

# Almost Euclidean sections of the N-dimensional cross-polytope using $O(N)$ random bits

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## Abstract

It is well known that  $\mathbb{R}^N$  has subspaces of dimension proportional to  $N$  on which the  $\ell_1$  norm is equivalent to the  $\ell_2$  norm; however, no explicit constructions are known. Extending earlier work by Artstein–Avidan and Milman, we prove that such a subspace can be generated using  $O(N)$  random bits.

## 1 Introduction

We study embeddings of  $\ell_2$  spaces into  $\ell_1$  spaces. Recall that the  $\ell_p$  norm on  $\mathbb{R}^N$  is defined by:

$$\|x\|_p = \left( \sum_{i=1}^N |x_i|^p \right)^{1/p} \quad (p \geq 1)$$

The following inequality holds on  $\mathbb{R}^N$ :

$$\|x\|_2 \leq \|x\|_1 \leq \sqrt{N} \|x\|_2$$

It is well known since the work of Figiel, Lindenstrauss and Milman [7] and Kashin [13] that there exists a subspace  $E$  of  $\mathbb{R}^N$  of dimension

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$\Theta(N)$  such that for all  $x \in E$ ,  $\|x\|_1 = \Theta(\sqrt{N}\|x\|_2)$  (for the convenience of the reader, we recall the  $\Theta$ -notation at the end of the introduction).

More formally put, for every  $0 < \eta < 1$  and every  $N \in \mathbb{N}$  (large enough), there exists an  $\eta N$ -dimensional subspace  $E \hookrightarrow \mathbb{R}^N$  such that for every  $x \in E$ :

$$c_\eta \sqrt{N} \|x\|_2 \leq \|x\|_1 \leq \sqrt{N} \|x\|_2 \quad (1)$$

where  $c_\eta > 0$  depends only on  $\eta$ .

The subspace  $E$  gives in particular an embedding of  $(\mathbb{R}^{\eta N}, \|\cdot\|_2)$  into  $(\mathbb{R}^N, \|\cdot\|_1)$ . This allows to reduce various problems in  $\ell_2$  norm to corresponding problem in  $\ell_1$  norm, with only a constant blowup in the dimension.

An explicit construction of  $E$  would therefore have various algorithmic applications. This was put forward by Indyk [10, 11], who proved several related results and applied them to problems in Computer Science.

No explicit subspace  $E$  satisfying (1) has been found so far (for large  $N$ ). However, it is known that a randomly chosen subspace, under various natural definitions of distributions of subspaces, satisfies (1) with probability very close to 1.

In a sense, this situation is typical for various problems in asymptotic convex geometry, as for numerous properties satisfied by “random” high-dimensional objects it is hard to generate a deterministic object satisfying the property.

To resolve this dissonance, a new line of research was introduced by Sh. Artstein-Avidan and V. Milman. In the innovating work [3], the authors proposed to reduce the randomness needed to generate the random objects. More precisely, they showed that the random constructions in the proofs of a broad range of theorems, from Milman’s Quotient of Subspace theorem to Zig-Zag approximation, can be performed on the finite probability space  $\{-1, +1\}^R$  equipped with the uniform probability measure, where  $R \in \mathbb{N}$  is reasonably small (the reader may refer to the work [4] by Artstein–Avidan and Milman for further developments and to the ICM lecture by Szarek [16] for a discussion of these and related issues).

In this case, we say informally that  $R$  random bits are used in the construction. For example, regarding the property (1), Artstein-Avidan and Milman showed that  $O(N \log N)$  random bits suffice to construct the subspace  $E$ .

Their proof uses  $\varepsilon$ -net arguments, and decreasing the number of random bits beyond  $\Omega(N)$  will probably require entirely new proof ideas. However, the  $\log N$  factor in [3] seemed to be an artefact of the proof.

In this work, we show that this is indeed the case, and reduce the number of random bits to  $O(N)$  using a modification of the construction from [3].

**Theorem 1.** *For every  $0 < \eta < 1$ , an  $\eta N$ -dimensional subspace of  $\mathbb{R}^N$  satisfying (1) can be generated using  $O(N)$  random bits. Moreover, the memory needed to generate the subspace is  $O(\log^2 N)$ .*

As promised, we recall now the  $\Theta$ -notation:

*Notation.* Let  $f, g$  be two functions from  $(a, +\infty)$  or  $(a, +\infty) \cap \mathbb{N}$  to  $\mathbb{R}_+$ . We will write:

1.  $f = O(g)$  if there exist two constants  $C > 0$  and  $x_0 \geq a$  such that  $f(x) \leq Cg(x)$  for every  $x \geq x_0$ ;
2.  $f = o(g)$  if  $f(x)/g(x) \rightarrow 0$  as  $x \rightarrow \infty$ ;
3.  $f = \Omega(g)$  if  $g = O(f)$ ;
4.  $f = \omega(g)$  if  $g = o(f)$ ;
5. and finally,  $f = \Theta(g)$  if  $f = O(g)$  and  $f = \Omega(g)$ .

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## 2 Construction

Denote  $\xi = 1 - \eta$ ,  $n = \xi N$ . We will construct a random  $n \times N$  sign matrix  $A$  (that is,  $A_{ij} = \pm 1$ ) using  $O(N)$  random bits, and then prove that the kernel

$$E = \text{Ker}A = \{x \in \mathbb{R}^N \mid Ax = 0\}$$

satisfies (1) with high probability.

Recall the following simple definition:

**Definition 1.** The Hadamard (or entrywise) product of two  $n \times N$  matrices  $A_1$  and  $A_2$  is the  $n \times N$  matrix  $A = A_1 \bullet A_2$ , defined by  $(A)_{i,j} = (A_1)_{i,j}(A_2)_{i,j}$ .

Our random matrix  $A$  will be the Hadamard product  $A_1 \bullet A_2$  of two random matrices  $A_1$  and  $A_2$ , independent of each other. The construction of  $A_1$  and  $A_2$  will use two different techniques, both of them quite common.

**Definition 2.** A sequence of random variables  $X_1, \dots, X_M$  is called  $k$ -wise independent if every  $k$  of them are independent.

It is well-known that it is possible to construct  $M$   $k$ -wise independent random signs from  $O(k \log M)$  truly independent random signs. More formally, we have:

**Lemma A.** For every  $k \leq M$ , there exists a subset

$$\Upsilon_{k,M} \subset \{-1, 1\}^M$$

such that  $|\Upsilon_{k,M}| = 2^{C_{k,M}}$ ,  $C_{k,M} = O(k \log M)$ , and for the randomly chosen vector  $X = (X_1, \dots, X_M)$  from  $\Upsilon_{k,M}$ , the following properties hold:

1. For  $1 \leq m \leq M$ ,  $\mathbb{P}\{X_m = -1\} = \mathbb{P}\{X_m = 1\} = 1/2$ .
2. The coordinates of  $X$  are  $k$ -wise independent.
3. The set  $\Upsilon_{k,M}$  is *explicit*, meaning that there exists a bijection  $v_{k,M} : \{-1, 1\}^{C_{k,M}} \rightarrow \Upsilon_{k,M}$  that can be computed in time polynomial in  $k$  and  $M$ .

**Definition 3.** The random variables  $(X_1, \dots, X_M)$  satisfying the conditions 1.-2. of Lemma A are called  $k$ -wise independent random signs.

For completeness, we reproduce a proof of Lemma A due to Alon, Babai and Itai [1] in Appendix A.

The elements of our first matrix  $A_1$  will be  $k$ -wise independent with  $k = \Theta(\log N)$ . That is,  $A_1$ , regarded as a vector in  $\{-1, 1\}^{nN}$ , will be a uniformly chosen element of  $\Upsilon_{k,nN}$ .

*Remark.* Regardless of the distribution of the random sign matrix  $A_2$ , the entries  $A_{ij}$  of the Hadamard product  $A = A_1 \bullet A_2$  are  $k$ -wise independent random signs (in the sense of Definition 3).

Recall the definition of  $\ell_2$  operator norm:

**Definition 4.** For a matrix  $A$ , we define its operator norm as

$$\|A\| = \max_{x \neq 0} \frac{\|Ax\|_2}{\|x\|_2} .$$

The  $k$ -wise independence of the elements of  $A_1$  allows to control the operator norm of  $A$ . The following technical lemma may be of independent interest:

**Lemma 2.** *Let  $V$  be any  $n \times N$  matrix of  $2k$ -wise independent random signs,  $k \leq c_2 \sqrt{N}$  (where  $c_2 > 0$  is a numerical constant). Denote  $\xi = n/N \leq 1$ . Then, for  $t \geq 0$ ,*

$$\begin{aligned} \mathbb{P} \left\{ \frac{1}{\sqrt{N}} \|V\| \geq 1 + \sqrt{\xi} + t \right\} &\leq 2n \left( 1 + \frac{t}{1 + \sqrt{\xi}} \right)^{-2k} \\ &\leq 2n \exp \left\{ \frac{-2kt}{1 + \sqrt{\xi} + t} \right\} . \end{aligned}$$

We prove the lemma in Section 3.

**Corollary 3.** *Let  $0 < \xi < 1$ ,  $n = \xi N$ ; let  $A_1$  be constructed as above with  $k$ -wise independent entries, and let  $A = A_1 \bullet A_2$ , where  $A_2$  is an arbitrary random sign matrix independent of  $A_1$ . There exists a numerical constant  $C_1 > 0$  such that for  $k \geq C_1 \log n$ ,*

$$\mathbb{P}[\|A\| > 3\sqrt{N}] < 1/n .$$

We now head to construct a probability space for  $A_2$ ; we use random walks on expander graphs (see Hoory, Linial and Wigderson [9] for an extensive survey). Let us recall the basic definitions.

Let  $G = (\mathcal{V}, \mathcal{E})$  be a  $d$ -regular graph; the value of  $d$  plays no significant role in the estimates, so the reader may assume  $d = 4$ . Let  $P^G$  be the transition matrix of the random walk of  $G$ :

$$P_{uv}^G = \begin{cases} 1/d, & (u, v) \in \mathcal{E} \\ 0, & (u, v) \notin \mathcal{E}. \end{cases}$$

Denote by  $1 = \lambda_1 \geq \lambda_2 \geq \lambda_3 \geq \dots$  the eigenvalues of  $P^G$  arranged in decreasing order, and denote  $\lambda = \max_{i \geq 2} |\lambda_i|$ .

In this notation, the graph  $G$  is called a  $(|\mathcal{V}|, d, \lambda)$ -graph. We will only need the following fact (cf. [9], [3]):

*Fact.* For any  $d \geq 3$  and any number of vertices  $|\mathcal{V}|$  (big enough), there exists a  $(|\mathcal{V}|, d, \lambda)$ -graph  $G = (\mathcal{V} = \{1, 2, \dots, |\mathcal{V}|\}, \mathcal{E})$  such that

1.  $\lambda < 0.95$  and
2.  $G$  is *explicit*, formally meaning that set of neighbours

$$\{u \in \mathcal{V} \mid (u, v) \in \mathcal{E}\}$$

of any vertex  $v \in \mathcal{V}$  can be computed in time that is polynomial in  $\log |\mathcal{V}|$ .

Sometimes we will call such a graph an expander graph with parameter  $\lambda$ .

Let  $G = (\mathcal{V}, \mathcal{E})$  be an expander graph, with vertices  $\mathcal{V}$  indexed by the elements of  $\Upsilon_{4,N}$ . Let  $v_1, v_2, \dots, v_n$  be a random walk of length  $n$  in  $G$ , starting from a random element of  $\mathcal{V}$ . Write the sign vectors corresponding to  $v_1, \dots, v_n$  in  $\Upsilon_{4,N}$  as the rows of  $A_2$ .

The use of expander graphs is similar to [3]; however, we use constant degree expanders. We also show it suffices to use 4-wise independent rows rather than truly independent rows. This enables the computation to be performed using less memory ( $O(\log^2 N)$ ).

Note that the construction uses in total

$$\begin{aligned} O(\log n \log(Nn)) + O(\log N) + O(n \log d) \\ = O(n + \log n \log N) = O(N) \end{aligned} \quad (2)$$

random bits. Also, we have the following:

**Lemma 4.** *Let  $A_1$  be any constant sign matrix, and let  $A_2$  be constructed as above. For every  $x \in \mathbb{R}^N$  and any  $\varepsilon \leq c_\lambda \sqrt{\xi}$ ,*

$$\mathbb{P} \left\{ \|Ax\|_2 < 6\varepsilon \sqrt{N} \|x\|_2 \right\} < C_\lambda p_\lambda^n,$$

where the constants  $C_\lambda, c_\lambda > 0$  and  $0 < p_\lambda < 1$  depend on the parameter  $\lambda \in [0, 1)$  of the graph  $G$ .

**Corollary 5.** *The statement of the lemma remains true if we change  $A_1$  from constant to drawn from any distribution.*

We prove this lemma in Section 4; the proof is a variation on the ideas from Artstein-Avidan and Milman [3].

Now we can reformulate our main result.

**Theorem 6.** *Let  $A_1$  and  $A_2$  be constructed as above ( $A_1$  has  $\Theta(\log n)$  independent entries, the rows of  $A_2$  come from a random walk on an expander); let  $A = A_1 \bullet A_2$ ,  $E = \text{Ker } A$ . Then, with probability  $1 - o(1)$ ,*

$$\frac{c'\xi}{\sqrt{\log 1/\xi}} \sqrt{N} \|x\|_2 \leq \|x\|_1 \leq \sqrt{N} \|x\|_2 \quad \text{for every } x \in E, \quad (3)$$

where  $c' > 0$  is a universal constant.

The proof uses the Lemmata formulated above as well as the following standard lemma from asymptotic convex geometry.

**Lemma B.** Let  $A$  be a random  $n \times N$  sign matrix such that:

1.  $\mathbb{P}[\|A\| > 3\sqrt{N}] \leq q$ ;
2. There exist  $0 < p < 1$ ,  $\varepsilon > 0$  and  $C > 0$  such that for every  $y \in \mathbb{R}^N$ ,

$$\mathbb{P} \left\{ \|Ay\|_2 < 6\varepsilon\sqrt{N}\|y\|_2 \right\} < Cp^n .$$

Then with probability at least

$$1 - q - p^{\Theta(n)}$$

over the choice of  $A$ , we have:

$$\|x\|_1 \geq \delta\sqrt{N}\|x\|_2 \quad \text{for every } x \in \text{Ker}A ,$$

where we can take

$$\delta = \frac{c\varepsilon}{\sqrt{\frac{1}{\xi} \log \frac{1}{p} \log \left( \frac{1}{\xi} \log \frac{1}{p} \right)}} ,$$

$c > 0$  being a universal constant.

For completeness, we prove Lemma B in Appendix B.

*Proof of Theorem 6.* According to Corollary 3 the random matrix  $A$  satisfies the condition 1. of Lemma B with  $q = 1/n$ . According to Corollary 5  $A$  also satisfies 2., with  $p = p_\lambda$ ,  $C = C_\lambda$  and  $\varepsilon = c_\lambda\sqrt{\xi}$ . Now apply Lemma B; note that  $\lambda \leq 0.95 < 1$  is bounded away from 1 and hence  $p_\lambda$  and  $C_\lambda$  may be replaced by universal constants ( $p_{0.95}$  and  $C_{0.95}$ , resp.)  $\square$

Clearly, Theorem 6 implies Theorem 1.

### 3 Operator norm of a matrix with $2k$ -wise independent entries

*Proof of Lemma 2.* We start by bounding the expectation of  $\|V\|^{2k}$ . For a real symmetric  $n \times n$  matrix  $W$ , denote by  $\lambda_1(W), \dots, \lambda_n(W)$  the eigenvalues of  $W$ , and let  $\lambda_{\max}(W) = \max_i \lambda_i(W)$ . Observe that

$$\|V\|^2 = \lambda_{\max}(V^t V) = \lambda_{\max}(V V^t)$$

and hence:

$$\begin{aligned} \mathbb{E}\|V/\sqrt{N}\|^{2k} &= \mathbb{E}\lambda_{\max}(V V^t/N)^k \\ &\leq \mathbb{E}\sum_{i=1}^n \lambda_i(V V^t/N)^k = \mathbb{E}\text{Tr}((V V^t/N)^k). \end{aligned}$$

The trace of  $(V V^t)^k$  is equal to

$$\sum V_{i_1, j_1} V_{i_2, j_1} V_{i_2, j_2} V_{i_3, j_2} \cdots V_{i_k, j_k} V_{i_1, j_k},$$

where the sum is over closed paths  $(i_1, j_1, \dots, i_k, j_k, i_1)$  in the bipartite graph  $K_{n, N}$ . The expectation of each term in the sum is 0 if there is some  $V_{i, j}$  that appears an odd number of times, and 1 if all the terms appear an even number of times. So, the expectation is equal to the number  $m(k; n, N)$  of closed even paths of length  $2k$  in  $K_{n, N}$ , starting on the side of size  $n$  (an even path is a path in which every edge appears an even number of times).

Instead of estimating this expectation directly, we follow an idea of Aubrun [5] and take a different route. The trace of  $(V V^t)^k$  is a sum over products of powers of at most  $2k$  elements from  $V$ , and so, since the elements of  $V$  come from a  $2k$ -wise independent probability space, the expectation is the same as if the elements of  $V$  were truly independent. Hence, we can use estimates known for matrices with i.i.d. elements.

We chose to use such an estimate for matrices with Gaussian i.i.d elements. Let  $\tilde{V}$  be an  $n \times N$  matrix, whose entries are independent,  $\tilde{V}_{i, j} \sim N(0, 1)$ . For every entry  $1 \leq i \leq n$ ,  $1 \leq j \leq N$  and every integer  $l \geq 1$  we have:

$$\mathbb{E}\tilde{V}_{i, j}^{2l} \geq (\mathbb{E}\tilde{V}_{i, j}^2)^l = 1 = \mathbb{E}V_{i, j}^{2l}; \quad \mathbb{E}\tilde{V}_{i, j}^{2l+1} = 0 = \mathbb{E}V_{i, j}^{2l+1}.$$



Therefore

$$\begin{aligned} \mathbb{E}\text{Tr}((VV^t/N)^k) &\leq \mathbb{E}\text{Tr}((\tilde{V}\tilde{V}^t/N)^k) = \mathbb{E} \sum_{i=1}^n \lambda_i(\tilde{V}\tilde{V}^t/N)^k \\ &\leq n\mathbb{E}\lambda_{\max}(\tilde{V}\tilde{V}^t/N)^k = n\mathbb{E}\|\tilde{V}/\sqrt{N}\|^{2k} . \end{aligned}$$

We use the following bound for Gaussian random matrices with independent entries (see Davidson–Szarek [6, Thm. II.13], extending an idea of Y. Gordon):

$$\mathbb{P} \left\{ \|\tilde{V}/\sqrt{N}\| \geq 1 + \sqrt{\xi} + t \right\} < \exp(-Nt^2/2) , \quad t \geq 0 .$$

Now,

$$\begin{aligned} \mathbb{E}\|\tilde{V}/\sqrt{N}\|^{2k} &= \int_0^\infty 2kt^{2k-1} \mathbb{P} \left\{ \|\tilde{V}/\sqrt{N}\| \geq t \right\} dt \\ &< (1 + \sqrt{\xi})^{2k} + 2k \int_0^\infty (1 + \sqrt{\xi} + u)^{2k-1} \exp(-Nu^2/2) du . \end{aligned}$$

It is easy to see that the second term is smaller than the first one:

$$\begin{aligned} &2k \int_0^\infty (1 + \sqrt{\xi} + u)^{2k-1} \exp(-Nu^2/2) du \\ &< 2k(1 + \sqrt{\xi})^{2k-1} \int_0^\infty \exp \left\{ \frac{2k-1}{1 + \sqrt{\xi}} u - Nu^2/2 \right\} du \\ &< \frac{2k}{\sqrt{N}} (1 + \sqrt{\xi})^{2k-1} \int_{-\infty}^\infty \exp \left\{ \frac{2k-1}{\sqrt{N} + \sqrt{n}} u - u^2/2 \right\} du \\ &= (1 + \sqrt{\xi})^{2k-1} \frac{\sqrt{8\pi} k}{\sqrt{N}} \exp \left\{ \frac{1}{2} \left( \frac{2k-1}{\sqrt{N} + \sqrt{n}} \right)^2 \right\} \\ &= (1 + \sqrt{\xi})^{2k} \times O(k/\sqrt{N}) \times e^{O(k^2/N)} . \end{aligned}$$

If  $k \leq c_2\sqrt{N}$  (for an appropriately chosen numerical constant  $c_2 > 0$ ), the product of the  $O$ -terms is not greater than 1. Hence

$$\mathbb{E}\|\tilde{V}/\sqrt{N}\|^{2k} < 2(1 + \sqrt{\xi})^{2k} ,$$

implying that

$$\mathbb{E}\|V/\sqrt{N}\|^{2k} < 2n(1 + \sqrt{\xi})^{2k} .$$

Now by Chebyshev's inequality

$$\mathbb{P} \left\{ \|V/\sqrt{N}\| \geq 1 + \sqrt{\xi} + t \right\} \leq \frac{\mathbb{E}\|V/\sqrt{N}\|^{2k}}{(1 + \sqrt{\xi} + t)^{2k}} < 2n \left( \frac{1 + \sqrt{\xi}}{1 + \sqrt{\xi} + t} \right)^{2k}$$

□

*Remarks.*

1. The lemma shows that for  $k = \Omega(\log N)$  the operator norm of  $V/\sqrt{N}$  is not much larger than  $1 + \sqrt{\xi}$ . This matches the bound for matrices with independent entries (cf. Geman [8]).
2. A more direct proof would be to bound the numbers  $m(k; n, N)$  directly, as in the work of Geman [8]. This would yield an estimate similar to the one we get.

## 4 Bound for a single vector

Fix  $x$ ,  $\|x\|_2 = 1$ ; let us bound the probability

$$\mathbb{P} \left\{ \|Ax\|_2 < 6\varepsilon\sqrt{N} \right\}$$

when  $A = A_1 \bullet A_2$ ,  $A_1$  is a fixed sign matrix and  $A_2$  is generated from a random walk on an expander as explained in Section 2.

Recall that  $G = (\mathcal{V}, \mathcal{E})$  is a  $d$ -regular graph with  $2^{O(\log N)}$  vertices, and  $P^G$  is the transition matrix of the random walk on  $G$ ;  $\lambda$  is the second largest absolute value of an eigenvalue of  $P^G$ .

First we bound from below the probability that a coordinate of  $Ax$  is not very small.

**Lemma 7.** *Let  $\Psi$  be a random vector in  $\{-1, +1\}^N$  with 4-wise independent coordinates. Then*

$$\mathbb{P} \left\{ \langle \Psi, x \rangle^2 \geq 1/2 \right\} \geq 1/12 .$$

*Proof.* First,

$$\begin{aligned} \mathbb{E} \langle \Psi, x \rangle^2 &= \sum_{i,j=1}^N x_i x_j \mathbb{E} \Psi_i \Psi_j = \sum_{i=1}^N x_i^2 = 1 ; \\ \mathbb{E} \langle \Psi, x \rangle^4 &= \sum_{i,j,k,l=1}^N x_i x_j x_k x_l \mathbb{E} \Psi_i \Psi_j \Psi_k \Psi_l \\ &= \sum_{i=1}^N x_i^4 + 6 \sum_{1 \leq i < j \leq N} x_i^2 x_j^2 < 3 \left( \sum_{i=1}^N x_i^2 \right)^2 = 3 . \end{aligned}$$

Recall the Paley–Zygmund inequality [14]:

**Lemma** (Paley–Zygmund). *If  $Z \geq 0$  is a random variable with finite second moment,  $0 < \theta < 1$ , then*

$$\mathbb{P}\{Z \geq \theta \mathbb{E}Z\} \geq (1 - \theta)^2 \frac{\mathbb{E}(Z)^2}{\mathbb{E}(Z^2)}.$$

Applying the inequality for  $Z = \langle \Psi, x \rangle^2$ ,  $\theta = 1/2$ , we obtain the statement of the lemma. □

*Proof of Lemma 4.* Let us show that a constant fraction of the rows  $\psi_i$  of  $A$  satisfy w.h.p

$$\langle \psi_i, x \rangle \geq 1/2. \quad (4)$$

For fixed  $A_1$  and  $1 \leq i \leq n$ , the coordinates of  $\psi_i$  are 4-wise independent; therefore by Lemma 7 there is a subset  $S_i \subset \mathcal{V}$  such that  $|S_i|/|\mathcal{V}| \geq 1/12$ , and the  $i$ -th  $\psi_i$  of  $A$  satisfies (4) iff the  $i$ -th row  $v_i$  of  $A_2$  lies in  $S_i$ .

We need a modification of Kahale’s Chernoff-type bound on expanders [12], see also Alon, Feige, Wigderson and Zuckerman [2, Theorem 4], Artstein-Avidan and Milman [3, Section 4], and Hoory, Linial and Wigderson [9, Theorem 3.11] for related results<sup>1</sup>

**Lemma 8.** *Let  $G = (\mathcal{V}, \mathcal{E})$  be a graph; as before, let  $1 = \lambda_1 \geq \lambda_2 \geq \lambda_3 \geq \dots$  be the eigenvalues of  $P^G$ ; denote  $\lambda = \max_{i \geq 2} |\lambda_i|$ . The probability that a random walk on  $G$ , starting from a random point in  $\mathcal{V}$ , is in  $S_i$  on the  $i$ -th step,  $i = 1, 2, \dots, k$ , is at most*

$$\prod_{i=1}^{k-1} \sqrt{\lambda + (1 - \lambda) \frac{|S_i|}{|\mathcal{V}|}} \sqrt{\lambda + (1 - \lambda) \frac{|S_{i+1}|}{|\mathcal{V}|}}.$$

*Proof of Lemma 8.* Denote  $e = (1, 1, \dots, 1)/\sqrt{|\mathcal{V}|}$ , and denote by  $\Pi_i$  the projector on the coordinates in  $S_i$ . Then the probability in question equals

$$\begin{aligned} & \langle \Pi_k P^G \Pi_{k-1} P^G \dots P^G \Pi_1 e, e \rangle \\ & \leq \|\Pi_k P^G \Pi_{k-1}\| \times \|\Pi_{k-1} P^G \Pi_{k-2}\| \times \dots \times \|\Pi_2 P^G \Pi_1\|, \quad (5) \end{aligned}$$

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<sup>1</sup>Added in proof: an even stronger result was recently proved. See theorem 5.4 in E. Mossel, R. O’Donnell, O. Regev, J. Steif and B. Sudakov, Non-Interactive Correlation Distillation, Inhomogeneous Markov Chains and the Reverse Bonami-Beckner Inequality, Israel Journal of Mathematics 154 (2006), 299-336.

where we used the submultiplicativity of operator norm and the equality  $\Pi_i^2 = \Pi_i$ . Let us bound the norms

$$\|\Pi_{i+1}P^G\Pi_i\| = \max_{\|g\|_2=1} \|\Pi_{i+1}P^G\Pi_i g\|_2 .$$

First of all, the vector  $g$  for which the maximum is attained is supported in  $S_i$ ; hence  $\Pi_i g = g$ . Let us decompose  $g = \alpha e + \beta v$ , where  $\alpha^2 + \beta^2 = 1$  and  $v$  is a unit vector orthogonal to  $e$ .

Note that

$$|\alpha| = |\langle g, e \rangle| \leq \|g\|_1 / \sqrt{|\mathcal{V}|} \leq \sqrt{\frac{|S_i|}{|\mathcal{V}|}} \|g\|_2 = \sqrt{\frac{|S_i|}{|\mathcal{V}|}} .$$

Therefore  $P^G g = \alpha e + \beta P^G v$ . Now,

$$\|\Pi_{i+1}P^G g\|_2 = \max_{\|h\|_2=1} \langle \Pi_{i+1}P^G g, h \rangle = \max_{\|h\|_2=1} \langle P^G g, \Pi_{i+1} h \rangle ;$$

we may assume that  $h$  is supported in  $S_{i+1}$ . Let  $h = \alpha' e + \beta' v'$ , where  $v'$  is a unit vector orthogonal to  $e$ ; as before,

$$\alpha'^2 + \beta'^2 = 1 \quad \text{and} \quad |\alpha'| \leq \sqrt{\frac{|S_{i+1}|}{|\mathcal{V}|}} .$$

Hence

$$\begin{aligned} \langle P^G g, h \rangle &= \alpha\alpha' + \beta\beta' \langle P^G v, v' \rangle \leq \alpha\alpha' + \lambda\beta\beta' \\ &\leq \sqrt{\alpha^2 + \lambda\beta^2} \sqrt{\alpha'^2 + \lambda\beta'^2} \\ &= \sqrt{\lambda + (1-\lambda)\alpha^2} \sqrt{\lambda + (1-\lambda)\alpha'^2} \\ &\leq \sqrt{\lambda + (1-\lambda)\frac{|S_i|}{|\mathcal{V}|}} \sqrt{\lambda + (1-\lambda)\frac{|S_{i+1}|}{|\mathcal{V}|}} . \end{aligned}$$

□

Now, if  $\|Ax\|_2 < 6\varepsilon\sqrt{N}$ ,  $A$  has at most  $72\varepsilon^2 N$  rows  $\psi$  such that

$$\langle \psi, x \rangle^2 \geq 1/2 .$$

By Lemma 8, the probability of this event is at most

$$\begin{aligned} &\binom{n}{\lceil 72\varepsilon^2 N \rceil} \left( \frac{11}{12}(1-\lambda) + \lambda \right)^{n-\lceil 72\varepsilon^2 N \rceil-1} \\ &\leq 2 \left( \frac{e\xi}{72\varepsilon^2} \right)^{72n\varepsilon^2/\xi} \left( \frac{11}{12}(1-\lambda) + \lambda \right)^{n-72n\varepsilon^2/\xi} . \quad (6) \end{aligned}$$

For  $\varepsilon$  small enough, this probability is exponentially small. More formally, it is easy to see that there exist some constants  $C_\lambda \geq 1 > c_\lambda > 0$  and  $0 < p_\lambda < 1$  depending only on  $\lambda$ , such that

$$\mathbb{P} \left\{ \|Ax\|_2 < 6\varepsilon\sqrt{N} \right\} \leq C_\lambda p_\lambda^n \quad \text{if } 0 < \varepsilon \leq c_\lambda \sqrt{\xi}. \quad (7)$$

Lemma 4 is proved. □

## A Construction of $k$ -wise independent random bits

For completeness, we recall the construction of  $2^r - 1$   $k$ -wise independent random bits from  $kr$  independent random bits due to Alon, Babai and Itai [1]. It will be more convenient to work with vectors of  $\{0, 1\}$  rather than  $\{-1, +1\}$ .

Let

$$\alpha_1, \dots, \alpha_{2^r-1} \in \text{GF}(2^r)$$

be the non-zero elements of the finite field of cardinality  $2^r$ .  $\text{GF}(2^r)$  is a linear space over  $\text{GF}(2)$ ; hence we may represent an element  $\alpha \in \text{GF}(2^r)$  as an  $r$ -tuple  $\tilde{\alpha} \in \text{GF}(2)^r$ .

Consider the matrix

$$M = \begin{pmatrix} 1 & \alpha_1 & \alpha_1^2 & \cdots & \alpha_1^{k-1} \\ 1 & \alpha_2 & \alpha_2^2 & \cdots & \alpha_2^{k-1} \\ \cdots & \cdots & \cdots & \cdots & \cdots \\ 1 & \alpha_{2^r-1} & \alpha_{2^r-1}^2 & \cdots & \alpha_{2^r-1}^{k-1} \end{pmatrix}.$$

Every  $k$  rows of  $M$  form a Van der Monde matrix, and in particular are linearly independent. Let

$$\tilde{M} = \begin{pmatrix} 1 & \tilde{\alpha}_1 & \tilde{\alpha}_1^2 & \cdots & \tilde{\alpha}_1^{k-1} \\ 1 & \tilde{\alpha}_2 & \tilde{\alpha}_2^2 & \cdots & \tilde{\alpha}_2^{k-1} \\ \cdots & \cdots & \cdots & \cdots & \cdots \\ 1 & \tilde{\alpha}_{2^r-1} & \tilde{\alpha}_{2^r-1}^2 & \cdots & \tilde{\alpha}_{2^r-1}^{k-1} \end{pmatrix}$$

be the corresponding  $kr \times (2^r - 1)$  matrix over  $\text{GF}(2)$ ; its rows are also linearly independent. Now let  $Z$  be a random vector distributed uniformly in  $\text{GF}(2)^{kr}$ ; let  $X = \tilde{M}Z$ .

*Claim.* The coordinates of the vector  $X$  are  $k$ -wise independent.

*Proof.* For every set of indices  $\emptyset \neq I \subset \{1, \dots, 2^r - 1\}$  such that  $|I| = k$ , the matrix  $\widetilde{M}_I$  formed from the corresponding rows of  $\widetilde{M}$  is of rank  $k$ ; that is,  $\widetilde{M}_I$  is surjective and the preimages of the vectors in  $\{0, 1\}^k$  are of equal size. The vector  $Z$  is distributed uniformly in  $\text{GF}(2)^{kr}$ ; hence the vector  $(X_i)_{i \in I} = \widetilde{M}_I Z$  is uniformly distributed in  $\text{GF}(2)^k$ .  $\square$

## B Proof of Lemma B

The proof of Lemma B is based on  $\varepsilon$ -net arguments.

**Definition 5.** Let  $S \subset \mathbb{R}^N$  be a convex set. A (finite) subset  $\mathcal{N} \subset S$  is called an  $\varepsilon$ -net in  $S$  if for every  $x \in S$  there exists  $y \in \mathcal{N}$  such that  $\|x - y\|_2 \leq \varepsilon$ .

*Notation.* Let  $t > 0$  and let  $K \subset \mathbb{R}^n$  be a convex body. As usual, denote

$$tK = \{tx \mid x \in K\} .$$

Similarly to [3], we use the following result, due to Schütt [15]:

**Theorem (Schütt).** *The exists a universal constant  $c > 0$  such that for any  $\zeta > 0$  and  $\theta \geq c\sqrt{\frac{1}{\zeta} \log \frac{1}{\zeta}}$  there exists a  $\theta$ -net  $\mathcal{N}$  in  $\sqrt{N}B_1^N$  such that  $|\mathcal{N}| \leq e^{\zeta N}$ .*

*Proof of Lemma B.* Pick  $0 < \zeta < \xi \log \frac{1}{p}$ ; then  $e^\zeta < 1/p^\xi$ . Set

$$\delta = \frac{\varepsilon}{c\sqrt{\frac{1}{\zeta} \log \frac{1}{\zeta}}} .$$

Scaling the result of Schütt's theorem times  $\delta$ , we get an  $\varepsilon$ -net  $\mathcal{N}$  in  $\delta\sqrt{N}B_1^N$ ,  $|\mathcal{N}| \leq e^{\zeta N}$ .

By our assumptions, for every  $y \in \mathcal{N}$

$$\mathbb{P} \left\{ \|Ay\|_2 < 6\varepsilon\sqrt{N}\|y\|_2 \right\} < Cp^n ,$$

and so the probability that there exists  $y \in \mathcal{N}$  with

$$\|Ay\|_2 < 6\varepsilon\sqrt{N}\|y\|_2$$

is at most

$$Ce^{\zeta N} p^n = p^{\Theta(n)} .$$

Assume that for every  $y \in \mathcal{N}$  we have

$$\|Ay\|_2 \geq 6\varepsilon\sqrt{N}\|y\|_2 ,$$

and also that  $\|A\| \leq 3\sqrt{N}$ . This event happens with probability at least  $1 - q - p^{\Theta(n)}$ . We will show that whenever these two conditions hold, every  $x \in \text{Ker}A$  satisfies

$$\|x\|_1 \geq \delta\sqrt{N}\|x\|_2 .$$

It is enough to show this for  $x$  with  $\|x\|_2 = 1$ .

Take any  $x \in \mathbb{R}^N$  with  $\|x\|_1 < \delta\sqrt{N}$  and  $\|x\|_2 = 1$ . We will show  $x \notin \text{Ker}(A)$ . First,  $x \in \delta\sqrt{N}B_1^N$ , and so there exists  $y \in \mathcal{N}$  such that  $\|x - y\| \leq \varepsilon$ . Now we have:

$$\begin{aligned} \|Ax\|_2 &\geq \|Ay\|_2 - \|A(x - y)\|_2 \geq 6\varepsilon\sqrt{N}\|y\|_2 - \|A\|\|x - y\|_2 \\ &\geq 6\varepsilon(1 - \varepsilon)\sqrt{N} - 3\varepsilon\sqrt{N} > 0 , \end{aligned}$$

where we used the fact that

$$\|y\|_2 \geq \|x\|_2 - \|x - y\|_2 \geq 1 - \varepsilon .$$

□

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