right. \\
& \left.+\sum_{j=1}^{n} \sum_{\substack{\alpha \in\{0,1\}^{n} \\
\alpha \neq \mathcal{F}_{j}}} \operatorname{Pr}[X=\alpha] \cdot H\left(X_{j} \mid X_{<\pi^{-1}(j)}^{\pi}=\alpha_{<\pi^{-1}(j)}^{\pi}\right)\right]
\end{aligned}
$$

By the linearity of expectation, we have

$$
\begin{aligned}
\underset{\pi}{\mathbb{E}}\left[H\left(X^{\pi}\right)\right] & =\sum_{j=1}^{n} \sum_{\substack{\alpha \in\{0,1\}^{n} \\
\alpha \neq \mathcal{F}_{j}}} \operatorname{Pr}[X=\alpha] \cdot \underset{\pi}{\mathbb{E}}\left[H\left(X_{j} \mid X_{<\pi^{-1}(j)}^{\pi}=\alpha_{<\pi^{-1}(j)}^{\pi}\right)\right] \\
& +\sum_{j=1}^{n} \sum_{\substack{\alpha \in\{0,1\}^{n} \\
\alpha \nLeftarrow \mathcal{F}_{j}}} \operatorname{Pr}[X=\alpha] \cdot \underset{\pi}{\mathbb{E}}\left[H\left(X_{j} \mid X_{<\pi^{-1}(j)}^{\pi}=\alpha_{<\pi^{-1}(j)}^{\pi}\right)\right]
\end{aligned}
$$

Applying Lemma 1 we get

$$
\begin{aligned}
\underset{\pi}{\mathbb{E}}\left[H\left(X^{\pi}\right)\right] & \leq \sum_{j=1}^{n} \sum_{\substack{\alpha \in\{0,1\}^{n} \\
\alpha \models \mathcal{F}_{j}}} \operatorname{Pr}[X=\alpha] \cdot\left(1-\frac{1}{q+1}\right)+\sum_{j=1}^{n} \sum_{\substack{\alpha \in\{0,1\}^{n} \\
\alpha \not \models \mathcal{F}_{j}}} \operatorname{Pr}[X=\alpha] \cdot 1 \\
& =\sum_{j=1}^{n}\left(\sigma_{j} \cdot\left(1-\frac{1}{q+1}\right)+\left(1-\sigma_{j}\right) \cdot 1\right)=\sum_{j=1}^{n}\left(1-\frac{\sigma_{i}}{q+1}\right) \\
& =n-\frac{n \cdot \bar{\sigma}}{q+1} .
\end{aligned}
$$

Having in mind that $\mathbb{E}_{\pi}\left[H\left(X^{\pi}\right)\right]=\mathbb{E}_{\pi}[H(X)]=H(X) \geq n-k$ we get $\bar{\sigma} \leq \frac{k \cdot(q+1)}{n}$.

## References

[1] O. Meir and A. Wigderson. Prediction from partial information and hindsight, with application to circuit lower bounds. Electronic Colloquium on Computational Complexity (ECCC), 24:149, 2017.
[2] R. Paturi, P. Pudlák, M. E. Saks, and F. Zane. An improved exponential-time algorithm for k-SAT. J. ACM, 52(3):337-364, 2005.
[3] R. Paturi, P. Pudlák, and F. Zane. Satisfiability coding lemma. Chicago J. Theor. Comput. Sci., 1999, 1999.


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