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Pseudorandom Sets in Grassmann Graph have Near-Perfect Expansion

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Abstract

We prove that pseudorandom sets in Grassmann graph have near-perfect expansion as hypothesized in [4]. This completes the proof of the 2-to-2 Games Conjecture (albeit with imperfect completeness) as proposed in [15, 3], along with a contribution from [2, 14].

The Grassmann graph $\mathsf{Gr}_{\mathsf{global}}$ contains induced subgraphs $\mathsf{Gr}_{\mathsf{local}}$ that are themselves isomorphic to Grassmann graphs of lower orders. A set is called pseudorandom if its density is $o(1)$ inside all subgraphs $\mathsf{Gr}_{\mathsf{local}}$ whose order is $O(1)$ lower than that of $\mathsf{Gr}_{\mathsf{global}}$. We prove that pseudorandom sets have expansion $1 - o(1)$, greatly extending the results and techniques in [4].

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1 Introduction

The Unique Games Conjecture is a prominent question in theoretical computer science (please see the surveys [19, 11, 9, 12]). The focus of this paper is the closely related 2-to-2 Games Conjecture and even more specifically, a combinatorial hypothesis that was posed in [4] towards proving the 2-to-2 Games Conjecture. For the purposes of this paper, it suffices to define the Unique Game and the 2-to-2 Game as the following computational problems. Let $\mathbb{F}_2^\ell$ denote the ℓ-dimensional vector space over the binary field $\mathbb{F}_2$, considered as an additive group with the $\oplus$ operation.

Definition 1.1. *An instance $\mathcal{U}$ of the* $\mathsf{UniqueGame}[\mathbb{F}_2^\ell]$ *problem consists of n variables $x_1, \ldots, x_n$ taking values over (the alphabet) $\mathbb{F}_2^\ell$ and m constraints $C_1, \ldots, C_m$ where each constraint is a linear equation of the form $T_{ij}x_i \oplus T'_{ij}x_j = b_{ij}$, T_{ij}, T'_{ij} are $\ell \times \ell$ invertible matrices, and $b_{ij} \in \mathbb{F}_2^\ell$. Let $\mathsf{OPT}(\mathcal{U})$ denote the maximum fraction of the constraints that can be satisfied by any assignment to the instance.*

For constants $0 < s < c < 1$, let $\mathsf{GapUG}[\mathbb{F}_2^\ell](c, s)$ be the gap-version where the instance $\mathcal{U}$ of the $\mathsf{UniqueGame}[\mathbb{F}_2^\ell]$ problem is promised to have either $\mathsf{OPT}(\mathcal{U}) \geqslant c$ or $\mathsf{OPT}(\mathcal{U}) \leqslant s$. The Unique Games Conjecture states that[1]

Conjecture 1.2. *For every constant $\varepsilon > 0$, there exists a sufficiently large integer $\ell = \ell(\varepsilon)$ such that* $\mathsf{GapUG}[\mathbb{F}_2^\ell](1-\varepsilon, \varepsilon)$ *is* NP*-hard.*

Definition 1.3. *An instance $\mathcal{U}_{2\leftrightarrow 2}$ of the* 2-to-2 $\mathsf{Game}[\mathbb{F}_2^\ell]$ *problem consists of n variables $x_1, \ldots, x_n$ taking values over (the alphabet) $\mathbb{F}_2^\ell$ and m constraints $C_1, \ldots, C_m$ where each constraint is of the form $T_{ij}x_i \oplus T'_{ij}x_j \in \{b_{ij}, b'_{ij}\}$, T_{ij}, T'_{ij} are $\ell \times \ell$ invertible matrices, and $b_{ij}, b'_{ij} \in \mathbb{F}_2^\ell$. Let $\mathsf{OPT}(\mathcal{U}_{2\leftrightarrow 2})$ denote the maximum fraction of the constraints that can be satisfied by any assignment to the instance.*

For constants $0 < s < c \leqslant 1$, let Gap 2-to-2$[\mathbb{F}_2^\ell](c, s)$ be the gap-version where the instance $\mathcal{U}_{2\leftrightarrow 2}$ of the 2-to-2 $\mathsf{Game}[\mathbb{F}_2^\ell]$ problem is promised to have either $\mathsf{OPT}(\mathcal{U}_{2\leftrightarrow 2}) \geqslant c$ or $\mathsf{OPT}(\mathcal{U}_{2\leftrightarrow 2}) \leqslant s$. We will refer to the statement below as the 2-to-2 Games Conjecture[2] (stated as a theorem thanks to the main result of this work and the previous works as explained below):

Theorem 1.4. *For every constant $\varepsilon > 0$, there exists a sufficiently large integer $\ell = \ell(\varepsilon)$ such that* Gap 2-to-2$[\mathbb{F}_2^\ell](1-\varepsilon, \varepsilon)$ *is* NP*-hard.*

Though weaker than the Unique Games Conjecture, the 2-to-2 Games Conjecture has important applications of its own and evidence in its favor also serves as evidence in favor of the Unique Games Conjecture. In recent line of work [15, 3, 4], the authors proposed an approach towards proving the 2-to-2 Games Conjecture. The authors were able to formulate a combinatorial hypothesis about the structure of non-perfectly expanding sets in the Grassmann graph and (along with an important contribution by Barak, Kothari, and Steurer [2]) prove that this hypothesis implies the 2-to-2 Games Conjecture. Instead of repeating what has

[1] The original statement in [10] refers to more general constraints. However it follows from [13] that the original conjecture is equivalent to the statement here, i.e. when the constraints are linear equations over the group $\mathbb{F}_2^\ell$ (and even when the matrices T_{ij}, T'_{ij} are identity matrices).

[2] Comments regarding the original formulation of this conjecture in [10]: (1) It was proposed with *perfect completeness*, i.e. stating that Gap 2-to-2$[\mathbb{F}_2^\ell](1, \varepsilon)$ is NP-hard. However, as far as this paper is considered, we view the issue of perfect versus imperfect completeness as being relatively minor. (2) It was proposed with more general constraints (rather than the special case with linear structure described herein) and with "2-to-1" constraints (rather than with "2-to-2" constraints described herein; the conjecture was referred to as the 2-to-1 Conjecture). Both these are non-issues however: the current and preceding works [15, 3, 4] now prove the conjecture *with* linear structure and the constraints are easily reinterpreted as being 2-to-1 constraints (hence proving the 2-to-1 Conjecture as well, which in any case is morally equivalent to the 2-to-2 Conjecture).

already been said before, we refer the reader to the introductory sections of the papers [15, 3, 4] for a detailed overview of this development. A brief overview along with the significance of proving the 2-to-2 Games Conjecture is sketched in Sections B and C. The focus of this paper is the combinatorial hypothesis itself, which we are able to prove, in turn proving the 2-to-2 Games Conjecture, completing this line of work. Preliminary steps towards proving the combinatorial hypothesis were already taken in [4] and the current paper has also benefited from the insights in [14] (please see the remark at the end of this section).

In the following, we introduce the Grassmann graph and state the combinatorial hypothesis. Consider an n-vertex d-regular graph $G(V, E)$ and a non-empty set of vertices $S \subseteq V$ with $|S| \leqslant \frac{n}{2}$. Its (edge-)expansion is defined as

$$\Phi(S) = \frac{|E(S, \overline{S})|}{d \cdot |S|},$$

where $E(S, \overline{S})$ denotes the set of edges with one endpoint in S and the other in $\overline{S} = V \setminus S$. Alternately, it is the probability that selecting a uniformly random vertex in S and moving along a uniformly random edge incident on that vertex, one lands outside S. It is clear that if S is a randomly selected set of size $o(n)$, then its expansion is $1 - o(1)$ (with high probability over the choice of the set), i.e. small random sets have near-perfect expansion.

Let k, ℓ be integer parameters with $1 \ll \ell \ll k$. We will be interested in the Grassmann graph $\mathsf{Gr}_{k,\ell}$ and subsets S of its vertices that have expansion strictly bounded away from 1 (say at most $\frac{31}{32}$). Such sets will be referred to as "non-perfectly expanding".

Definition 1.5. *The Grassmann graph $\mathsf{Gr}_{k,\ell}$ is defined as follows. Its vertex set consists of all ℓ-dimensional subspaces L of $\mathbb{F}_2^k$ and (L, L') is an edge if and only if $\mathsf{dim}(L \cap L') = \ell - 1$.*

Definition 1.6. *Suppose $A \subseteq B \subseteq \mathbb{F}_2^k$ are subspaces. Let $\mathsf{dim}(A) = a, \mathsf{codim}(B) = b$ and think of a, b as small constants (say $a = b = 2$). Then the subgraph $\mathsf{Gr}_{k,\ell}[A, B]$ is an induced subgraph of $\mathsf{Gr}_{k,\ell}$ induced on precisely the set of vertices L such that $A \subseteq L \subseteq B$. It is easily seen that $\mathsf{Gr}_{k,\ell}[A, B]$ is an isomorphic copy of a lower order Grassmann graph $\mathsf{Gr}_{k-a-b,\ell-a}$. We call $a + b$ as the co-order of $\mathsf{Gr}_{k,\ell}[A, B]$ with respect to $\mathsf{Gr}_{k,\ell}$.*

The sets $\mathsf{Gr}_{k,\ell}[A, B]$ are natural examples of sets in $\mathsf{Gr}_{k,\ell}$ that have expansion strictly bounded away from 1 (when a, b are small constants). Indeed, the expansion of $\mathsf{Gr}_{k,\ell}[A, B]$, when seen as a subset of $\mathsf{Gr}_{k,\ell}$, has expansion precisely $1 - 2^{-(a+b)}$ (up to an error $O(2^{-\ell})$ which is thought of as negligible and ignored for the ease of presentation). The reasoning is as follows. For a vertex $L \in \mathsf{Gr}_{k,\ell}[A, B]$, its random neighbor L' is obtained by picking a random subspace $T \subseteq L$, $\mathsf{dim}(T) = \ell - 1$ and a random point $x \in \mathbb{F}_2^k \setminus L$ and letting $L' = T \oplus \mathsf{Span}(x)$. Now $L' \in \mathsf{Gr}_{k,\ell}[A, B]$ if and only if $A \subseteq T$ and $x \in B$ and these events happen independently with probabilities 2^{-a} and 2^{-b} respectively (up to an error $O(2^{-\ell})$). Thus a random neighbor of a random vertex in $\mathsf{Gr}_{k,\ell}[A, B]$ is also inside it with probability $2^{-(a+b)}$ and hence its expansion is $1 - 2^{-(a+b)}$. Furthermore, we observe that if $S \subseteq \mathsf{Gr}_{k,\ell}[A, B] \subseteq \mathsf{Gr}_{k,\ell}$ is such that

$$\frac{|S|}{|\mathsf{Gr}_{k,\ell}[A, B]|} = \varepsilon,$$

then $\Phi(S) \leqslant 1 - \varepsilon \cdot 2^{-(a+b)}$. This is because (we skip the easy proof) any set of density ε inside a Grassmann graph has at least ε^2 fraction of the edges inside it (and hence has expansion at most $1 - \varepsilon$). Therefore, a random neighbor of a random vertex in $S \subseteq \mathsf{Gr}_{k,\ell}[A, B]$ lies inside $\mathsf{Gr}_{k,\ell}[A, B]$ with probability $2^{-(a+b)}$ as seen above and then inside S with probability at least ε, justifying the observation. We summarize the overall observation as:

Fact 1.7. *(Informal): A subset of constant density inside a constant co-order copy of Grassmann graph inside a Grassmann graph has expansion strictly bounded away from* 1.

(Formal): Let $S \subseteq \mathsf{Gr}_{k,\ell}[A,B] \subseteq \mathsf{Gr}_{k,\ell}$ *be such that* $\mathsf{dim}(A) = a, \mathsf{codim}(B) = b$ *and the density of* S *inside* $\mathsf{Gr}_{k,\ell}[A,B]$ *is* ε*. Then* $\Phi(S) \leqslant 1 - \varepsilon \cdot 2^{-(a+b)}$.

The authors of [4] hypothesize, essentially, that the converse of the above fact is true. Informally, their hypothesis is that any set S in the Grassmann graph $\mathsf{Gr}_{k,\ell}$ whose expansion is strictly bounded away from 1 has constant density inside *some* copy of Grassmann graph of constant co-order. A precise statement appears below (now as a theorem and the main result in this paper):

Theorem 1.8. *For every constant* $0 < \alpha < 1$*, there exists a constant* $\varepsilon > 0$ *and an integer* $r \geqslant 0$ *such that for all sufficiently large integers* ℓ *and (after fixing it) for all sufficiently large integers* k*, the following holds. let* $S \subseteq \mathsf{Gr}_{k,\ell}$ *be such that* $\Phi(S) \leqslant \alpha$*. Then there exist subspaces* $A \subseteq B \subseteq \mathbb{F}_2^k$ *such that* $\mathsf{dim}(A) = a, \mathsf{codim}(B) = b$*,* $a + b \leqslant r$ *and*

$$\frac{|S \cap \mathsf{Gr}_{k,\ell}[A,B]|}{|\mathsf{Gr}_{k,\ell}[A,B]|} \geqslant \varepsilon.$$

It has already been shown in [15, 3, 4] (along with an important contribution from Barak, Kothari, and Steurer [2]) that the above Theorem 1.8 implies the 2-to-2 Games Conjecture! In [4], the authors prove Theorem 1.8 when $\alpha < \frac{7}{8}$, via spectral analysis of the Grassmann graph, introduced therein (the eigenvalues and eigenspaces of the Grassmann graph were known before). Roughly speaking, given a set S with expansion at most $\alpha < 1 - 2^{-(s+1)}$, it is easily observed that the indicator vector of the set $\mathbf{1}_S$ must have a significant projection onto the eigenspace at "level" at most s (s is a constant when α is strictly bounded away from 1). The spectral analysis then attempts to use this projection to deduce the desired structure of S. The approach is worked out in [4] when $s = 2$, corresponding to $\alpha < \frac{7}{8}$. It already requires rather difficult and lengthy case analysis. In principle, the same approach could be extended to higher levels $s \geqslant 3$, but the number of cases to handle seems to explode beyond control. Instead, we are able to argue in a more systematic fashion and avoid the explosion in potential case analysis (easier said than done of course).

We end this section with some remarks on Theorem 1.8: (1) The subspaces A and B therein are referred to as "zoom-in" and "zoom-out" spaces respectively [15, 3, 4]. This makes sense if one imagines searching for the appropriate subgraph $\mathsf{Gr}_{k,\ell}[A,B]$ where the set S happens to have significant density. (2) We note that if S has density $\geqslant \varepsilon$, then the conclusion of the theorem is vacuously true without any need for a zoom-in or a zoom-out (i.e. $a = b = 0, A = \{0\}, B = \mathbb{F}_2^k$), so the theorem is really about "small" sets. (3) Our proof gives correct dependence of the required zoom-in-out dimension r on the upper bound on expansion α. For $\alpha < 1 - 2^{-(s+1)}$, one gets a significant projection onto the eigenspace at level at most s and then in our proof, a (combined) zoom-in-out dimension of at most $r = s$ is needed. This is tight (i.e. a lesser zoom-in-out dimension is not sufficient) since we know that subgraphs $\mathsf{Gr}_{k,\ell}[A,B]$ have expansion $1 - 2^{-(a+b)}$ and the zoom-in-out dimension $a + b$. (4) We note that towards proving the theorem, it will be easier to work with the contra-positive statement: a set S that has very small density inside every copy of the Grassmann graph with constant co-order (such a set will be called pseudorandom) has near-perfect expansion (i.e. very near 1). (5) The phenomenon as in Theorem 1.8 occurs also in the Johnson graph and has been analyzed in [14]. In a Johnson graph, the vertices are ℓ-subsets of a k-set and the edges are pairs of ℓ-subsets with intersection of size t (we are concerned with the case when $t = \lfloor \frac{\ell}{2} \rfloor$). Therein the notion of zoom-out and Fourier analysis are not needed. The Johnson case can informally be seen as a special case of the Grassmann case and the analysis of the former in [14] has been insightful in the analysis of the latter in the current paper.

2 Preliminaries

In this section, we recall and summarize the high-level plan towards proving Theorem 1.8 (or rather the contra-positive), developed already in [4]. The task boils down to upper-bounding the fourth moment of the "projection of the indicator function $\mathbf{1}_S$ onto the Fourier level-r". While this was accomplished in [4] for $r \leqslant 2$, the authors therein were unable to extend it further for two reasons:

- It relied on rather ad hoc case analysis.
- The Fourier analysis on $\mathsf{Gr}_{k,\ell}$ is unfriendly. It is futile to write down the eigenvectors explicitly and one instead works with the (eigen)space spanned by all eigenvectors with a specific eigenvalue, referred to as the "Fourier level".
  - The r^{th} level has eigenvalue very close to 2^{-r}, but there is a error term.
  - The "union" of the eigenspaces at levels $0, 1, \ldots, r$, has a clean description: it is spanned by the indictor functions of the subgraphs $\mathsf{Gr}_{k,\ell}[A, B]$ of co-order at most r. However to get the hands on precisely the r^{th} eigenspace requires "subtracting" the contribution of the previous levels. This leads to rather unfriendly inclusion-exclusion type recursive formulas even for $r = 2$ and it is not clear how to extend these to higher levels.

  For these and additional reasons, authors of [4] worked with approximations to all the quantities and formulas of interest. Extending these approximate formulas to higher Fourier levels seems to incur error terms that are unaffordable.

We are able to circumvent both these obstacles (which to some extent go hand in hand). Firstly, while we still have to consider a large number of elaborate cases, the proof is systematic and works simultaneously for all Fourier levels r (i.e. without the number of cases exploding with r). Secondly, instead of working with the Grassmann graph $\mathsf{Gr}_{k,\ell}$, we instead work with a related graph $\mathsf{H}_{k,\ell}$ (see the definition below). Surprisingly (or perhaps not so surprisingly in hindsight) we are able to write down exact recursive formulas relating quantities at successive Fourier levels. The eigenvalues are exactly 2^{-i} providing a hint that things would fall in place, but it still takes significant effort to develop the full Fourier analytic machinery, described in Section 3. The recursive formulas therein are systematic extensions, to higher Fourier levels, of the approximate and ad hoc formulas for the second level in [4].

We now define the new graph $\mathsf{H}_{k,\ell}$ and show that the task of proving Theorem 1.8 reduces to the task of proving an analogous theorem for $\mathsf{H}_{k,\ell}$ (i.e. Theorem 2.6). Then we describe the very basics of Fourier analysis, just enough to recall the high-level plan (from [4]) that reduces the task further to showing that the indictor function $\mathbf{1}_S$ of a pseudorandom set $S \subseteq \mathsf{H}_{k,\ell}$ has low Fourier weight on low Fourier levels (see Section 2.4 and Theorem 2.13). This task is in turn reduced to that of upper-bounding the fourth moment of the Fourier level-r component of the indictor function $\mathbf{1}_S$ (see Sections 2.5, 2.6), which finally is reduced to our main technical result, Theorem 2.15, about upper-bounding closely related "4-wise correlations".

The main contribution of the paper is Section 3 onwards: the full Fourier analytic machinery is developed in Section 3, upper-bounds on pairwise and 3-wise correlations are presented in Section 4 as a warm-up, and the heart of the paper, the desired upper-bound on the 4-wise correlations, is presented in Sections 5, 6, 7.

2.1 Switching to the Graph $\mathsf{H}_{k,\ell}$

Definition 2.1. *Let $2 \leqslant \ell \leqslant k$ be integers. The vertices of the graph $\mathsf{H}_{k,\ell}$ are given by $(\{0,1\}^k)^\ell$. The edges of the graph are best described by describing how to sample a uniformly random neighbor z of an*

arbitrary vertex x. Fix a vertex $x \in (\{0,1\}^k)^\ell$ and write $x = (x_1, ..., x_\ell)$ where $x_1, ..., x_\ell \in \{0,1\}^k$. Sample $y \leftarrow \{0,1\}^k$, $b_1, \dots b_\ell \leftarrow \{0,1\}$ independently and uniformly at random. Let the neighbor of x be $z = (x_1 + b_1 \cdot y,\ x_2 + b_2 \cdot y, \dots, x_\ell + b_\ell \cdot y)$.

The two graphs $\mathsf{H}_{k,\ell}$ and $\mathsf{Gr}_{k,\ell}$ are closely related as follows:

- The vertices of $\mathsf{H}_{k,\ell}$ are ℓ-tuples of vectors in $\mathbb{F}_2^k$. The vertices of $\mathsf{Gr}_{k,\ell}$ are ℓ-dimensional subspaces of $\mathbb{F}_2^k$, or equivalently, ℓ-tuples of vectors in $\mathbb{F}_2^k$ that are linearly independent and two tuples are considered the same if their vectors have the same linear span.
- When a random vertex $x = (x_1, \dots, x_\ell)$ in $\mathsf{H}_{k,\ell}$ is sampled and then a random edge incident on it is sampled by sampling $y \leftarrow \{0,1\}^k$ and $b_1, \dots, b_\ell \in \{0,1\}$, with probability $\approx 2^{-k} + 2^{-\ell}$, either $y = 0$ or $b = (b_1, \dots, b_\ell) = 0$, and the edge is a self-loop. Otherwise $y \neq 0, b \neq 0$ and the other endpoint is $z = (x_1 + b_1 \cdot y, \dots, x_\ell + b_\ell \cdot y)$. Provided that both x, z have full ℓ-dimensional linear span (which happens with probability except $\approx 2^{\ell - k}$ and we think of $\ell \ll k$), the edge (x, z) corresponds to a uniformly random edge of the Grassmann graph.

Remark 2.2. *Barak, Kothari, and Steuter [2] have made a similar suggestion. They consider a graph whose vertices are $k \times \ell$ matrices and the edges are pairs of matrices that differ by a rank 1 matrix. In terms of our notation, this amounts to an edge (x, z) with $z = x \oplus y \otimes b, y \neq 0, b \neq 0$ and x, z are thought of as $k \times \ell$ matrices. They seem to be interested in "reducing" the $k \times \ell$ case to the $k \times k$ case; the latter is same as the graph of the "degree-2 short code test" as in [1].*

2.2 It Suffices to Work with $\mathsf{H}_{k,\ell}$

We show that Theorem 1.8 for $\mathsf{Gr}_{k,\ell}$ follows easily from the corresponding Theorem 2.6 for $\mathsf{H}_{k,\ell}$ (see below) and then we work with the graph $\mathsf{H}_{k,\ell}$ for the rest of the paper. It will be convenient to restate Theorem 1.8 in the contra-positive and in terms of "pseudo-random sets".

Definition 2.3. *A subset of vertices $S \subseteq \mathsf{Gr}_{k,\ell}$ is called (r, ε)-pseudorandom if for any subspaces $A \subseteq B \subseteq \mathbb{F}_2^k$ such that $\mathsf{dim}(A) = a, \mathsf{codim}(B) = b$, $a + b \leqslant r$, we have*

$$\mu_{\mathsf{in}(A),\mathsf{out}(B)}(S) \stackrel{def}{=} \frac{|S \cap \mathsf{Gr}_{k,\ell}[A, B]|}{|\mathsf{Gr}_{k,\ell}[A, B]|} \leqslant \varepsilon.$$

Theorem 2.4. (Theorem 1.8 restated) *For every constant $\zeta > 0$, there exists a constant $\varepsilon > 0$ and an integer $r \geqslant 0$ such that for all sufficiently large integers ℓ and (after fixing it) for all sufficiently large integers k, the following holds. If $S \subseteq \mathsf{Gr}_{k,\ell}$ is (r, ε)-pseudorandom, then $\Phi(S) \geqslant 1 - \zeta$.*

Now we show how to reduce this theorem to Theorem 2.6 below. The reasoning is straightforward. We will show that for every $S \subseteq \mathsf{Gr}_{k,\ell}$, there is a natural corresponding set $S^* \subseteq \mathsf{H}_{k,\ell}$ such that

- Lemma 2.8 below: If S is (r, ε)-pseudorandom, then S^* is (r, ε)-pseudorandom (for a similar notion of being pseudorandom in $\mathsf{H}_{k,\ell}$ and up to a negligible additive change in the parameter ε).
- Theorem 2.6 below: If S^* is (r, ε)-pseudorandom, then $\Phi(S^*) \geqslant 1 - \zeta$.
- Lemma 2.7 below: $\Phi(S) = \Phi(S^*)$ (up to a negligible additive difference) and hence $\Phi(S) \geqslant 1 - \zeta$ as desired.

We elaborate on each of the three items. The reader may wish to skip the self-evident proofs of Lemmas 2.7 and 2.8. For a set $S \subseteq \mathsf{Gr}_{k,\ell}$, the corresponding set $S^* \subseteq \mathsf{H}_{k,\ell}$ is defined naturally as

$$S^* \stackrel{def}{=} \{(x_1, \ldots, x_\ell) \mid \mathsf{dim}(\mathsf{Span}(x_1, \ldots, x_\ell)) = \ell,\ \mathsf{Span}(x_1, \ldots, x_\ell) \in S\}.$$

We note that S^* is invariant under change of basis, i.e. if $\mathsf{Span}(x_1, \ldots, x_\ell) = \mathsf{Span}(y_1, \ldots, y_\ell)$, then $(x_1, \ldots, x_\ell) \in S^*$ if and only if $(y_1, \ldots, y_\ell) \in S^*$. We call such subsets of $\mathsf{H}_{k,\ell}$ basis-invariant. Throughout the paper, we will only concern ourselves with basis-invariant subsets of $\mathsf{H}_{k,\ell}$. We note moreover that tuples in S^* are linearly independent (this is a minor issue; the only place where this is used is in the proof of Lemmas 2.7 and 2.8). The notion of pseudorandom sets in $\mathsf{H}_{k,\ell}$ is defined in a similar manner.

Definition 2.5. *A basis-invariant subset of vertices $S^* \subseteq \mathsf{H}_{k,\ell}$ is called (r, ε)-pseudorandom if for any sequence $Q = (x_1, \ldots, x_q)$ of points in $\mathbb{F}_2^k$ and a subspace $W \subseteq \mathbb{F}_2^k$ and $q + \mathsf{codim}(W) \leqslant r$, we have*

$$\mu_{\mathsf{in}(Q),\mathsf{out}(W)}(S^*) \stackrel{def}{=} \Pr_{z_{q+1},\ldots,z_\ell \in W}[(x_1, \ldots, x_q, z_{q+1}, \ldots, z_\ell) \in S^*] \leqslant \varepsilon.$$

There is a slight difference between Definitions 2.3 and 2.5. In the latter, we allow Q to be a sequence of points (so there can be linear dependencies among them) and we do not necessarily require that $Q \subseteq W$. This difference however has no significance and is to be ignored. The following is the main result in the paper. As noted, together with Lemmas 2.7 and 2.8 below, it implies Theorem 1.8 and hence proves the 2-to-2 Games Conjecture.

Theorem 2.6. *For every constant $\zeta > 0$, there exists a constant $\varepsilon > 0$ and an integer $r \geqslant 0$ such that for all sufficiently large integers ℓ and (after fixing it) for all sufficiently large integers k, the following holds. If $S^* \subseteq \mathsf{H}_{k,\ell}$ is a basis-invariant (r, ε)-pseudorandom set, then $\Phi(S) \geqslant 1 - \zeta$.*

Lemma 2.7.

$$|\Phi(S^*) - \Phi(S)| \leqslant 2^{-\ell} + 2 \cdot 2^{\ell - k}.$$

Proof. Towards proving the lemma, let $x = (x_1, ..., x_\ell) \in S^*$. Denote its random neighbor by $z = (x_1 + b_1 \cdot y, \ldots, x_\ell + b_\ell \cdot y)$ and $b = (b_1, ..., b_\ell)$. Then

$$\Phi(S^*) = \Pr_{x \in S^*, y, b}[z \notin S^*] = 2^{-\ell} \cdot 0 + (1 - 2^{-\ell}) \Pr_{x \in S^*, y, b \neq 0}[z \notin S^*]. \tag{1}$$

Note that for any $x \in S^*$, the vectors $x_1, ..., x_\ell$ are linearly independent. Let $L = \mathsf{Span}(x_1, ..., x_\ell)$ and $L' = \mathsf{Span}(x_1 + b_1 \cdot y, \ldots, x_\ell + b_\ell \cdot y)$. Conditioned on $y \notin L$, L' is ℓ-dimensional subspace. Moreover, for $b \neq 0$, its distribution is uniform over all ℓ-dimensional subspaces that intersect L in dimension $\ell - 1$. Therefore

$$\Pr_{\substack{x \in S^*, y, \\ b \neq 0}}[z \notin S^*] = \Pr_{x \in S^*, y}[y \notin L] \cdot \Pr_{\substack{x \in S^*, y, \\ b \neq 0}}[z \notin S^* \mid y \notin L] + \Pr_{x \in S^*, y}[y \in L] \cdot \Pr_{\substack{x \in S^*, y, \\ b \neq 0}}[z \notin S^* \mid y \in L].$$

Since choosing $x \in S^*$ uniformly at random corresponds to choosing $L \in S$ uniformly at random and picking a random basis, we have that the first summand equals $(1 - 2^{\ell - k})\Phi(S)$. The second summand is at most $\Pr_{x \in S^*, y}[y \in L] \leqslant 2^{\ell - k}$. Combining everything finishes the proof. □

Lemma 2.8. *If S is (r, ε)-pseudorandom, then S^* is $(r, \varepsilon + 2^{\ell + r - k})$-pseudorandom.*

Proof. Towards proving the lemma, we recall Definition 2.5 and consider any sequence $Q = (x_1, \ldots, x_q) \subseteq \mathbb{F}_2^k$ and a subspace $W \subseteq \mathbb{F}_2^k$ such that $q + \mathsf{codim}(W) \leqslant r$. If Q is a linearly dependent set, then $\mu_{\mathsf{in}(Q),\mathsf{out}(W)}(S^*) = 0$ and there is nothing to prove. So assume that Q is linearly independent.

$$\mu_{\mathsf{in}(Q),\mathsf{out}(W)}(S^*) = \Pr_{z_{q+1},\ldots,z_\ell \in W}[(x_1, \ldots, x_q, z_{q+1}, \ldots, z_\ell) \in S^*].$$

Denoting by $\mathcal{E}$ the event that $\{x_1, \ldots, x_q, z_{q+1}, ..., z_\ell\}$ are linearly independent, we have

$$\begin{aligned}\mu_{\mathsf{in}(Q),\mathsf{out}(W)}(S^*) &\leqslant \Pr[\mathcal{E}] \cdot \Pr_{z_{q+1},\ldots,z_\ell \in W}[(x_1, \ldots, x_q, z_{q+1}, \ldots, z_\ell) \in S^* \mid \mathcal{E}] + \Pr\left[\,\overline{\mathcal{E}}\,\right] \\ &= \Pr[\mathcal{E}] \cdot \mu_{\mathsf{in}(\mathsf{Span}(Q)),\mathsf{out}(Q \oplus W)}(S) + \Pr\left[\,\overline{\mathcal{E}}\,\right].\end{aligned}$$

The last equality follows from the fact that conditioned on $\mathcal{E}$, $L = \mathsf{Span}(x_1, \ldots, x_q, z_{q+1}, \ldots, z_\ell)$ is distributed uniformly among all ℓ-dimensional subspaces containing Q and contained in $Q \oplus W$, and $L \in S$ if and only if $(x_1, \ldots, x_q, z_{q+1}, \ldots, z_\ell) \in S^*$. We conclude that $\mu_{\mathsf{in}(Q),\mathsf{out}(W)}(S^*) \leqslant \varepsilon + 2^{\ell+r-k}$ by noting that S is (r, ε)-pseudorandom and hence $\mu_{\mathsf{in}(\mathsf{Span}(Q)),\mathsf{out}(Q\oplus W)}(S) \leqslant \varepsilon$ and that

$$\Pr_{z_{q+1},\ldots,z_\ell \in W}[\mathcal{E}] \geqslant 1 - \sum_{i=q}^{\ell-1} 2^{i-(k-r)} \geqslant 1 - 2^{\ell+r-k}.$$

We note moreover that $q + \mathsf{codim}(Q \oplus W) \leqslant q + \mathsf{codim}(W) \leqslant r$, so we may appeal to the (r, ε)-pseudorandomness of S. □

2.3 The Eigenvectors and Eigenvalues of $\mathsf{H}_{k,\ell}$ and Fourier Levels

One advantage of working with the graph $\mathsf{H}_{k,\ell}$ is that its vertex set is the Boolean hypercube $(\{0,1\}^k)^\ell$, it is a Cayley graph, and determining its eigenvectors and eigenvalues is straightforward.

Definition 2.9. *For $T_1, \ldots, T_k \in \{0,1\}^k$, define $\chi_{T_1,\ldots,T_\ell} \colon (\{0,1\}^k)^\ell \to \{-1, 1\}$ by*

$$\chi_{T_1,\ldots T_\ell}(x_1, \ldots, x_\ell) = \prod_{i=1}^{\ell} \chi_{T_i}(x_i),$$

where $\chi_{T_i}(x_i) = (-1)^{T_i \cdot x_i}$ is the standard Fourier character (here '$\cdot$' is the inner product over $\mathbb{F}_2$).

We denote by $\mathsf{H}_{k,\ell}$ also the normalized transition matrix of the graph $\mathsf{H}_{k,\ell}$ (i.e. its entry (x, z) equals the probability that a random neighbor of x equals z). We will be interested in the eigenvectors and eigenvalues of $\mathsf{H}_{k,\ell}$. Since $\mathsf{H}_{k,\ell}$ is a Cayley graph on the Boolean hypercube, its eigenvectors are precisely the characters $\chi_{T_1,\ldots,T_\ell}$.

Lemma 2.10. *If $\dim(\mathsf{Span}(T_1, \ldots, T_\ell)) = r$, then $\chi_{T_1,\ldots,T_\ell}$ is a eigenvector of $\mathsf{H}_{k,\ell}$ with eigenvalue 2^{-r}, i.e.*

$$\mathsf{H}_{k,\ell} \cdot \chi_{T_1,\ldots,T_\ell} = 2^{-r} \cdot \chi_{T_1,\ldots,T_\ell}.$$

Proof. Considering a random choice of $y \in \{0,1\}^k$ and $b = (b_1, \ldots, b_\ell) \in \{0,1\}^\ell$,

$$\begin{aligned}\mathsf{H}_{k,\ell} \cdot \chi_{T_1,\ldots,T_\ell}(x) &= \mathop{\mathbb{E}}_{y,b}[\,\chi_{T_1,\ldots,T_\ell}(x_1 + b_1 y, \ldots, x_\ell + b_\ell(y))] \\ &= \chi_{T_1,\ldots,T_\ell}(x) \cdot \mathop{\mathbb{E}}_{y,b}\left[\chi_{\oplus_{i=1}^\ell b_i T_i}(y)\right].\end{aligned}$$

The expectation over y vanishes if $\oplus_{i=1}^{\ell} b_i T_i \neq 0$ and equals 1 otherwise. Since $\oplus_{i=1}^{\ell} b_i T_i$ is a uniformly random vector in $\mathsf{Span}(T_1, \ldots, T_\ell)$, the probability over the choice of b that $\oplus_{i=1}^{\ell} b_i T_i = 0$ is precisely 2^{-r}. □

Definition 2.11. *(Clearly) any function $F\colon (\{0,1\}^k)^\ell \to \mathbb{R}$ can be written as*

$$F[x_1, \ldots, x_\ell] = \sum_{T_1,\ldots,T_\ell \in \{0,1\}^k} \widehat{F}(T_1, \ldots, T_\ell) \cdot \chi_{T_1,\ldots,T_\ell}(x_1, ..., x_\ell).$$

Its i^{th} level component is defined as its projection onto the eigenspace with eigenvalue 2^{-i}, i.e.

$$F_{=i}[x_1, \ldots, x_\ell] = \sum_{\substack{T_1,\ldots,T_\ell \in \{0,1\}^k \\ \mathsf{dim}(\mathsf{Span}(T_1,\ldots,T_\ell))=i}} \widehat{F}(T_1, \ldots, T_\ell) \cdot \chi_{T_1,\ldots,T_\ell}(x_1, ..., x_\ell).$$

The decomposition $F = \sum_{i=0}^{\ell} F_{=i}$ into "Fourier levels" satisfies Parseval's identity: $\|F\|_2^2 = \sum_{i=0}^{\ell} \|F_{=i}\|_2^2$.

Definition 2.12. *A function $F\colon (\{0,1\}^k)^\ell \to \mathbb{R}$ is called basis-invariant if for every $x_1, \ldots, x_\ell \in \{0,1\}^k$ and an invertible $\ell \times \ell$ matrix M over $\mathbb{F}_2$, we have that*

$$F[x_1, \ldots, x_\ell] = F[M(x_1, \ldots, x_\ell)].$$

Here $M(x_1, \ldots, x_\ell) = (y_1, \ldots, y_\ell)$ such that $y_i = \sum_{j=1}^{\ell} M_{ij}\, x_j$.

In words, a function is basis-invariant if its value is preserved under invertible linear transformation of its arguments. All functions that we deal with in this paper are basis invariant and in particular the indicators of sets $S \subseteq \mathsf{H}_{k,\ell}$ that "arise" from corresponding sets in $\mathsf{Gr}_{k,\ell}$.

2.4 Pseudorandomness implies Low Weight at Low Levels implies Near-Perfect Expansion

Fix a basis-invariant set $S \subseteq \mathsf{H}_{k,\ell} = (\{0,1\}^k)^\ell$. Let $F : (\{0,1\}^k)^\ell \to \{0,1\}$ be its indicator function. Let $\delta = \mu(F) = \|F\|_2^2$ denote its density and let $\|F_{=i}\|_2^2$ be its "weight at the i^{th} Fourier level". We note that the weight at the 0^{th} level is δ^2 and the sum of the weights at all Fourier levels equals δ. Theorem 2.6 requires us to show that if S is pseudorandom, then it has near-prefect expansion. At a high-level, this is accomplished in two steps:

- One shows that a pseudorandom set must have low (say $\leqslant \zeta\delta$) weight at all lower (say up to r) levels.
- One shows that if there is low weight at all lower levels, then the set must have near-perfect expansion ($\geqslant 1 - \zeta(r+1) - 2^{-(r+1)}$).

We include a quick proof of the second step below for the sake of completeness. The main task remains thereafter to prove the first step. Assume therefore that F has weight at most $\zeta\delta$ at each level up to r. Below a random neighbor of x is denoted as $z \sim x$ and the inner product is $\langle F_1, F_2 \rangle = \mathbb{E}_x [F_1(x) F_2(x)]$. We have

$$1 - \Phi(S) = \Pr_{x \in S, z \sim x} [z \in S] = (1/\delta) \cdot \Pr_{x, z \sim x} [x \in S \wedge z \in S] = (1/\delta) \cdot \langle F,\, \mathsf{H}_{k,\ell}\, F \rangle.$$

Using the decomposition $F = \sum_{i=0}^{\ell} F_{=i}$ into mutually orthogonal eigenspaces $F_{=i}$ of eigenvalues 2^{-i}, and that $\delta = \sum_{i=0}^{\ell} \|F_{=i}\|_2^2$, we get that

$$\delta(1 - \Phi(S)) = \sum_{i=0}^{\ell} 2^{-i} \|F_{=i}\|_2^2 \leqslant \sum_{i=0}^{r} \|F_{=i}\|^2 + 2^{-(r+1)} \sum_{i=r+1}^{\ell} \|F_{=i}\|_2^2 \leqslant \zeta\delta(r+1) + \delta\, 2^{-(r+1)}.$$

Dividing by δ gives us $\Phi(S) \geqslant 1 - \zeta(r+1) - 2^{-(r+1)}$ as claimed. To summarize, to prove Theorem 2.6, it suffices to prove (hence this is our main result):

Theorem 2.13. *Let S be a basis-invariant set of vertices in $\mathsf{H}_{k,\ell}$ that has density δ and is (r, ε) pseudo-random. Let $F\colon \mathsf{H}_{k,\ell} = (\{0,1\}^k)^\ell \to \{0,1\}$ be the indicator function of S. Then for any $i = 0, 1, \ldots, r$,*

$$\eta = \|F_{=i}\|_2^2 \leqslant 2^{7r^3+3} \varepsilon^{\frac{1}{4}} \delta.$$

We now summarize the high-level plan to prove Theorem 2.13 as in [4]. The idea is to consider the fourth moment of $F_{=i}$ and prove both a lower bound and an upper bound on it. Specifically, let S be a set that has density δ and is (r, ε) pseudo-random as in the statement of the theorem. Let $0 \leqslant i \leqslant r$ and let $\eta = \|F_{=i}\|_2^2$. The theorem follows by showing that (the expectation is over $x \in (\{0,1\}^k)^\ell$; one cancels η from both sides, moves $2^9\delta^4$ on the right and then takes a fourth root)

$$\frac{\eta^5}{2^9 \cdot \delta^4} \leqslant \mathbb{E}\left[F_{=i}^4\right] \leqslant 2^{25r^3} \eta\varepsilon. \tag{2}$$

2.5 Lower-bounding the Fourth Moment of $F_{=i}$

Lemma 2.14. *Under the condition and notation of Theorem 2.13,* $\mathbb{E}\left[F_{=i}^4\right] \geqslant \frac{\eta^5}{2^9 \cdot \delta^4}$.

Proof. We note the decomposition $F = \sum_{j=0}^{\ell} F_{=j}$ into mutually orthogonal components and that $\|F\|_2^2 = \delta, \|F_{=i}\|_2^2 = \eta$. Hence $\mathbb{E}\left[(F - F_{=i})^2\right] = \delta - \eta$. By Markov's inequality,

$$\Pr\left[(F - F_{=i})^2 \geqslant 1 - \frac{\eta}{2\delta}\right] \leqslant \delta - \frac{\eta}{2}.$$

On the other hand, F is Boolean and $\Pr[F = 1] = \delta$. Thus with probability at least $\frac{\eta}{2}$, both the events below occur:

$$F = 1, \qquad (F - F_{=i})^2 \leqslant 1 - \frac{\eta}{2\delta},$$

in which case it holds that $(1 - F_{=i})^2 \leqslant 1 - \frac{\eta}{2\delta}$ and in turn that $F_{=i} \geqslant \frac{\eta}{4\delta}$. Hence as claimed,

$$\mathbb{E}\left[F_{=i}^4\right] \geqslant \frac{\eta}{2} \cdot \left(\frac{\eta}{4\delta}\right)^4.$$

□

2.6 Upper-bounding the Fourth Moment of $F_{=i}$

To summarize, the task of proving Theorem 2.13 is now reduced to proving the upper bound in Equation (2), i.e. under the condition and notation of Theorem 2.13, to prove that, for $0 \leqslant i \leqslant r$,

$$\mathbb{E}\left[F_{=i}^4\right] \leqslant 2^{25r^3}\eta\varepsilon, \qquad \eta \stackrel{def}{=} \mathbb{E}\left[F_{=i}^2\right].$$

Proving this upper bound is really the main result of this paper. We describe the first step of the proof below, take a lengthy detour in Section 3 to develop the required analytic machinery, and then return to the proof in Section 5. As shown in Section 3, Lemma 3.13, $F_{=i}$ has an alternate characterization (in addition to that in Definition 2.11 and the two characterizations are related): there exists a (unique) function $f_{=i} : (\{0,1\}^k)^i \to \mathbb{R}$ such that for all $x = (x_1, \ldots, x_\ell) \in (\{0,1\}^k)^\ell$,

$$F_{=i}[x] = \frac{1}{\beta_{i,i}} \sum_{M \in \mathcal{M}[i,\ell]} f_{=i}(Mx).$$

Let's explain the notation: here $\mathcal{M}[i,\ell]$ is the set of all $i \times \ell$ matrices over $\mathbb{F}_2$ that have full row-rank i. For $M \in \mathcal{M}[i,\ell]$, $Mx \in (\{0,1\}^k)^i$ is a i-tuple where $(Mx)_j = \sum_{t=1}^{\ell} M_{jt}x_t$. And $\beta_{i,i}$ is a normalizing factor that equals the number of invertible $i \times i$ matrices. To compute (or rather upper bound) $\mathbb{E}\left[F_{=i}^4\right]$, we simply take the sum to the fourth power, expand, and take the expectation over x:

$$\mathbb{E}\left[F_{=i}^4\right] = \frac{1}{\beta_{i,i}^4} \sum_{M_1,M_2,M_3,M_4 \in \mathcal{M}[i,\ell]} \mathop{\mathbb{E}}_{x} [f_{=i}(M_1x)f_{=i}(M_2x)f_{=i}(M_3x)f_{=i}(M_4x)].$$

We partition the sum according to the direct sum of row spaces of $M_1, \ldots, M_4$, that is according to $A = \oplus_{s=1}^4 \mathsf{rowspan}(M_s)$. We note that $A \subseteq \{0,1\}^\ell$ is a subspace and $i \leqslant d = \mathsf{dim}(A) \leqslant 4i$.

$$\mathbb{E}\left[F_{=i}^4\right] = \frac{1}{\beta_{i,i}^4} \sum_{d=i}^{4i} \sum_{A:\mathsf{dim}(A)=d} \sum_{\substack{M_1,M_2,M_3,M_4 \in \mathcal{M}[i,\ell] \\ \oplus_{s=1}^4 \mathsf{rowspan}(M_s)=A}} \mathop{\mathbb{E}}_{x} [f_{=i}(M_1x)f_{=i}(M_2x)f_{=i}(M_3x)f_{=i}(M_4x)].$$

The main task is to upper bound each individual expectation above. A crude upper bound on the sum is taken thereafter. We note that the number of choices for A is at most $2^{d\ell}$ (the number of d-dimensional subspaces of an ℓ-dimensional space), for fixed A, the number of choices for each of M_1, M_2, M_3, M_4 is at most 2^{id}, and $\beta_{i,i} \geqslant 1$. Hence to show the desired upper bound of $2^{25r^3}\eta\varepsilon$ on the entire sum, it is sufficient to show an upper bound of $\frac{2^{7r^3}\eta\varepsilon}{2^{d\ell}}$ on each individual expectation. The main technical result in the paper is therefore:

Theorem 2.15 (Main Technical Theorem). *Let S be a basis-invariant set of vertices in $\mathsf{H}_{k,\ell}$ that is (r,ε) pseudo-random. Let $F\colon H_{k,\ell} = (\{0,1\}^k)^\ell \to \{0,1\}$ be the indicator function of S and $\eta = \|F_{=i}\|_2^2$. Then for any $0 \leqslant i \leqslant r$, $i \leqslant d \leqslant 4i$, $A \subseteq \{0,1\}^\ell$ of dimension d and $M_1, \ldots, M_4 \in \mathcal{M}[i,\ell]$ such that $\oplus_{s=1}^4 \mathsf{rowspan}(M_s) = A$, we have that*

$$\left| \mathop{\mathbb{E}}_{x \in (\{0,1\}^k)^\ell} [f_{=i}(M_1x)f_{=i}(M_2x)f_{=i}(M_3x)f_{=i}(M_4x)] \right| \leqslant 2^{7r^3} \frac{\eta\varepsilon}{2^{d\ell}}. \tag{3}$$

3 Analytic Machinery

In this section, we present the Fourier analytic machinery needed towards our main results. Unfortunately, we are unable to provide extra insight into the statements of various lemmas in addition to what may be inferred per se from their statements (but please do see Section 3.1 for a high-level picture). In terms of which of these lemmas are to be considered central and which ones more auxiliary in nature, we recommend that Lemmas 3.13, 3.19, 3.20 be treated as the key ones, at least in the sense that these will be referred to and used directly in the main proof.

In what follows, $F\colon (\{0,1\}^k)^\ell \to \mathbb{R}$ is a basis-invariant function in the sense of Definition 2.12. Much of what is said applies to all such functions and not necessarily Boolean functions that are indicators of a basis-invariant set $S \subseteq \mathsf{H}_{k,\ell} = (\{0,1\}^k)^\ell$. However, the latter type of functions are the ones that we are mainly interested in, and the reader may assume that F is of this type. We recall Definition 2.11 of the Fourier representation and the decomposition into Fourier levels, $F = \sum_{r=0}^{\ell} F_{=r}$:

$$F[x_1,\ldots,x_\ell] = \sum_{T_1,\ldots,T_\ell \in \{0,1\}^k} \widehat{F}(T_1,\ldots,T_\ell)\cdot \chi_{T_1,\ldots,T_\ell}(x_1,\ldots,x_\ell).$$

$$F_{=r}[x_1,\ldots,x_\ell] = \sum_{\substack{T_1,\ldots,T_\ell \in \{0,1\}^k \\ \mathsf{dim}(\mathsf{Span}(T_1,\ldots,T_\ell))=r}} \widehat{F}(T_1,\ldots,T_\ell)\cdot \chi_{T_1,\ldots,T_\ell}(x_1,\ldots,x_\ell).$$

Lemma 3.1. *$\widehat{F}(T_1,\ldots,T_\ell)$ depends only on $\mathsf{Span}(T_1,\ldots,T_\ell)$.*

Proof. Suppose $\mathsf{dim}(\mathsf{Span}(T_1,\ldots,T_\ell)) = r$ and let $A_1,\ldots,A_r$ be a basis for the span. Then there is a $r\times\ell$ matrix of row-rank r such that $(T_1,\ldots,T_\ell) = M^{\mathsf{Tr}}(A_1,\ldots,A_r)$, where M^{Tr} is the $\ell\times r$ transposed matrix. Moreover, in this case, defining vectors $(y_1,\ldots,y_r)$ such that $(y_1,\ldots,y_r) = M(x_1,\ldots,x_\ell)$,

$$\prod_{j=1}^{\ell} \chi_{T_j}(x_j) = (-1)^{\oplus_{j=1}^{\ell} T_j\cdot x_j} = (-1)^{\oplus_{s=1}^{r} A_s\cdot y_s} = \prod_{s=1}^{r}\chi_{A_s}(y_s).$$

We extend M to a $\ell\times\ell$ invertible matrix M' by appropriately appending $\ell-r$ rows. Let $(y_1,\ldots,y_r,y_{r+1},\ldots,y_\ell) = M'(x_1,\ldots,x_\ell)$. It follows, using basis-invariance of F, that

$$\begin{aligned}
\widehat{F}(T_1,\ldots,T_\ell) &= \mathop{\mathbb{E}}_{x_1,\ldots,x_\ell}\left[F[x_1,\ldots,x_\ell]\prod_{j=1}^{\ell}\chi_{T_j}(x_j)\right] \\
&= \mathop{\mathbb{E}}_{x_1,\ldots,x_\ell}\left[F[x_1,\ldots,x_\ell]\prod_{s=1}^{r}\chi_{A_s}(y_s)\right] \\
&= \mathop{\mathbb{E}}_{y_1,\ldots,y_\ell}\left[F[y_1,\ldots,y_\ell]\prod_{s=1}^{r}\chi_{A_s}(y_s)\right] \\
&= \widehat{F}(A_1,\ldots,A_r,0,\ldots,0).
\end{aligned}$$

□

Thanks to this lemma, we write $\widehat{F}(T_1,\ldots,T_r)$ instead of $\widehat{F}(T_1,\ldots,T_\ell)$ if $\mathsf{dim}(\mathsf{Span}(T_1,\ldots,T_\ell)) = r$ and the first r characters $T_1,\ldots,T_r$ are linearly independent.

3.1 High-level Picture

We recall the goal outlined earlier: to show that for a pseudorandom set $S \subseteq \mathsf{H}_{k,\ell}$, its indictor function F has low Fourier weight at low levels. Clearly, there are two notions of interest here:

- The zoom-in-out densities $\mu_{\mathsf{in}(Q),\mathsf{out}(W)}(S)$.

  S is (r,ε)-pseudorandom if, by definition, all zoom-in-out densities, for $|Q| + \mathsf{codim}(W) \leqslant r$, are at most ε.

- The Fourier level functions $F_{=r}$.

  The Fourier weight at level r is, by definition, $\|F_{=r}\|_2^2$.

And then there is a third notion: as mentioned in Section 2.6, $F_{=r}$ has an alternate characterization: there exists a (unique) function $f_{=r} : (\{0,1\}^k)^r \to \mathbb{R}$ such that for all $x = (x_1, \ldots, x_\ell) \in (\{0,1\}^k)^\ell$,

$$F_{=r}[x] = \frac{1}{\beta_{i,i}} \sum_{M \in \mathcal{M}[i,\ell]} f_{=i}(Mx).$$

The functions $f_{=r}$ will play a crucial role in our analysis. We will avoid giving them a name. While these functions do capture the zoom-in-out densities, unfortunately we do not have a good intuition as to how.

A large part of our Fourier analytic machinery is devoted to relating the three notions, $\mu_{\mathsf{in}(Q),\mathsf{out}(W)}(S)$, $F_{=r}$, $f_{=r}$ to each other. Interestingly (and rather bafflingly), we only work with zoom-in densities $\mu_{\mathsf{in}(Q)}(S)$, and not with the zoom-out densities. The zoom-out densities enter the picture only in an indirect fashion, as Fourier sums of $f_{=r}$ (see Lemmas 3.19, 3.20). The relationship between $f_{=r}$ and the zoom-in densities is somewhat clearer, see Definition 3.5. Especially for $r = 1$, the relationship is immediate: $f_{=1}(x_1)$ is precisely the change in density of the set S after zooming into point x_1. For higher levels r, there is an inclusion-exclusion type formula that relates $f_{=r}$ to densities of the set S after zooming into up to r points.

3.2 Zoom-out Restriction Lemma

In this section, we prove a recursive formula that relates the Fourier coefficients of F to those of restrictions of F to a hyperplane.

Definition 3.2. *Let $F\colon (\{0,1\}^k)^\ell \to \mathbb{R}$ be a function. For a subspace $W \subseteq \{0,1\}^k$, we define the function $F_W\colon W^\ell \to \mathbb{R}$ to be the restriction of F to W^ℓ (referred to as the zoom-out function).*

Definition 3.3. *For a character T, the subspace orthogonal to T is $W_T = \left\{ x \in \{0,1\}^k \;\middle|\; T \cdot x = 0 \right\}$.*

Lemma 3.4. *Let $A, T_1, \ldots, T_r$ be linearly independent characters. Then*

$$\widehat{F}(A, T_1, \ldots, T_r) = \frac{1}{2^\ell - 2^r} \left(\widehat{F}_{W_A}(T_1, \ldots, T_r) - \sum_{\substack{D \subseteq \mathsf{Span}(A,T_1,\ldots,T_r) \\ \dim(D)=r, A \notin D}} \widehat{F}(D) \right).$$

Proof. The computation proceeds as below where in the third step we use the Fourier representation of F and in the fourth step we use the observation that for a character R, the expectation $\mathbb{E}_{y \in W_A}[\chi_R(y)]$ equals 1 when $R = 0$ or when $R = A$ and vanishes otherwise.

$$\begin{aligned}
\widehat{F}_{W_A}(T_1, \ldots, T_r) &= \mathop{\mathbb{E}}_{x_1,\ldots,x_\ell \in W_A} [F_{W_A}(x_1, \ldots, x_\ell) \cdot \chi_{T_1,\ldots,T_r}(x_1, \ldots, x_r)] \\
&= \mathop{\mathbb{E}}_{x_1,\ldots,x_\ell \in W_A} \left[F(x_1, \ldots, x_\ell) \cdot \prod_{i=1}^{r} \chi_{T_i}(x_i) \right] \\
&= \mathop{\mathbb{E}}_{x_1,\ldots,x_\ell \in W_A} \left[\sum_{Q_1,\ldots,Q_\ell} \widehat{F}(Q_1, \ldots, Q_\ell) \prod_{i=1}^{r} \chi_{T_i \oplus Q_i}(x_i) \cdot \prod_{j=r+1}^{\ell} \chi_{Q_j}(x_j) \right] \\
&= \sum_{1 \leqslant i \leqslant r:\ Q_i \in \{T_i, T_i \oplus A\}} \quad \sum_{r+1 \leqslant j \leqslant \ell:\ Q_j \in \{0, A\}} \widehat{F}(Q_1, ..., Q_\ell) \\
&= (2^\ell - 2^r)\widehat{F}(A, T_1, \ldots, T_r) + \sum_{\substack{D \subseteq \mathsf{Span}(A, T_1, \ldots, T_r) \\ \dim(D) = r, A \notin D}} \widehat{F}(D).
\end{aligned}$$

The last equality is justified as follows. There are 2^ℓ terms in the summation which split into two groups:

- In $2^\ell - 2^r$ terms, there is some $j \geqslant r + 1$ such that $Q_j = A$. In this case $\mathsf{Span}(Q_1, \ldots, Q_\ell)$ is same as $\mathsf{Span}(A, T_1, \ldots, T_r)$ and since the Fourier coefficients depend only on this span, $\widehat{F}(Q_1, ..., Q_\ell) = \widehat{F}(A, T_1, \ldots, T_r)$.
- For the remaining 2^r terms, for all $j \geqslant r + 1$, $Q_j = 0$. In this case, $\mathsf{Span}(Q_1, \ldots, Q_\ell)$ is same as $\mathsf{Span}(Q_1, \ldots, Q_r)$ which is an r-dimensional subspace of $\mathsf{Span}(A, T_1, \ldots, T_r)$ that does not contain A. Moreover each subspace of this kind is counted exactly once.

□

3.3 Defining $f_{=r}$ and Relating $F_{=r}$, $f_{=r}$ and Zoom-in Densities

For a (basis-invariant) function $F\colon (\{0,1\}^k)^\ell \to \mathbb{R}$, we have the decomposition $F = \sum_{r=0}^{\ell} F_{=r}$ where

$$F_{=r}[x_1, \ldots, x_\ell] = \sum_{\substack{T_1,\ldots,T_\ell \subseteq [k] \\ \dim(\mathsf{Span}(T_1,\ldots,T_\ell)) = r}} \widehat{F}(T_1, \ldots, T_\ell) \cdot \chi_{T_1,\ldots,T_\ell}(x_1, ..., x_\ell).$$

As mentioned, we will need an alternate formula for $F_{=r} : (\{0,1\}^k)^\ell \to \mathbb{R}$ in terms of related functions $f_{=r} : (\{0,1\}^k)^r \to \mathbb{R}$. Deriving this formula turns out to be rather cumbersome (but quite interesting at the same time). Next few subsections are devoted to this derivation. Sometimes we write $f_{=r,F}$ to make the relation to F explicit. We will use the following notations:

- For integers $1 \leqslant i \leqslant r$, $\mathcal{M}[i, r]$ denotes the set of $i \times r$ matrices over $\mathbb{F}_2$ with (row)-rank i. We have $|\mathcal{M}[i, r]| = \prod_{j=0}^{i-1} (2^r - 2^j) \stackrel{def}{=} \beta_{i,r}$.
- For $r \geqslant 0$, we will pretend that $\beta_{0,r} = 1$ and that there is a single matrix $\{0\}$ in $\mathcal{M}[0, r]$.
- For $x = (x_1, \ldots, x_r)$ and $M \in \mathcal{M}[i, r]$, Mx denotes the tuple $(y_1, \ldots, y_i)$ where $y_j = \sum_{t=1}^{r} M_{jt} x_t$.

Defining $f_{=r}$ in terms of Zoom-in Densities

Definition 3.5. *For $0 \leqslant r \leqslant \ell$, define $f_{=r} \colon (\{0,1\}^k)^r \to \mathbb{R}$ inductively as*

$$f_{=0}(\{0\}) = \mu(F) \stackrel{def}{=} \mathop{\mathbb{E}}_{x_1,\ldots,x_\ell}[F[x_1,\ldots,x_\ell]],$$

$$f_{=r}(x_1,\ldots,x_r) = \mu_{\mathsf{in}(\{x_1,\ldots,x_r\})}(F) - \sum_{d=0}^{r-1} \frac{1}{\beta_{d,d}} \sum_{M \in \mathcal{M}[d,r]} f_{=d}(Mx).$$

We note that

- $\mu_{\mathsf{in}(\{x_1,\ldots,x_r\})}(F) = \mathbb{E}_{z_{r+1},\ldots,z_\ell}[F[x_1,\ldots,x_r,z_{r+1},\ldots,z_\ell]]$ *is the zoom-in density.*
- *The term corresponding to $d = 0$ in the summation above equals $\mu(F)$.*
- *In the case $r = 1$,*
$$f_{=1}(x_1) = \mu_{\mathsf{in}(\{x_1\})}(F) - \mu(F).$$

Lemma 3.6. *The function $f_{=r}$ is basis-invariant, i.e. for every $x = (x_1,\ldots,x_r) \in (\{0,1\}^k)^r$ and $M \in \mathcal{M}[r,r]$, we have that*

$$f_{=r}(x) = f_{=r}(Mx).$$

Proof. By induction on r. For $r = 0, 1$, this is trivial. Let $r \geqslant 2$ and fix $x_1,\ldots,x_r$ and an $r \times r$ invertible matrix M as in the claim. By definition

$$f_{=r}(Mx) = \mu_{\mathsf{in}(Mx)}(F) - \sum_{d=0}^{r-1} \frac{1}{\beta_{d,d}} \sum_{M' \in \mathcal{M}[d,r]} f_{=d}(M'Mx).$$

Observe that for any $d \geqslant 0$, the mapping $M' \to M'M$ is a bijection on $\mathcal{M}[d,r]$. Also observe that $\mu_{\mathsf{in}(Mx)}(F) = \mu_{\mathsf{in}(x)}(F)$, since F is basis invariant. Thus the last expression equals

$$\mu_{\mathsf{in}(x)}(F) - \sum_{d=0}^{r-1} \frac{1}{\beta_{d,d}} \sum_{M' \in \mathcal{M}[d,r]} f_{=d}(M'x) = f_{=r}(x).$$

□

Zoom-in Restriction Lemma

Definition 3.7. *Let $F \colon (\{0,1\}^k)^\ell \to \mathbb{R}$ be a function and let $Q = \{a_1,\ldots,a_j\} \subseteq \{0,1\}^k$ where $j \leqslant \ell - 1$. We define the function $F_Q \colon (\{0,1\}^k)^{\ell-j} \to \mathbb{R}$ (the zoom-in restriction function) by*

$$F_Q[x_1,\ldots,x_{\ell-j}] = F[a_1,\ldots,a_j,x_1,\ldots,x_{\ell-j}].$$

We have the following recursive formula for $f_{=r}$. Here e_1 refers to a vector with the first coordinate 1 and all other coordinates zero.

Lemma 3.8. *Let $F\colon (\{0,1\}^k)^\ell \to \mathbb{R}$ and $r \geqslant 0$. Let $D = (a, x_1, ..., x_r) = (a, x)$. Then*

$$f_{=r+1,F}(D) = f_{=r,F_{\{a\}}}(x) - \frac{1}{\beta_{r,r}} \sum_{\substack{M' \in \mathcal{M}[r,r+1] \\ e_1 \notin \mathsf{rowspan}(M')}} f_{=r,F}(M'D).$$

Proof. The proof is by induction. The case $r = 0$ is trivial, both sides being $\mu_{\mathsf{in}(\{a\})}(F) - \mu(F)$, so assume $r \geqslant 1$. For convenience, we write $f_{=r}$ instead of $f_{=r,F}$, but do write $f_{=r,F_a}$ when it is the zoom-in function that is concerned. From Definition 3.5,

$$f_{=r+1}(D) = \mu_D(F) - \mu(F) - \sum_{j=1}^{r} \frac{1}{\beta_{j,j}} \sum_{M^* \in \mathcal{M}[j,r+1]} f_{=j}(M^*D).$$

Using $D = (a, x),\ \mu_D(F) = \mu_{\mathsf{in}(x)}(F_a)$ and splitting the summation into two parts, we get

$$f_{=r+1}(D) = \mu_{\mathsf{in}(x)}(F_a) - \mu(F) - \sum_{j=1}^{r} \frac{1}{\beta_{j,j}} \left(\sum_{\substack{M^* \in \mathcal{M}[j,r+1] \\ e_1 \in \mathsf{rowspan}(M^*)}} f_{=j}(M^*D) + \sum_{\substack{M^* \in \mathcal{M}[j,r+1] \\ e_1 \notin \mathsf{rowspan}(M^*)}} f_{=j}(M^*D) \right). \tag{4}$$

For fixed j, let's call the two terms above Γ_j and Δ_j respectively. Below, computation of Γ_j results in an "extra" $-\Delta_{j-1}$ term that cancels with the previous Δ-term in a telescoping manner.

$$\begin{aligned} \Gamma_j = \frac{1}{\beta_{j,j}} \sum_{\substack{M^* \in \mathcal{M}[j,r+1] \\ e_1 \in \mathsf{rowspan}(M^*)}} f_{=j}(M^*D) &= \frac{\beta_{j,r+1}}{\beta_{j,j}} \frac{2^j - 1}{2^{r+1} - 1} \cdot \mathop{\mathbb{E}}_{\substack{M^* \in \mathcal{M}[j,r+1] \\ e_1 \in \mathsf{rowspan}(M^*)}} [f_{=j}(M^*D)] \\ &= \frac{\beta_{j-1,r}}{\beta_{j-1,j-1}} \cdot \mathop{\mathbb{E}}_{\substack{M^* \in \mathcal{M}[j,r+1] \\ e_1 \in \mathsf{rowspan}(M^*)}} [f_{=j}(M^*D)], \end{aligned} \tag{5}$$

where we replaced summation by expectation for the sake of convenience with appropriate normalization factor and then used the definition of the β-parameters. The normalization factor is justified by observing that there are $\beta_{j,r+1}$ matrices in $\mathcal{M}[j, r+1]$ and a fraction $\frac{2^j-1}{2^{r+1}-1}$ of them will contain e_1 in their row-span (all non-zero vectors being symmetric in this regard).

Using Lemma 3.9, we see that M^* can be sampled by sampling $M \in \mathcal{M}[j-1, r]$, constructing M' from M, sampling M'', and then letting $M^* = M'' \cdot M'$. Since M'' is invertible and $f_{=j}$ is basis invariant,

$$f_{=j}(M^*D) = f_{=j}(M''M'D) = f_{=j}(M'D) = f_{=j}(M'(a, x)) = f_{=j}(a, Mx).$$

Hence the expectation in (5) is same as $\mathbb{E}_M[f_{=j}(a, Mx)]$. Applying the induction hypothesis (note that Mx is a $(j-1)$-tuple):

$$\mathop{\mathbb{E}}_{M \in \mathcal{M}[j-1,r]} \left[f_{=j-1,F_a}(Mx) - \frac{1}{\beta_{j-1,j-1}} \sum_{\substack{M'' \in \mathcal{M}[j-1,j] \\ e_1 \notin \mathsf{rowspan}(M'')}} f_{=j-1}(M''(a, Mx)) \right]$$

$$= \mathop{\mathbb{E}}_{M \in \mathcal{M}[j-1,r]} [f_{=j-1,F_a}(Mx)] - \frac{\beta_{j-1,j}}{\beta_{j-1,j-1}} \frac{2^j - 2^{j-1}}{2^j - 1} \mathop{\mathbb{E}}_{M \in \mathcal{M}[j-1,r]} \left[\mathop{\mathbb{E}}_{\substack{M'' \in \mathcal{M}[j-1,j] \\ e_1 \notin \mathsf{rowspan}(M'')}} \left[f_{=j-1}(M''(a, Mx)) \right] \right],$$

where the normalization factor is justified as before. Using Lemma 3.10, the distribution of $M^* = M''M'$ here (M' is constructed from M as in the lemma) is uniform over matrices in $\mathcal{M}[j-1, r+1]$ whose row-span does not contain e_1. It is also observed that

$$f_{=j-1}(M''(a, Mx)) = f_{=j-1}(M''M'(a, x)) = f_{=j-1}(M^*D).$$

Using the definition of the β-parameters, the expression can be re-written as

$$\mathop{\mathbb{E}}_{M\in\mathcal{M}[j-1,r]}[f_{=j-1,F_a}(Mx)] - 2^{j-1}\mathop{\mathbb{E}}_{\substack{M^*\in\mathcal{M}[j-1,r+1]\\ e_1\notin\mathsf{rowspan}(M^*)}}[f_{=j-1}(M^*D)].$$

Substituting in (5), we get

$$\begin{aligned}\Gamma_j &= \frac{\beta_{j-1,r}}{\beta_{j-1,j-1}}\mathop{\mathbb{E}}_{M\in\mathcal{M}[j-1,r]}[f_{=j-1,F_a}(Mx)] - \frac{\beta_{j-1,r}}{\beta_{j-1,j-1}}\cdot 2^{j-1}\mathop{\mathbb{E}}_{\substack{M^*\in\mathcal{M}[j-1,r+1]\\ e_1\notin\mathsf{rowspan}(M^*)}}[f_{=j-1}(M^*D)]\\ &= \frac{1}{\beta_{j-1,j-1}}\sum_{M\in\mathcal{M}[j-1,r]} f_{=j-1,F_a}(Mx) - \frac{1}{\beta_{j-1,j-1}}\sum_{\substack{M^*\in\mathcal{M}[j-1,r+1]\\ e_1\notin\mathsf{rowspan}(M^*)}} f_{=j-1}(M^*D)\\ &= \frac{1}{\beta_{j-1,j-1}}\sum_{M\in\mathcal{M}[j-1,r]} f_{=j-1,F_a}(Mx) - \Delta_{j-1}.\end{aligned}$$

Substituting in (4), telescoping, and noting that $\Delta_0 = \mu(F)$ (one can think of Δ_0 as the "extra" term while calculating Γ_1 as above), we get

$$\begin{aligned}f_{=r+1}(D) &= \left(\mu_{\mathsf{in}(x)}(F_a) - \mu(F) + \Delta_0 - \sum_{j=0}^{r-1}\frac{1}{\beta_{j,j}}\sum_{M\in\mathcal{M}[j,r]} f_{=j,F_a}(Mx)\right) - \Delta_r\\ &= f_{=r,F_{\{a\}}}(x) - \Delta_r,\end{aligned}$$

completing the proof. □

Some Auxiliary Lemmas

Lemma 3.9. *A uniformly random matrix M^* in $\mathcal{M}[j, r+1]$ whose row-span contains the vector e_1 can be sampled as:*

- *Pick a uniformly random matrix $M \in \mathcal{M}[j-1, r]$.*
- *Augment M to a matrix $M' \in \mathcal{M}[j, r+1]$ so that M' has top-left corner entry 1, the rest of the entries in the first column and the first row are 0 and deleting the first column and the first row yields M.*
- *Pick a uniformly random matrix $M'' \in \mathcal{M}[j, j]$ and output $M^* = M'' \cdot M'$.*

Proof. Let W be the r-dimensional subspace of $\mathbb{F}_2^{r+1}$ consisting of vectors whose first coordinate is 0. Clearly, a random j-dimensional subspace $L' \subseteq \mathbb{F}_2^{r+1}$ that contains e_1 is obtained by picking a random $(j-1)$-dimensional subspace $L \subseteq W$ and letting $L' = \mathsf{Span}(e_1) \oplus L$. Writing a random basis of L as rows of a matrix yields M. Writing e_1 followed by a random basis of L as rows of a matrix yields M' and its row-span equals L'. Thus it follows that the row-span of M' is a random j-dimensional subspace of $\mathbb{F}_2^{r+1}$ containing e_1. Now pre-multiplying M' by a random invertible matrix M'' yields the matrix M^* whose rows now form a random basis of a random j-dimensional subspace of $\mathbb{F}_2^{r+1}$ containing e_1 as claimed. □

Lemma 3.10. *A uniformly random matrix M^* in $\mathcal{M}[j-1, r+1]$ whose row-span does not contain the vector e_1 can be sampled as (the two incarnations of e_1 in the statement of this lemma are different):*

- *Pick a uniformly random matrix $M \in \mathcal{M}[j-1, r]$.*
- *Augment M to a matrix $M' \in \mathcal{M}[j, r+1]$ so that M' has top-left corner entry 1, the rest of the entries in the first column and the first row are 0 and deleting the first column and the first row yields M.*
- *Pick a matrix M'' that is uniformly distributed over matrices in $\mathcal{M}[j-1, j]$ whose row-span does not contain e_1 and output $M^* = M'' \cdot M'$.*

Proof. Let W, L, L' be as in the proof of the previous lemma. As therein, L' is a random j-dimensional subspace of $\mathbb{F}_2^{r+1}$ that contains e_1 and the row-span of M' equals L' and its rows are $v_1 = e_1$ followed by a basis $v_2, \ldots, v_j$ of L. From Lemma 3.11 below, the rows of $M^* = M'' \cdot M'$ then form a random basis of a random $(j-1)$-dimensional subspace of L' that does not contain $v_1 = e_1$. □

Lemma 3.11. *Let $v_1, \ldots, v_j$ be vectors that are linearly independent (over $\mathbb{F}_2$). Let M'' be a uniformly random matrix in $\mathcal{M}[j-1, j]$ whose row-span does not contain the vector e_1. Let*

$$(w_1, \ldots, w_{j-1}) = M'' \cdot (v_1, \ldots, v_j).$$

Then $(w_1, \ldots, w_{j-1})$ is a random basis of a random $(j-1)$-dimensional subspace of $\mathsf{Span}(v_1, \ldots, v_j)$ that does not contain v_1.

Proof. It is clear that

- Since the rows of M'' are linearly independent, so are $w_1, \ldots, w_{j-1}$.
- The matrix M'' and the tuple $(w_1, \ldots, w_{j-1})$ determine each other.
- $e_1 \notin \mathsf{rowspan}(M'')$ if and only if $v_1 \notin \mathsf{Span}(w_1, \ldots, w_{j-1})$.

Thus there is a one-to-one correspondence between matrices M'' in $\mathcal{M}[j-1, j]$ whose row-span does not contain the vector e_1 and tuples $(w_1, \ldots, w_{j-1})$ that span a $(j-1)$-dimensional subspace of $\mathsf{Span}(v_1, \ldots, v_j)$ that does not contain the vector v_1. □

Relating $F_{=r}$ to Zoom-in Densities

Lemma 3.12. *For every $0 \leqslant r \leqslant \ell$ and $x_1, \ldots, x_\ell \in \{0,1\}^k$,*

$$\left(\sum_{i=0}^{r} \frac{\beta_{i,r}}{\beta_{r,r}} \frac{\beta_{r,\ell}}{\beta_{i,\ell}} F_{=i} \right) [x_1, \ldots, x_\ell] = \frac{1}{\beta_{r,r}} \sum_{\substack{M \in \mathcal{M}[r,\ell] \\ y = Mx}} \mu_{\mathsf{in}((y_1, \ldots, y_r))}(F).$$

Proof. By definition,

$$\begin{aligned} F_{=i}[x_1, \ldots, x_\ell] &= \sum_{\substack{T_1, \ldots, T_\ell \\ \mathsf{dim}(T_1, \ldots, T_\ell) = i}} \widehat{F}(T_1, \ldots, T_\ell) \cdot \chi_{T_1, \ldots, T_\ell}(x_1, \ldots, x_\ell) \\ &= \sum_{\mathsf{dim}(D) = i} \widehat{F}(Q_1, \ldots, Q_i, Q_{i+1}, \ldots, Q_\ell) \sum_{\substack{T_1, \ldots, T_\ell \\ \mathsf{Span}(T_1, \ldots, T_\ell) = D}} \prod_{j=1}^{\ell} \chi_{T_j}(x_j), \end{aligned}$$

where the outer summation is over all i-dimensional subspaces D and for given D, $(Q_1, \ldots, Q_i)$ is a specific ordered basis for it, and $Q_{i+1} = \cdots = Q_\ell = 0$. We used the fact that the Fourier coefficient depends only on the span of its arguments. For given D, consider the inner sum. It is not difficult to see that all ℓ-tuples $(T_1, \ldots, T_\ell)$ such that $\mathsf{Span}(T_1, \ldots, T_\ell) = D$ are obtained precisely as

$$(T_1, \ldots, T_\ell) = M^{\mathsf{Tr}}(Q_1, \ldots, Q_i)$$

where M^{Tr} is a $\ell \times i$ matrix that is a transpose of a matrix $M \in \mathcal{M}[i, \ell]$. Moreover, in that case, defining vectors $(y_1, \ldots, y_i)$ such that $(y_1, \ldots, y_i) = M(x_1, \ldots, x_\ell)$ (which we abbreviate as $y = Mx$)

$$\prod_{j=1}^{\ell} \chi_{T_j}(x_j) = (-1)^{\oplus_{j=1}^{\ell} T_j \cdot x_j} = (-1)^{\oplus_{s=1}^{i} Q_s \cdot y_s} = \prod_{s=1}^{i} \chi_{Q_s}(y_s).$$

Thus we can write

$$F_{=i}[x_1, \ldots, x_\ell] = \sum_{\mathsf{dim}(D)=i} \widehat{F}(Q_1, \ldots, Q_i, Q_{i+1}, \ldots, Q_\ell) \sum_{\substack{M \in \mathcal{M}[i,\ell] \\ y = Mx}} \prod_{j=1}^{i} \chi_{Q_j}(y_j).$$

Using the definition of $\widehat{F}(Q_1, \ldots, Q_i, Q_{i+1} = 0, \ldots, Q_\ell = 0)$ and interchanging the order of summation,

$$\begin{aligned}
F_{=i}[x_1, \ldots, x_\ell] &= \sum_{\mathsf{dim}(D)=i} \mathop{\mathbb{E}}_{z_1, \ldots, z_\ell} \left[F(z_1, \ldots, z_\ell) \prod_{j=1}^{i} \chi_{Q_j}(z_j) \right] \cdot \sum_{\substack{M \in \mathcal{M}[i,\ell] \\ y = Mx}} \prod_{j=1}^{i} \chi_{Q_j}(y_j) \\
&= \sum_{\substack{M \in \mathcal{M}[i,\ell] \\ y = Mx}} \mathop{\mathbb{E}}_{z_1, \ldots, z_\ell} \left[F(z_1, \ldots, z_\ell) \sum_{\mathsf{dim}(D)=i} \prod_{j=1}^{i} \chi_{Q_j}(y_j \oplus z_j) \right] \\
&= \frac{\beta_{i,\ell}}{\beta_{r,\ell}} \sum_{\substack{M \in \mathcal{M}[r,\ell] \\ y = Mx}} \mathop{\mathbb{E}}_{z_1, \ldots, z_\ell} \left[F(z_1, \ldots, z_\ell) \sum_{\mathsf{dim}(D)=i} \prod_{j=1}^{i} \chi_{Q_j}(y_j \oplus z_j) \right].
\end{aligned}$$

Note that in the last line, the summation is over $r \times \ell$ matrices instead of $i \times \ell$ matrices and out of the vectors $(y_1, \ldots, y_r)$, only the first i vectors are "used". The normalization factor takes into account the number of matrices of the two different sizes. For a randomly chosen $r \times r$ invertible matrix M', consider the change of basis $y = M'y'$, $(z_1, \ldots, z_r) = M'(z'_1, \ldots, z'_r)$ and $(A_1, \ldots, A_r) = M'^{\mathsf{Tr}}(Q_1, \ldots, Q_i, Q_{i+1} = 0, \ldots, Q_r = 0)$. By similar reasoning as before

$$\prod_{j=1}^{i} \chi_{Q_j}(y_j \oplus z_j) = \prod_{j=1}^{r} \chi_{Q_j}(y_j \oplus z_j) = \prod_{j=1}^{r} \chi_{A_j}(y'_j \oplus z'_j).$$

Since F is basis-invariant and the distribution of y and y' is the same, we may as well write the above equation as

$$F_{=i}[x_1, \ldots, x_\ell] = \frac{\beta_{i,\ell}}{\beta_{r,\ell}} \sum_{\substack{M \in \mathcal{M}[r,\ell] \\ y = Mx}} \mathop{\mathbb{E}}_{z_1, \ldots, z_\ell} \left[F(z_1, \ldots, z_\ell) \sum_{\mathsf{dim}(D)=i} \mathop{\mathbb{E}}_{M'} \left[\prod_{j=1}^{r} \chi_{A_j}(y_j \oplus z_j) \right] \right].$$

In the above expression, first an i-dimensional subspace D is chosen along with a fixed ordered basis $Q_1, \ldots, Q_i$ and then $(A_1, \ldots, A_r) = M'^{\mathsf{Tr}}(Q_1, \ldots, Q_i, Q_{i+1} = 0, \ldots, Q_r = 0)$ for a random $r \times r$ invertible matrix M'. Up to a factor of $\frac{1}{\beta_{r,r}}$, one can instead consider a summation over all M', and then every tuple $(A_1, \ldots, A_r)$ such that $\mathsf{dim}(A_1, \ldots, A_r) = i$ occurs exactly $\frac{\beta_{r,r}}{\beta_{i,r}}$ times. Hence the above equation can be written as

$$F_{=i}[x_1, \ldots, x_\ell] = \frac{\beta_{i,\ell}}{\beta_{r,\ell}} \frac{1}{\beta_{i,r}} \sum_{\substack{M \in \mathcal{M}[r,\ell] \\ y = Mx}} \mathop{\mathbb{E}}_{z_1,\ldots,z_\ell} \left[F(z_1, \ldots, z_\ell) \sum_{\mathsf{dim}(A_1,\ldots,A_r)=i} \prod_{j=1}^{r} \chi_{A_j}(y_j \oplus z_j) \right].$$

Moving the β-factors to the left hand side and summing over $i = 0, 1, \ldots, r$ counts every r-tuple $(A_1, \ldots, A_r)$ exactly once (irrespective of its dimension). Hence

$$\left(\sum_{i=0}^{r} \beta_{i,r} \frac{\beta_{r,\ell}}{\beta_{i,\ell}} F_{=i} \right) [x_1, \ldots, x_\ell] = \sum_{\substack{M \in \mathcal{M}[r,\ell] \\ y = Mx}} \mathop{\mathbb{E}}_{z_1,\ldots,z_\ell} \left[F(z_1, \ldots, z_\ell) \sum_{A_1,\ldots,A_r} \prod_{j=1}^{r} \chi_{A_j}(y_j \oplus z_j) \right].$$

We observe finally that the inner summation equals 2^{kr} if $z_j = y_j$ for $1 \leqslant j \leqslant r$ and vanishes otherwise. In the former case, we can "fix" $z_j = y_j$ for $1 \leqslant j \leqslant r$ and drop the 2^{kr} factor (since 2^{-kr} is the probability that randomly chosen z_j happens to equal y_j for $1 \leqslant j \leqslant r$). This yields

$$\left(\sum_{i=0}^{r} \beta_{i,r} \frac{\beta_{r,\ell}}{\beta_{i,\ell}} F_{=i} \right) [x_1, \ldots, x_\ell] = \sum_{\substack{M \in \mathcal{M}[r,\ell] \\ y = Mx}} \mathop{\mathbb{E}}_{z_{r+1},\ldots,z_\ell} [F(y_1, \ldots, y_r, z_{r+1}, z_\ell)].$$

The proof of Lemma 3.12 is completed by dividing both sides by $\beta_{r,r}$ and noting that the expectation is precisely $\mu_{\mathsf{in}(y=Mx)}(F)$. □

Relating $F_{=r}$ and $f_{=r}$

Lemma 3.13. *For every $0 \leqslant r \leqslant \ell$ and $x_1, \ldots, x_\ell \in \{0,1\}^k$,*

$$F_{=r}[x_1, \ldots, x_\ell] = \frac{1}{\beta_{r,r}} \sum_{M \in \mathcal{M}[r,\ell]} f_{=r}(Mx).$$

Proof. The proof is by induction on r. The case $r = 0$ is trivial (both sides equal $\mu(F)$). Otherwise, using Lemma 3.12 we have

$$\left(\sum_{i=0}^{r} \frac{\beta_{i,r}}{\beta_{r,r}} \frac{\beta_{r,\ell}}{\beta_{i,\ell}} F_{=i} \right) [x_1, \ldots, x_\ell] = \frac{1}{\beta_{r,r}} \sum_{M \in \mathcal{M}[r,\ell]} \mu_{\mathsf{in}((Mx))}(F).$$

We note that the coefficient of $F_{=r}$ on the left side is 1. Therefore we get

$$F_{=r}[x_1, \ldots, x_\ell] = \frac{1}{\beta_{r,r}} \sum_{M \in \mathcal{M}[r,\ell]} \mu_{\mathsf{in}((Mx))}(F) - \left(\sum_{i=0}^{r-1} \frac{\beta_{i,r}}{\beta_{r,r}} \frac{\beta_{r,\ell}}{\beta_{i,\ell}} F_{=i} \right) [x_1, \ldots, x_\ell]. \tag{6}$$

Using the induction hypothesis, we get that

$$\left(\sum_{i=0}^{r-1}\frac{\beta_{i,r}}{\beta_{r,r}}\frac{\beta_{r,\ell}}{\beta_{i,\ell}}F_{=i}\right)[x_1,\ldots,x_\ell] = \sum_{i=0}^{r-1}\frac{\beta_{i,r}}{\beta_{r,r}}\frac{\beta_{r,\ell}}{\beta_{i,\ell}}\frac{1}{\beta_{i,i}}\sum_{M\in\mathcal{M}[i,\ell]} f_{=i}(Mx)$$
$$= \frac{1}{\beta_{r,r}}\sum_{A\in\mathcal{M}[r,\ell]}\sum_{i=0}^{r-1}\frac{1}{\beta_{i,i}}\sum_{Q\in\mathcal{M}[i,r]} f_{=i}(QAx). \tag{7}$$

The last equality follows by observing that a full row-rank $i\times\ell$ matrix M can be obtained as a product of full row-rank $i\times r$ and $r\times\ell$ matrices Q and A respectively (in uniform manner). In both summations, each M is counted exactly $\frac{\beta_{i,r}\beta_{r,\ell}}{\beta_{i,\ell}}$ times. Substituting (7) into (6) yields

$$F_{=r}[x_1,\ldots,x_\ell] = \frac{1}{\beta_{r,r}}\sum_{M\in\mathcal{M}[r,\ell]}\mu_{\mathsf{in}((Mx))}(F) - \frac{1}{\beta_{r,r}}\sum_{M\in\mathcal{M}[r,\ell]}\sum_{i=0}^{r-1}\frac{1}{\beta_{i,i}}\sum_{Q\in\mathcal{M}[i,r]} f_{=i}(QMx)$$
$$= \frac{1}{\beta_{r,r}}\sum_{M\in\mathcal{M}[r,\ell]} f_{=r}(Mx),$$

where in the last equality we combined the two sums over M, and used Definition 3.5 of $f_{=r}$, thereby finishing the inductive step. □

3.4 Relating $F_{=r}$ and $f_{=r}$ in the Fourier Domain

Lemma 3.14. *Let $0\leqslant j\leqslant r-1$ and let $a_1,\ldots,a_j\in\{0,1\}^k$. Then*

$$\mathop{\mathbb{E}}_{y_{j+1},\ldots,y_r\in\{0,1\}^k}[f_{=r}(a_1,\ldots,a_j,y_{j+1},\ldots,y_r)] = 0.$$

Proof. The proof is by double induction, first by induction on r with $j=0$, and then by induction on j (as long as $j\leqslant r-1$). So assume first that $j=0$. In the case $r=1$,

$$\mathop{\mathbb{E}}_{y_1}[f_{=1}(y_1)] = \mathop{\mathbb{E}}_{y_1}\left[\mu_{\mathsf{in}(y_1)}(F)-\mu(F)\right] = 0.$$

Now assume $j=0$ and $r\geqslant 2$.

$$\mathop{\mathbb{E}}_{y_1,\ldots,y_r}[f_{=r}(y_1,\ldots,y_r)] = \mathop{\mathbb{E}}_{y_1,\ldots,y_r}\left[\mu_{\mathsf{in}(y_1,\ldots,y_r)}(F) - \sum_{i=0}^{r-1}\frac{1}{\beta_{i,i}}\sum_{M\in\mathcal{M}[i,r]} f_{=i}(My)\right]$$
$$= -\sum_{i=1}^{r-1}\frac{1}{\beta_{i,i}}\sum_{M\in\mathcal{M}[i,r]}\mathop{\mathbb{E}}_{y_1,\ldots,y_r}[f_{=i}(My)],$$

where we used the fact that in the summation, the term with index $i=0$ is $\mu(F)$ and

$$\mathop{\mathbb{E}}_{y_1,\ldots,y_r}\left[\mu_{\mathsf{in}(y_1,\ldots,y_r)}(F)\right] = \mu(F)$$

as well. We note that for any $1\leqslant i\leqslant r-1$ and $M\in\mathcal{M}[i,r]$, the distribution of My is uniform over $(\{0,1\}^k)^i$ and hence by the induction hypothesis

$$\mathop{\mathbb{E}}_{y_1,\ldots,y_r}[f_{=i}(My)] = 0,$$

proving the inductive step. We now know that the claim is true for $j = 0$ and all $r \geqslant 1$. In the following, we consider the case $1 \leqslant j \leqslant r-1$ and "reduce" it to the case $j-1$. Using Lemma 3.8 we see

$$\begin{aligned}
&\mathop{\mathbb{E}}_{y_{j+1},\ldots,y_r}[f_{=r}(a_1,\ldots,a_j,y_{j+1},\ldots,y_r)] \\
&= \mathop{\mathbb{E}}_{y_{j+1},\ldots,y_r}\left[f_{=r-1,F_{\{a_1\}}}(a_2,\ldots,a_j,y_{j+1},\ldots,y_r) - \frac{1}{\beta_{r-1,r-1}}\sum_{\substack{M\in\mathcal{M}[r-1,r]\\ e_1\notin\mathsf{rowspan}(M)}} f_{=r-1}(M(a,y))\right]. \qquad (8)
\end{aligned}$$

The expectation of the first term vanishes by induction hypothesis. For the second term, fix any matrix M therein. Since $f_{=r-1}$ is basis-invariant, we can rewrite the rows of M as long as the row-span is preserved. By Lemma 3.15 below, as may assume that M is semi-diagonal. Hence

$$M(a,y) = M(a_1,\ldots,a_j,y_{j+1},\ldots,y_r) = (a'_2,\ldots,a'_j,y'_{j+1},\ldots,y'_r),$$

where $a'_i = a_i$ or $a_i + a_1$ (hence fixed) and similarly, $y'_i = y_i$ or $y_i + a_1$ (hence distributed same as y_i). By induction hypothesis

$$\mathop{\mathbb{E}}_{y_{j+1},\ldots,y_r}[f_{=r-1}(M(a,y))] = \mathop{\mathbb{E}}_{y'_{j+1},\ldots,y'_r}\left[f_{=r-1}(a'_2,\ldots,a'_j,y'_{j+1},\ldots,y'_r)\right] = 0,$$

completing the proof. □

Lemma 3.15. *A $(r-1)\times r$ matrix is called semi-diagonal if deleting the first column gives a square matrix that is diagonal and has ones on the diagonal. Then for any matrix $M \in \mathcal{M}[r-1,r]$ such that $e_1 \notin \mathsf{rowspan}(M)$, there is a semi-diagonal matrix $M' \in \mathcal{M}[r-1,r]$ with the same row-span.*

Proof. Let D be the row-span of M and for $2 \leqslant s \leqslant r$, let $W_s \subseteq \{0,1\}^r$ be the subspace of vectors with the last $r-s$ coordinates zero. Since $\mathsf{dim}(D) = r-1, \mathsf{dim}(W_s) = s, e_1 \in W_s \setminus D$, it must be the case that $\mathsf{dim}(D \cap W_s) = s-1$. Thus $\{D \cap W_s\}_{s=2}^r$ is an "increasing" sequence of subspaces that finally equals D. Hence a basis for D can be chosen so that its successive members are in $W_2 \setminus \{e_1\}, W_3 \setminus W_2, \ldots, W_r \setminus W_{r-1}$ respectively. Moreover, in this process, when we choose a vector $v_s \in W_s \setminus W_{s-1}$, the s^{th} coordinate of v equals one, and we can zero-out its coordinates $2,\ldots,s-1$ by adding to it $v_2,\ldots,v_{s-1}$ if necessary. □

We now consider the Fourier representation of $f_{=r} : (\{0,1\}^k)^r \to \mathbb{R}$:

$$f_{=r}(x_1,\ldots,x_r) = \sum_{T_1,\ldots,T_r} \widehat{f}_{=r}(T_1,\ldots,T_r)\ \chi_{T_1}(x_1)\cdots\chi_{T_r}(x_r).$$

Lemma 3.16. *$\widehat{f}_{=r}(T_1,\ldots,T_r)$ depends only on $\mathsf{Span}(T_1,\ldots,T_r)$.*

Proof. Follows from the basis-invariance of $f_{=r}$ (Lemma 3.6) and a proof identical to that of Lemma 3.1. □

Lemma 3.17. *Suppose $T_1,\ldots,T_r$ are linearly dependent. Then $\widehat{f}_{=r}(T_1,\ldots,T_r) = 0$.*

Proof. Suppose w.l.o.g. that T_r depends on $T_1, \dots, T_{r-1}$. By Lemma 3.16,

$$\begin{aligned}\widehat{f}_{=r}(T_1, \dots, T_r) &= \widehat{f}_{=r}(T_1, \dots, T_{r-1}, 0) \\ &= \mathop{\mathbb{E}}_{y_1,\dots,y_r}\left[f_{=r}(y_1, \dots, y_r)\prod_{j=1}^{r-1}\chi_{T_j}(y_j)\right] \\ &= \mathop{\mathbb{E}}_{y_1,\dots,y_{r-1}}\left[\mathop{\mathbb{E}}_{y_r}\left[f_{=r}(y_1, \dots, y_{r-1}, y_r)\right]\prod_{j=1}^{r-1}\chi_{T_j}(y_j)\right],\end{aligned}$$

and the inner expectation vanishes according to Lemma 3.14. □

Therefore, we may write

$$f_{=r}(y_1, \dots, y_r) = \sum_{\substack{T_1,\dots,T_r \\ \mathsf{dim}(T_1,\dots,T_r)=r}} \widehat{f}_{=r}(T_1, \dots, T_r)\chi_{T_1}(y_1)\cdots\chi_{T_r}(y_r). \tag{9}$$

Lemma 3.18. *Let $0 \leqslant r \leqslant \ell$ and let $T_1, \dots, T_r$ be characters such that $\mathsf{dim}(T_1, \dots, T_r) = r$. Then*

$$\widehat{f}_{=r}(T_1, \dots, T_r) = \widehat{F}(T_1, \dots, T_r).$$

Proof. For any $x_1, \dots, x_\ell \in \{0,1\}^k$, by Lemma 3.13,

$$\begin{aligned}F_{=r}[x_1, \dots, x_\ell] &= \frac{1}{\beta_{r,r}}\sum_{M\in\mathcal{M}[r,\ell]} f_{=r}(Mx) \\ &= \frac{1}{\beta_{r,r}}\sum_{M\in\mathcal{M}[r,\ell]}\sum_{\substack{T_1,\dots,T_r \\ \mathsf{dim}(T_1,\dots,T_r)=r}} \widehat{f}_{=r}(T_1, \dots, T_r)\prod_{j=1}^{r}\chi_{T_j}((Mx)_j) \\ &= \frac{1}{\beta_{r,r}}\sum_{\substack{T_1,\dots,T_r \\ \mathsf{dim}(T_1,\dots,T_r)=r}} \widehat{f}_{=r}(T_1, \dots, T_r)\sum_{M\in\mathcal{M}[r,\ell]}\prod_{j=1}^{r}\chi_{T_j}((Mx)_j).\end{aligned}$$

As we have done previously, if $T = (T_1, \dots, T_r)$ and M^{Tr} is the transposed matrix, we have

$$\prod_{j=1}^{r}\chi_{T_j}((Mx)_j) = \prod_{i=1}^{\ell}\chi_{(M^{\mathsf{Tr}}T)_i}(x_i).$$

Hence,

$$\begin{aligned}F_{=r}[x_1, \dots, x_\ell] &= \frac{1}{\beta_{r,r}}\sum_{\substack{T_1,\dots,T_r \\ \mathsf{dim}(T_1,\dots,T_r)=r}} \widehat{f}_{=r}(T_1, \dots, T_r)\sum_{M\in\mathcal{M}[r,\ell]}\prod_{i=1}^{\ell}\chi_{(M^{\mathsf{Tr}}T)_i}(x_i) \\ &= \sum_{\substack{P_1,\dots,P_\ell \\ \mathsf{dim}(P_1,\dots,P_\ell)=r}} \widehat{f}_{=r}(\mathsf{basis}(P_1, \dots, P_\ell))\prod_{i=1}^{\ell}\chi_{P_i}(x_i),\end{aligned}$$

where it is easily checked that each tuple $(P_1, \ldots, P_\ell)$ with dimension of the span r is counted exactly once. On the other hand, by definition

$$F_{=r}[x_1, \ldots, x_\ell] = \sum_{\substack{P_1,\ldots,P_\ell \\ \mathsf{dim}(P_1,\ldots,P_\ell)=r}} \widehat{F}(\mathsf{basis}(P_1, \ldots, P_\ell)) \prod_{i=1}^{\ell} \chi_{P_i}(x_i).$$

By uniqueness of Fourier representation, we conclude the assertion of the lemma. □

3.5 Bounding Restricted Fourier Sums of $f_{=r}$

Lemma 3.19. *Let S be a (basis invariant) set of vertices in $\mathsf{H}_{k,\ell}$ that is (r, ε) pseudo-random. Let $0 \leqslant j \leqslant r \leqslant \frac{\ell}{2}$ and $F\colon (\{0,1\}^k)^\ell \to \{0,1\}$ be the indicator function of S. Then for any characters $A_1, \ldots, A_j$,*

$$\sum_{T_{j+1},\ldots,T_r} \widehat{f}^2_{=r}(A_1, \ldots, A_j, T_{j+1}, \ldots T_r) \leqslant \frac{2^{4r^2}}{2^{(r+j)\ell}}\, \varepsilon.$$

Proof. We will prove an upper bound of $\frac{C_r}{2^{(r+j)\ell}}\, \varepsilon$ with $C_r = 2^{4r^2}$. We note first that a (r, ε)-pseudorandom set automatically has density at most ε and hence $\|F\|_2^2 \leqslant \varepsilon$. For $j = r = 0$, the upper bound clearly holds with $C_0 = 1$, so we assume $r \geqslant 1$. The proof is by induction on j. When $j = 0$, we have (note that non-zero Fourier coefficients of $f_{=r}$ have linearly independent arguments)

$$\begin{aligned}\sum_{\substack{T_1,\ldots,T_r \\ \mathsf{dim}(T_1,\ldots,T_r)=r}} \widehat{f}^2_{=r}(T_1, \ldots T_r) &= \sum_{\substack{T_1,\ldots,T_r \\ \mathsf{dim}(T_1,\ldots,T_r)=r}} \widehat{F}^2(T_1, \ldots T_r) \\ &= \frac{\beta_{r,r}}{\beta_{r,\ell}} \sum_{\substack{Q_1,\ldots,Q_\ell \\ \mathsf{dim}(Q_1,\ldots,Q_\ell)=r}} \widehat{F}^2(Q_1, \ldots Q_\ell),\end{aligned}$$

since for a fixed r-dimensional span of the arguments, there are $\beta_{r,r}$ terms $(T_1, \ldots, T_r)$ in the first summation and $\beta_{r,\ell}$ terms $(Q_1, \ldots, Q_\ell)$ in the second summation. The expression now equals and is then upper bounded as ($C_r = 2^{4r^2}$),

$$\frac{\beta_{r,r}}{\beta_{r,\ell}} \|F_{=r}\|_2^2 \leqslant \frac{2^{r^2}}{\frac{1}{2} \cdot 2^{r\ell}} \|F\|_2^2 \leqslant \frac{2^{r^2+1}}{2^{r\ell}} \varepsilon \leqslant \frac{C_r}{2^{r\ell}} \varepsilon.$$

Now assume $j \geqslant 1$. By Lemma 3.18 and 3.4, for any $\mathsf{dim}(A_1, \ldots, A_j, T_{j+1}, \ldots, T_r) = r$,

$$\begin{aligned}&\widehat{f}^2_{=r}(A_1, \ldots, A_j, T_{j+1}, \ldots T_r) = \widehat{F}^2(A_1, \ldots, A_j, T_{j+1}, \ldots T_r) \\ &= \frac{1}{(2^\ell - 2^{r-1})^2} \left(\widehat{F}_{W_{A_1}}(A_2, \ldots, A_j, T_{j+1}, \ldots, T_r) - \sum_{\substack{D \subseteq \mathsf{Span}(A_1,\ldots,A_j,T_{j+1},\ldots T_r) \\ \mathsf{dim}(D)=r-1, A_1 \notin D}} \widehat{F}(D) \right)^2 \\ &\leqslant \frac{4 \cdot 2^r}{2^{2\ell}} \left(\widehat{F}^2_{W_{A_1}}(A_2, \ldots, A_j, T_{j+1}, \ldots, T_r) + \sum_{\substack{D \subseteq \mathsf{Span}(A_1,\ldots,A_j,T_{j+1},\ldots T_r) \\ \mathsf{dim}(D)=r-1, A_1 \notin D}} \widehat{F}^2(D) \right), \qquad (10)\end{aligned}$$

the last inequality is by Cauchy-Schwarz (there are 2^{r-1} choices for D). Summing over $T_{j+1}, \ldots, T_r$, the first term sums up to at most

$$\begin{aligned}
&\sum_{\substack{T_{j+1},\ldots,T_r \\ \mathsf{dim}(A_2,\ldots,A_j,T_{j+1},\ldots,T_r)=r-1}} \widehat{F}^2_{W_{A_1}}(A_2,\ldots,A_j,T_{j+1},\ldots,T_r) \\
&= \sum_{\substack{T_{j+1},\ldots,T_r \\ \mathsf{dim}(A_2,\ldots,A_j,T_{j+1},\ldots,T_r)=r-1}} \widehat{f}^2_{=r-1,F_{W_{A_1}}}(A_2,\ldots,A_j,T_{j+1},\ldots T_r) \\
&\leqslant \frac{C_{r-1}}{2^{(r-1+j-1)\ell}}\,\varepsilon,
\end{aligned}$$

using the induction hypothesis and since the (r,ε) pseudorandomness of F implies $(r-1,\varepsilon)$ pseudorandomness of $F_{W_{A_1}}$. For the second term, consider any $D \subseteq \mathsf{Span}(A_1,\ldots,A_j,T_{j+1},\ldots T_r)$ of dimension $r-1$ not containing A_1. By Lemma 3.15, we may assume that D has basis

$$D = \mathsf{Span}(A'_2,\ldots,A'_j,\ T'_{j+1},\ \ldots,\ T'_r)$$

where $A'_i = A_i + b_i \cdot A_1$ and $T'_i = T_i + b_i \cdot A_1$ for some $b = (b_2,\ldots,b_r)$. In the following calculation, b is thought of as fixed, determining D. Summing over all $T_{j+1},\ldots,T_r$,

$$\begin{aligned}
\sum_{\substack{T_{j+1},\ldots,T_r \\ \mathsf{dim}(A_1,\ldots,A_j,T_{j+1},\ldots,T_r)=r}} \widehat{F}^2(D) &\leqslant \sum_{\substack{T'_{j+1},\ldots,T'_r \\ \mathsf{dim}(A'_2,\ldots,A'_j,T'_{j+1},\ldots,T'_r)=r-1}} \widehat{F}^2(A'_2,\ldots,A'_j,T'_{j+1},\ldots,T'_r) \\
&= \sum_{\substack{T'_{j+1},\ldots,T'_r \\ \mathsf{dim}(A'_2,\ldots,A'_j,T'_{j+1},\ldots,T'_r)=r-1}} \widehat{f}^2_{=r-1}(A'_2,\ldots,A'_j,T'_{j+1},\ldots,T'_r) \\
&\leqslant \frac{C_{r-1}}{2^{(r-1+j-1)\ell}}\,\varepsilon,
\end{aligned}$$

using the induction hypothesis. We note that there are 2^{r-1} choices for b (or equivalently D). Combining both the upper bounds, gets us an upper bound of

$$\left(\frac{\varepsilon}{2^{(r+j)\ell}}\right) \cdot (4 \cdot 2^r) \cdot (1 + 2^{r-1}) \cdot C_{r-1}.$$

This is upper bounded by $\frac{C_r}{2^{(r+j)\ell}}\varepsilon$ provided $C_r \geqslant 2^{2r+2}C_{r-1}$ (and $C_0 = 1$). Letting $C_r = 2^{4r^2}$ proves the lemma. □

Lemma 3.20. *Let S be a (basis invariant) set of vertices in $\mathsf{H}_{k,\ell}$ that is (r,ε) pseudo-random. Let s,t,p,q be non-negative integers such that $s+t+p+q = r \leqslant \frac{\ell}{2}$. Let $F\colon (\{0,1\}^k)^\ell \to \{0,1\}$ be the indicator function of S. Let $a_1,\ldots,a_s$ be points and $A_1,\ldots,A_p$ be characters. Define the restriction*

$$g_{a_1,\ldots,a_s,x_1,\ldots,x_t}(y_1,\ldots,y_p,z_1,\ldots,z_q) = f_{=r}(a_1,\ldots,a_s,x_1,\ldots,x_t,y_1,\ldots,y_p,z_1,\ldots,z_q).$$

Then

$$\mathop{\mathbb{E}}_{x_1,\ldots,x_t\in\{0,1\}^k}\left[\sum_{T_1,\ldots,T_q} \widehat{g}^2_{a_1,\ldots,a_s,x_1,\ldots,x_t}(A_1,\ldots,A_p,T_1,\ldots T_q)\right] \leqslant \frac{2^{2sr^2+4r^2}}{2^{(t+2p+q)\ell}}\,\varepsilon \leqslant \frac{2^{6r^3}}{2^{(t+2p+q)\ell}}\,\varepsilon.$$

Proof. We will prove an upper bound of $\frac{C_{s,r}}{2^{(t+2p+q)\ell}}\,\varepsilon$ where $C_{s,r} = 2^{2sr^2+4r^2}$. The proof is by induction on s. First let us consider the case $s = 0$. The expectation to be upper-bounded equals (we denote by x, y, z the respective tuples of variables)

$$\mathop{\mathbb{E}}_{x}\left[\sum_{T_1,\ldots,T_q}\left(\mathop{\mathbb{E}}_{y,z}\left[f_{=r}(x,y,z)\prod_{i=1}^{p}\chi_{A_i}(y_i)\prod_{i=1}^{q}\chi_{T_i}(z_i)\right]\right)^2\right]. \tag{11}$$

Consider the expectation inside. Expanding $f_{=r}$ into Fourier, it equals

$$\sum_{\substack{Q_1,\ldots,Q_t\\ R_1,\ldots,R_p,S_1,\ldots,S_q}}\mathop{\mathbb{E}}_{y,z}\left[\widehat{f}_{=r}(Q,R,S)\prod_{i=1}^{t}\chi_{Q_i}(x_i)\prod_{i=1}^{p}\chi_{A_i\oplus R_i}(y_i)\prod_{i=1}^{q}\chi_{T_i\oplus S_i}(z_i)\right]$$

$$=\sum_{Q_1,\ldots,Q_t}\widehat{f}_{=r}(Q,A,T)\prod_{i=1}^{t}\chi_{Q_i}(x_i).$$

Squaring this, taking the expectation over x, and then summing over $T_1,\ldots,T_q$ shows that (11) equals,

$$\sum_{Q_1,\ldots,Q_t,T_1,\ldots,T_q}\widehat{f}^2_{=r}(Q_1,\ldots,Q_t,A_1,\ldots,A_p,T_1,\ldots,T_q),$$

which is upper bounded by $\frac{2^{4r^2}}{2^{(t+2p+q)\ell}}\,\varepsilon$ by Lemma 3.19. Now consider the case $s \geqslant 1$. As before, the expectation to be upper-bounded equals (with an additional argument $a = (a_1,\ldots,a_s)$)

$$\mathop{\mathbb{E}}_{x}\left[\sum_{T_1,\ldots,T_q}\left(\mathop{\mathbb{E}}_{y,z}\left[f_{=r}(a,x,y,z)\prod_{i=1}^{p}\chi_{A_i}(y_i)\prod_{i=1}^{q}\chi_{T_i}(z_i)\right]\right)^2\right]. \tag{12}$$

Applying Lemma 3.8 we get

$$\begin{aligned}f_{=r}(a_1,\ldots,a_s,x,y,z) &= f_{=r-1,F_{a_1}}(a_2,\ldots,a_s,x,y,z)\\ &\quad+\frac{1}{\beta_{r-1,r-1}}\sum_{\substack{M\in\mathcal{M}[r-1,r]\\ e_1\notin\mathsf{rowspan}(M)}}f_{=r-1,F_{a_1}}(M(a,x,y,z)).\end{aligned}$$

We take expectation over y, z. For the first term, we have

$$\mathop{\mathbb{E}}_{y,z}\left[f_{=r-1,F_{a_1}}(a_2,\ldots,a_s,x,y,z)\prod_{i=1}^{p}\chi_{A_i}(y_i)\prod_{i=1}^{q}\chi_{T_i}(z_i)\right]=\widehat{h}_{a_2,\ldots,a_s,x}(A_1,\ldots,A_p,T_1,\ldots,T_q), \tag{13}$$

where $h_{a_2,\ldots,a_s,x}$ is the restriction of the function $f_{=r-1,F_{a_1}}$ in a manner similar to g is the restriction of $f_{=r}$. For the second term, let $M \in \mathcal{M}[r-1,r]$ whose row-span does not not contain e_1. By Lemma 3.15, we may assume that M is semi-diagonal and then

$$M(a,x,y,z) = (a' = (a'_2,\ldots,a'_s), x', y', z'),$$

where each new coordinate is same as earlier except possibly adding a_1. Hence

$$\mathop{\mathbb{E}}_{y,z}\left[f_{=r-1}(M(a,x,y,z))\prod_{i=1}^{p}\chi_{A_i}(y_i)\prod_{i=1}^{q}\chi_{T_i}(z_i)\right]=\mathsf{sign}\cdot\widehat{h}_{a',x'}(A_1,\ldots,A_p,T_1,\ldots,T_q) \tag{14}$$

for a $\mathsf{sign} \in \{-1, 1\}$ (which takes into account the possible additions of a_1 to get the new coordinates y', z'). Towards upper-bounding (12), we can now sum up the absolute values of (13) and (14) (the latter summed over $\beta_{r-1,r}$ matrices along with the leading coefficient $\frac{1}{\beta_{r-1}r-1}$), square the sum, upper bound it by Cauchy-Schwartz, and finally take the outer summation over $T = (T_1, \ldots, T_q)$ and expectation over x. We end up with an overall upper bound

$$\left(1 + \frac{\beta_{r-1,r}}{\beta^2_{r-1,r-1}}\right)\left(\mathop{\mathbb{E}}_{x}\left[\sum_{T} \widehat{h}^2_{a_2,\ldots,a_s,x}(A,T)\right] + \sum_{M \in \mathcal{M}[r-1,r]} \mathop{\mathbb{E}}_{x}\left[\sum_{T} \widehat{h}^2_{a',x'}(A,T)\right]\right).$$

We may now apply the induction hypothesis since sequences $(a_2, \ldots, a_s)$ and a' have length $s-1$, x' is distributed the same as x, and furthermore, F_{a_1} is $(r-1, \varepsilon)$ pseudo-random. Thus we get an upper bound of

$$\frac{\varepsilon}{2^{(t+2p+q)\ell}} \cdot C_{s-1,r-1} \cdot \left(1 + \frac{\beta_{r-1,r}}{\beta^2_{r-1,r-1}}\right)(1 + \beta_{r-1,r})).$$

Using very crude estimates $\beta_{r-1,r-1} \geqslant 1$ and $\beta_{r-1,r} \leqslant 2^{r^2} - 1$, we upper bound by $\frac{\varepsilon}{2^{(t+2p+q)\ell}} \cdot C_{s,r}$. It suffices to have $C_{s,r} \geqslant 2^{2r^2} C_{s-1,r-1}$ and $C_{0,r} = 2^{4r^2}$, i.e. $C_{s,r} = 2^{2sr^2+4r^2}$. □

4 Pair-wise and Three-wise Correlations of $f_{=i}$

The rest of the paper is devoted to the proof of our main technical result, Theorem 2.15, that upper bounds the four-wise correlations of $f_{=i}$. It is natural and instructive to first understand pair-wise and three-wise correlations of $f_{=i}$.

4.1 Pairwise Correlations

Studying pairwise correlations is simple. There are two cases depending on whether $\mathsf{rowspan}(M_1), \mathsf{rowspan}(M_2)$ are distinct or the same. In the latter case, due to basis-invariance of $f_{=i}$, we may assume that the two matrices are the same.

Lemma 4.1. *Let $M_1, M_2 \in \mathcal{M}[i, \ell]$. If $\mathsf{rowspan}(M_1) \neq \mathsf{rowspan}(M_2)$, then*

$$\mathop{\mathbb{E}}_{x_1,\ldots,x_\ell \in \{0,1\}^k}[f_{=i}(M_1 x) f_{=i}(M_2 x)] = 0.$$

Proof. It is clearly possible to choose linearly independent vectors $v_1, \ldots, v_s, u_1, \ldots, u_{i-s}, w_1, \ldots, w_{i-s}$ in $\{0,1\}^\ell$ such that

- $(v_1, \ldots, v_s)$ is a basis for $\mathsf{rowspan}(M_1) \cap \mathsf{rowspan}(M_2)$,
- $(v_1, \ldots, v_s, u_1, \ldots, u_{i-s})$ is a basis for $\mathsf{rowspan}(M_1)$, and
- $(v_1, \ldots, v_s, w_1, \ldots, w_{i-s})$ is a basis for $\mathsf{rowspan}(M_2)$.

By the assumption $i - s \geqslant 1$. By the basis-invariance of $f_{=i}$, we can assume that in the last two items, the respective sets are in fact the rows of the two matrices. For a row-vector $a \in \{0,1\}^\ell$ and $x = (x_1, \ldots, x_\ell)$, let us denote $a' = \langle a, x \rangle = \sum_{j=1}^{\ell} a_j x_j$. Let us define

$$v'_j = \langle v_j, x \rangle, \qquad u'_j = \langle u_j, x \rangle, \qquad w'_j = \langle w_j, x \rangle.$$

Thus $\{v'_j, u'_j, w'_j\}$ are uniformly and independently distributed over $\{0,1\}^k$. Moreover

$$M_1 x = (v'_1, \ldots, v'_s, u'_1, \ldots, u'_{i-s}), \qquad M_2 x = (v'_1, \ldots, v'_s, w'_1, \ldots, w'_{i-s}).$$

It follows using Lemma 3.14 that

$$\begin{aligned}
\mathop{\mathbb{E}}_{x}\left[f_{=i}(M_1 x) f_{=i}(M_2 x)\right] &= \mathop{\mathbb{E}}_{v'_j, u'_j, w'_j}\left[f_{=i}(v'_1, \ldots, v'_s, u'_1, \ldots, u'_{i-s}) f_{=i}(v'_1, \ldots, v'_s, w'_1, \ldots, w'_{i-s})\right] \\
&= \mathop{\mathbb{E}}_{v'_j, w'_j}\left[\mathop{\mathbb{E}}_{u'_1, \ldots, u'_{i-s}}\left[f_{=i}(v'_1, \ldots, v'_s, u'_1, \ldots, u'_{i-s})\right] \cdot f_{=i}(v'_1, \ldots, v'_s, w'_1, \ldots, w'_{i-s})\right] \\
&= 0.
\end{aligned}$$

□

Lemma 4.2. *Let $M \in \mathcal{M}[i,\ell]$. Then*

$$\mathop{\mathbb{E}}_{x_1, \ldots, x_\ell \in \{0,1\}^k}\left[f^2_{=i}(Mx)\right] = \frac{\beta_{i,i}}{\beta_{i,\ell}} \|F_{=i}\|_2^2.$$

Proof. Denoting $(y_1, \ldots, y_i) = Mx$, we note that $y_1, \ldots, y_i$ are uniformly and independently distributed in $\{0,1\}^k$ and hence the expectation above, call it Γ, does not depend on the choice of $M \in \mathcal{M}[i,\ell]$. Also, due to basis variance, the expectation is the same as $\Gamma = \mathbb{E}_x\left[f_{=i}(M_1 x) f_{=i}(M_2 x)\right]$ as long as $\mathsf{rowspan}(M_1) = \mathsf{rowspan}(M_2)$. Using Lemma 3.13, squaring, and taking expectation over x,

$$F_{=i}[x_1, \ldots, x_\ell] = \frac{1}{\beta_{i,i}} \sum_{M \in \mathcal{M}[i,\ell]} f_{=i}(Mx).$$

$$\beta^2_{i,i} \cdot \|F_{=i}\|_2^2 = \sum_{\substack{M_1, M_2 \in \mathcal{M}[i,\ell] \\ \mathsf{rowspan}(M_1) = \mathsf{rowspan}(M_2)}} \Gamma + \sum_{\substack{M_1, M_2 \in \mathcal{M}[i,\ell] \\ \mathsf{rowspan}(M_1) \neq \mathsf{rowspan}(M_2)}} \mathop{\mathbb{E}}_{x}\left[f_{=i}(M_1 x) f_{=i}(M_2 x)\right].$$

The lemma follows by noting that that are $\beta_{i,\ell}\beta_{i,i}$ pairs M_1, M_2 with the same row-span and by previous Lemma 4.1, the expectation vanishes when the row-spans are distinct. □

The left hand side in the statement of this Lemma equals $\|f_{=i}\|_2^2$ and we have $\beta_{i,i} \leqslant 2^{i^2-1}$, $\beta_{i,\ell} \geqslant \frac{1}{2} \cdot 2^{i\ell}$. We record this very useful fact for future:

Lemma 4.3. $\|f_{=i}\|_2^2 \leqslant \frac{2^{i^2}}{2^{i\ell}} \|F_{=i}\|_2^2$.

4.2 Three-wise Correlations

Understanding three-wise correlations is more difficult. Here, we will need to use the Fourier analytic machinery developed in Section 3. Our formal result is:

Theorem 4.4. *Let S be a basis-invariant set of vertices in $\mathsf{H}_{k,\ell}$ that is (r,ε) pseudo-random. Let $F\colon H_{k,\ell} = (\{0,1\}^k)^\ell \to \{0,1\}$ be the indicator function of S and $\eta = \|F_{=i}\|_2^2$. Then for any $0 \leqslant i \leqslant r$, $i \leqslant d \leqslant 3i$, $A \subseteq \{0,1\}^\ell$ of dimension d and $M_1, M_2, M_3 \in \mathcal{M}[i,\ell]$ such that $\oplus_{s=1}^3 \mathsf{rowspan}(M_s) = A$, we have that*

$$\left|\mathop{\mathbb{E}}_{x \in (\{0,1\}^k)^\ell}\left[f_{=i}(M_1 x) f_{=i}(M_2 x) f_{=i}(M_3 x)\right]\right| \leqslant 2^{4r^2} \frac{\eta\sqrt{\varepsilon}}{2^{d\ell}}. \tag{15}$$

The rest of this section is devoted to the proof of the above lemma. Fix i, $i \leqslant d \leqslant 3i$, $A \subseteq \{0,1\}^\ell$ of dimension d, and $M_1, M_2, M_3 \in \mathcal{M}[i,\ell]$ whose direct sum of row spaces is A. Since $f_{=i}$ is basis invariant, we are free to rewrite the rows of each matrix as long as the row-span is preserved. We will spend some effort into bringing the matrices into a convenient form. We begin with the following simple observation.

Lemma 4.5. *Suppose* $\mathsf{rowspan}(M_3) \not\subseteq \mathsf{rowspan}(M_1) \oplus \mathsf{rowspan}(M_2)$ *(or the other two symmetric cases). Then*

$$\mathop{\mathbb{E}}_{x \in (\{0,1\}^k)^\ell} [f_{=i}(M_1 x) f_{=i}(M_2 x) f_{=i}(M_3 x)] = 0.$$

Proof. The proof is essentially the same as that of Lemma 4.1. We can choose linearly independent vectors $w_1, \ldots, w_t, v_1, \ldots, v_s$ such that $t \geqslant 1$ and

- $w_1, \ldots, w_t$ are the first t rows of M_3.
- $v_1, \ldots, v_s$ span the remaining $i - t$ rows of M_3 as well as $\mathsf{rowspan}(M_1) \oplus \mathsf{rowspan}(M_2)$.

We define $v'_j = \langle v_j, x \rangle$ and $w'_j = \langle w_j, x \rangle$ so that $\{v'_j, w'_j\}$ are uniformly and independently distributed in $\{0,1\}^k$. Clearly, $M_1 x, M_2 x$ depend only on v'_j and $M_3 x = (w'_1, \ldots, w'_t, y_1, \ldots, y_{i-t})$ where $y_1, \ldots, y_{i-t}$ depend only on v'_j. Fixing an arbitrary choice of v'_j fixes $a^* = M_1 x, b^* = M_2 x$ and $c^* = (y_1, \ldots, y_{i-t})$. The expectation is then

$$f_{=i}(a^*) f_{=i}(b^*) \mathop{\mathbb{E}}_{w'_1, \ldots, w'_t} \left[f_{=i}(w'_1, \ldots, w'_t, c^*) \right],$$

which vanishes according to Lemma 3.14. □

Thanks to Lemma 4.5, we assume henceforth that the row-span of each matrix is contained in the direct sum of the other two. Let $H = \cap_{j=1}^3 \mathsf{rowspan}(M_j), \mathsf{dim}(H) = s$, and let $g_1, \ldots, g_s$ be a basis for it. We may assume w.l.o.g. that the first s rows of each matrix are $g_1, \ldots, g_s$. Let M'_1, M'_2, M'_3 be the matrices M_1, M_2, M_3 after removing these first s rows. By our assumptions, we have $\cap_{j=1}^3 \mathsf{rowspan}(M'_j) = \{0\}$ and moreover the row-span of each is contained in the direct sum of the other two. We show that one can assume a strong structure on the row-spans of M'_1, M'_2, M'_3 as below. We recommend reading the proof as similar tricks are used hereafter.

Lemma 4.6. *The row spans of* M'_1, M'_2, M'_3 *have the form*

$$\begin{aligned}
&\mathsf{Span}(w_1, \ldots, w_t, \quad y_1, \ldots, y_n, \quad y_{n+1}, \ldots, y_{i-s-t}) \\
&\mathsf{Span}(w_1, \ldots, w_t, \quad z_1, \ldots, z_n, \quad z_{n+1}, \ldots, z_{i-s-t}) \\
&\mathsf{Span}(y_1 + z_1, \ldots, y_n + z_n, \quad y_{n+1}, \ldots, y_{i-s-t}, \quad z_{n+1}, \ldots, z_{i-s-t})
\end{aligned}$$

where the vectors $w_1, \ldots, w_t, y_1, \ldots, y_{i-s-t}, z_1, \ldots, z_{i-s-t}$ *are linearly independent.*

Proof. Let $\{w_1, \ldots, w_t\}$ be the basis for $\mathsf{rowspan}(M'_1) \cap \mathsf{rowspan}(M'_2)$. Let

$$A = \mathsf{rowspan}(M'_1), B = \mathsf{rowspan}(M'_2), C = \mathsf{rowspan}(M'_3)$$

so that $A \cap B \cap C = \{0\}$ and each is contained in the direct sum of the remaining two. If we pretend that $w_1 = \ldots = w_t = 0$ (which really amounts to working with a quotient space, but we find this informal description clearer), we can pretend that $A \cap B = \{0\}$. Apply Lemma A.1 to get the desired form. One caveat however is that each variable y_j above (the same goes for z_j) is really $y_j + \sigma(w)$ where $\sigma(w)$ denotes some arbitrary linear combination of $w_1, \ldots, w_t$ (not necessarily the same for different y_j, z_j)), a side effect of "pulling back" from the quotient space. Nevertheless, this can be fixed by simply redefining $y_j \leftarrow y_j + \sigma(w)$. □

We now turn back to the task of upper bounding the expectation in (15). We make a change of variables: $g'_j = \langle g_j, x\rangle$ for $j = 1, \ldots, s$, $w'_j = \langle w_j, x\rangle$ for $j = 1, \ldots, t$, and $y'_j = \langle y_j, x\rangle$ and $z'_j = \langle z_j, x\rangle$ for $j = 1, \ldots, i-s-t$. Since these vectors are linearly independent, we have that our g', w', y', z' variables are independent and uniform over $\{0,1\}^k$. For notational simplicity, we will just drop the primes in the superscripts, and relabel these variables as g, w, y, z. Thus the expectation in (15) equals

$$\mathop{\mathbb{E}}_{g,w,y,z}\Big[f_{=i}(g_1,\ldots,g_s,\ \ w_1,\ldots,w_t,\ \ y_1,\ldots,y_n,\ \ y_{n+1},\ldots,y_{i-s-t}) \\ f_{=i}(g_1,\ldots,g_s,\ \ w_1,\ldots,w_t,\ \ z_1,\ldots,z_n,\ \ z_{n+1},\ldots,z_{i-s-t}) \\ f_{=i}(g_1,\ldots,g_s,\ \ y_1+z_1,\ldots,y_n+z_n,\ \ y_{n+1},\ldots,y_{i-s-t},\ \ z_{n+1},\ldots,z_{i-s-t})\Big].$$

Denote $h_{g_1,\ldots,g_s}(a_1,\ldots,a_{i-s}) = f_{=i}(g_1,\ldots,g_s,a_1,\ldots,a_{i-s})$. To reduce cumbersome notation, we drop the subscript from h for now and remember that it is $g_1,\ldots,g_s$ throughout. Then our expectation is

$$\mathop{\mathbb{E}}_{g,w,y,z}\Big[h(w_1,\ldots,w_t,\ \ y_1,\ldots,y_n,\ \ y_{n+1},\ldots,y_{i-s-t}) \\ h(w_1,\ldots,w_t,\ \ z_1,\ldots,z_n,\ \ z_{n+1},\ldots,z_{i-s-t}) \\ h(y_1+z_1,\ldots,y_n+z_n,\ \ y_{n+1},\ldots,y_{i-s-t},\ \ z_{n+1},\ldots,z_{i-s-t})\Big].$$

For a tuple $(b_1, b_2, \ldots, b_n)$ and $m_1 \leqslant m_2$, we will denote by $b_{[m_1:m_2]}$ the sub-tuple $(b_{m_1}, b_{m_1+1}, \ldots, b_{m_2})$. Applying the Fourier transform on h and using the expectation over w, y, z, we see that the expectation equals

$$\mathop{\mathbb{E}}_{g}\ \sum_{\substack{T_1,\ldots,T_n\\ W_1,\ldots,W_t\\ P_{n+1},\ldots,P_{i-s-t}\\ Q_{n+1},\ldots,Q_{i-s-t}}} \Big[\widehat{h}(W_{[1:t]}, T_{[1:n]}, P_{[n+1:i-s-t]}) \cdot \widehat{h}(W_{[1:t]}, T_{[1:n]}, Q_{[n+1:i-s-t]}) \\ \widehat{h}(T_{[1:n]}, P_{[n+1:i-s-t]}, Q_{[n+1:i-s-t]})\Big].$$

For ease of notation, we will denote by T the tuple $(T_1, \ldots, T_n)$ and similarly for W, P, Q. Thus the expression above can be written as

$$\mathop{\mathbb{E}}_{g}\left[\sum_{T,W,P,Q} \widehat{h}(W,T,P)\cdot\widehat{h}(W,T,Q)\cdot\widehat{h}(T,P,Q)\right].$$

For a fixed g and T, the sum over W, P, Q can be upper bounded in absolute value by repeated Cauchy-

Schwartz as (this technique will be very useful; it is summarized as Lemma A.4):

$$\begin{aligned}
&\sum_{W,P} |\widehat{h}(W,T,P)| \left(\sum_{Q} |\widehat{h}(W,T,Q)| \cdot |\widehat{h}(T,P,Q)| \right) \\
&\leqslant \sum_{W,P} |\widehat{h}(W,T,P)| \left(\sqrt{\sum_{Q} \widehat{h}^2(W,T,Q)} \sqrt{\sum_{Q} \widehat{h}^2(T,P,Q)} \right) \\
&= \sum_{W} \sqrt{\sum_{Q} \widehat{h}^2(W,T,Q)} \left(\sum_{P} |\widehat{h}(W,T,P)| \cdot \sqrt{\sum_{Q} \widehat{h}^2(T,P,Q)} \right) \\
&\leqslant \sum_{W} \sqrt{\sum_{Q} \widehat{h}^2(W,T,Q)} \left(\sqrt{\sum_{P} \widehat{h}^2(W,T,P)} \sqrt{\sum_{P,Q} \widehat{h}^2(T,P,Q)} \right) \\
&= \sqrt{\sum_{P,Q} \widehat{h}^2(T,P,Q)} \sum_{W} \left(\sqrt{\sum_{Q} \widehat{h}^2(W,T,Q)} \sqrt{\sum_{P} \widehat{h}^2(W,T,P)} \right) \\
&\leqslant \sqrt{\sum_{P,Q} \widehat{h}^2(T,P,Q)} \sqrt{\sum_{W,Q} \widehat{h}^2(W,T,Q)} \sqrt{\sum_{W,P} \widehat{h}^2(W,T,P)} \\
&\stackrel{def}{=} \sqrt{A_1(T)} \sqrt{A_2(T)} \sqrt{A_2(T)},
\end{aligned}$$

where we labeled the three expressions inside the square roots as $A_1(T), A_2(T), A_2(T)$ respectively, noting that the second and the third are really the same. Considering the expectation over g, and further upper bounding

$$\mathop{\mathbb{E}}_{g} \left[\sum_{T} \sqrt{A_1(T)}\, A_2(T) \right] \leqslant \sqrt{\max_{g,T} A_1(T)} \cdot \mathop{\mathbb{E}}_{g} \left[\sum_{T} A_2(T) \right].$$

By Parseval and by Lemma 4.3,

$$\mathop{\mathbb{E}}_{g} \left[\sum_{T} A_2(T) \right] = \mathop{\mathbb{E}}_{g} \left[\sum_{T,W,P} \widehat{h}^2(W,T,P) \right] = \mathop{\mathbb{E}}_{g} \left[\|h\|_2^2 \right] = \|f_{=i}\|_2^2 \leqslant \frac{2^{i^2}}{2^{i\ell}}\, \eta \leqslant \frac{2^{r^2}}{2^{i\ell}}\, \eta. \tag{16}$$

By Lemma 3.20, we have

$$\max_{g,T} A_1(T) = \max_{g,T} \sum_{P,Q} \widehat{h}^2(T,P,Q) \leqslant \frac{2^{6i^3}}{2^{2n+(2\cdot(i-s-t-n))\ell}}\, \varepsilon \leqslant \frac{2^{6r^3}}{2^{2(i-s-t)\ell}}\, \varepsilon. \tag{17}$$

Combining both upper bounds (16) and (17), and noting that $d = 2i - s - t$, we get the desired upper bound

$$\left(\frac{2^{6r^3}}{2^{2(i-s-t)\ell}}\, \varepsilon \right)^{\frac{1}{2}} \frac{2^{r^2}}{2^{i\ell}}\, \eta \;\leqslant\; 2^{4r^2} \frac{\eta\sqrt{\varepsilon}}{2^{(2i-s-t)\ell}}.$$

5 Four-wise Correlations: Getting the Matrices into Convenient Form

We now begin the proof of our main technical Theorem 2.15. This section is devoted to bringing the matrices M_1, M_2, M_3, M_4 into a convenient form and the actual analysis is presented in subsequent sections. We emphasize again that we can rewrite rows of the four matrices as long as each row-span is preserved. Thanks to the lemma below, we assume henceforth that the row-span of each matrix is contained in the direct sum of the remaining three.

Lemma 5.1. *Suppose* $\mathsf{rowspan}(M_4) \not\subseteq \oplus_{j=1}^{3}\mathsf{rowspan}(M_j)$ *(or the other three symmetric cases). Then*

$$\mathop{\mathbb{E}}_{x\in(\{0,1\}^k)^\ell}[f_{=i}(M_1x)f_{=i}(M_2x)f_{=i}(M_3x))f_{=i}(M_4x)] = 0.$$

Proof. Essentially the same as that of Lemma 4.5. □

5.1 Removing 4-wise and 3-wise Intersections of Rowspaces

Consider the subspace $\cap_{j=1}^{4}\mathsf{rowspan}(M_j)$. Let H_4 be a basis for it and $h_4 = |H_4|$ be its dimension. We may assume w.l.o.g. that the first h_4 rows of each matrix are precisely H_4 and the rest of their rows are linear combinations of vectors $v_1, \ldots, v_r$ that are linearly independent of H_4. The rows H_4 are removed now from each matrix; they will only come into play at the very end of the analysis. For notational convenience, we refer to the matrices with these rows removed also as M_1, M_2, M_3, M_4 respectively. We assume henceforth that $\cap_{j=1}^{4}\mathsf{rowspan}(M_j) = \{0\}$ and that the row-span of each matrix is contained in the direct sum of the remaining three.

We handle 3-wise intersections of the row-spaces in the same manner. Suppose there is a non-zero vector $w \in \cap_{j=1}^{3}\mathsf{rowspan}(M_j)$. Since we assumed that the 4-wise intersection of the row-spaces is trivial, $w \notin \mathsf{rowspan}(M_4)$. We may assume w.l.o.g. that w is the first row of M_1, M_2, M_3 and their rest of the rows as well as the rows of M_4 are linear combinations of vectors $v_1, \ldots, v_r$ (not necessarily the same as in the previous para) that are linearly independent of w. The row w is removed now from M_1, M_2, M_3; it will only come into play at the very end of the analysis. For notational convenience, we refer to the matrices with this row removed also as M_1, M_2, M_3 respectively (and M_4 is unaffected). This process is repeated as long as there is a non-trivial 3-wise intersection of the row-spaces. At the end of this process, let H_3 denote the set of all row-vectors thus removed, $h_3 = |H_3|$, and s_1, s_2, s_3, s_4 be the number of remaining rows of the respective matrices. Since the original number of rows was i and h_4 were removed in the earlier step, the number of rows removed from the j^{th} matrix in the current step is $i - h_4 - s_j$.

We assume henceforth that the matrices M_1, M_2, M_3, M_4 do not have non-trivial 3-wise intersection of their row-spaces, that the row-space of each is contained in the direct sum of the remaining three, and that their number of rows is s_1, s_2, s_3, s_4 respectively.

5.2 Getting M_1, M_2, M_3 into Form

We first write M_1, M_2, M_3 in a convenient form. Letting $A = \mathsf{rowspan}(M_1)$, $B = \mathsf{rowspan}(M_2)$, and

$$C = \mathsf{rowspan}(M_3) \cap (\mathsf{rowspan}(M_1) \oplus \mathsf{rowspan}(M_2)),$$

and applying an argument similar to Lemma 4.6 and Lemma A.1, we can write A, B, C as

$$\begin{aligned}
&\mathsf{Span}(v_1, \ldots, v_t,\ \ p_1, \ldots, p_n,\ \ u,\ \ y)\\
&\mathsf{Span}(v_1, \ldots, v_t,\ \ q_1, \ldots, q_n,\ \ w,\ \ z)\\
&\mathsf{Span}(p_1 + q_1, \ldots, p_n + q_n, u, w)
\end{aligned}$$

where u, y, w, z denote tuples of vectors (we do not wish to use an index/subscript to denote their length) and the vectors $\{v_j, p_j, q_j, u_j, y_j, w_j, z_j)$ are linearly independent. Now we complete the basis for C to that of $\mathsf{rowspan}(M_3)$ by adding linearly independent vectors $a = (a_1, \ldots, a_h)$ from $\mathsf{rowspan}(M_3) \setminus C$. Hence the row-spans of M_1, M_2, M_3 can be assumed to be in the form (p, q have the same length n):

$$\begin{aligned}
&\mathsf{Span}(v, p, u, y)\\
&\mathsf{Span}(v, q, w, z)\\
&\mathsf{Span}(a, p_1 + q_1, \ldots, p_n + q_n, u, w).
\end{aligned} \tag{18}$$

5.3 Getting M_4 into Form: Part I

Now we begin the rather tedious process of getting M_4 into a convenient form given the form (18) for the first three matrices.

We pretend first that $v = p = u = q = w = 0$ (formally, taking a quotient). The first three row-spaces now amount to $Y = \mathsf{Span}(y)$, $Z = \mathsf{Span}(z)$, and $A = \mathsf{Span}(a)$. Denoting $W = \mathsf{rowspan}(M_4)$ (its quotient to be precise), we have that $W \subseteq A \oplus Y \oplus Z$. Using Lemma A.2, there is a basis for W of the following form $\cup_{s=1}^{7} A_s$ where

$$\begin{array}{lll}
A_1 = \{ & a_i + y_j + z_k & \mid\ i \in \Sigma_1,\ \ j \in \Phi_1,\ \ k \in \Psi_1\}\\
A_2 = \{ & a_i + y_j & \mid\ i \in \Sigma_2,\ \ j \in \Phi_2\ \ \}\\
A_3 = \{ & a_i \quad + z_k & \mid\ i \in \Sigma_3,\ \ k \in \Psi_3\}\\
A_4 = \{ & a_i \quad + \sigma(y_{\Phi_1,\Phi_5}) & \mid\ i \in \Sigma_4\ \ \}\\
A_5 = \{ & \quad y_j + z_k & \mid\ \ j \in \Phi_5,\ \ k \in \Psi_5\}\\
A_6 = \{ & \quad y_j & \mid\ \ j \in \Phi_6\ \ \}\\
A_7 = \{ & \quad\ z_k & \mid\ \ k \in \Psi_7\}.
\end{array}$$

Here $\sigma(y_{\Phi_1,\Phi_5})$ are arbitrary linear forms in $\{y_j | j \in \Phi_1 \cup \Phi_5\}$ (possibly different ones, but we hide this fact as it will be essentially irrelevant). We emphasize that the notation (and similar ones) $\{a_i + y_j + z_k \mid i \in \Sigma_1, j \in \Phi_1, k \in \Psi_1\}$ is imprecise, but chosen for the sake of ease. Here $|\Sigma_1| = |\Phi_1| = |\Psi_1|$ and there are exactly $|\Sigma_1|$ vectors in this set, forming a kind of a perfect matching. We further emphasize the following observation.

Informally speaking, if all forms $\sigma(y_{\Phi_1,\Phi_5})$ are ignored, then each a, y, z variable appears in the above representation exactly once. Formally,

- $\Sigma_1 \cup \Sigma_2 \cup \Sigma_3 \cup \Sigma_4$ (disjointly) cover all a-variables.
- $\Phi_1 \cup \Phi_2 \cup \Phi_5 \cup \Phi_6$ (disjointly) cover all y-variables.
- $\Psi_1 \cup \Psi_3 \cup \Psi_5 \cup \Psi_7$ (disjointly) cover all z-variables.

These three statements are simply consequences of $A \subseteq Y \oplus Z \oplus W$, $Y \subseteq A \oplus Z \oplus W$, $Z \subseteq A \oplus Y \oplus W$ respectively. Now we "pull back" from the quotient space. This has the effect of adding a $\sigma(v,p,u,q,w)$ term to each vector that denotes an arbitrary linear form in those variables (and since these forms would be essentially irrelevant, we use the same notation for all). This yields a partial basis for $\mathsf{rowspan}(M_4)$ summarized below.

Lemma 5.2. $\mathsf{rowspan}(M_4)$ *has a partial basis of the following form* $\cup_{s=1}^{7} A_s$ *where*

$$\begin{array}{lllll}
A_1 = \{ & a_i + y_j + z_k & +\sigma(v,p,u,q,w) & \mid & i \in \Sigma_1, \quad j \in \Phi_1, \quad k \in \Psi_1\} \\
A_2 = \{ & a_i + y_j & +\sigma(v,p,u,q,w) & \mid & i \in \Sigma_2, \quad j \in \Phi_2 \quad \} \\
A_3 = \{ & a_i \quad + z_k & +\sigma(v,p,u,q,w) & \mid & i \in \Sigma_3, \quad k \in \Psi_3\} \\
A_4 = \{ & a_i & +\sigma(v,p,u,q,w) + \sigma(y_{\Phi_1,\Phi_5}) & \mid & i \in \Sigma_4 \quad \} \\
A_5 = \{ & y_j + z_k & +\sigma(v,p,u,q,w) & \mid & j \in \Phi_5, \quad k \in \Psi_5\} \\
A_6 = \{ & y_j & +\sigma(v,p,u,q,w) & \mid & j \in \Phi_6 \quad \} \\
A_7 = \{ & z_k & +\sigma(v,p,u,q,w) & \mid & k \in \Psi_7\}.
\end{array}$$

Moreover, if all forms $\sigma(y_{\Phi_1,\Phi_5})$ *are ignored, then each* a, y, z *variable appears in the above representation exactly once.*

5.4 Getting M_4 into Form: Part II

In the previous subsection, we obtained a partial basis for $\mathsf{rowspan}(M_4)$ by pretending that $v = p = u = q = w = 0$ (but did add $\sigma(v,p,u,q,w)$ terms back to account for this). This basis can now be extended to a basis for $\mathsf{rowspan}(M_4)$ by adding in a basis for

$$W = \mathsf{rowspan}(M_4) \cap \mathsf{Span}(v,p,u,q,w).$$

We do this in two steps. First, we pretend that $v = u = w = 0$. Let $P = \mathsf{Span}(p)$ and $Q = \mathsf{Span}(q)$ (note that $n = \mathsf{dim}(P) = \mathsf{dim}(Q)$). Since $W \subseteq P \oplus Q$, by Lemma A.3, for a partition of the index set $\{1,\ldots,n\} = \Delta_0 \cup \Delta_1 \cup \Delta_2 \cup \ldots \cup \Delta_m \cup \Omega_1 \cup \Omega_2$, we may assume that W has a basis

$$\begin{array}{lllll}
A_8 = & \{ & p_i + \sigma(q) & \mid i \in \Delta_1 & \} \\
A_9 = B_2 \cup \ldots \cup B_m & & & & \\
A_{10} = & \{ & p_i + \sigma(q_{\Omega_1}) & \mid \quad i \in \Omega_1 \cup \Omega_2 & \} \\
A_{11} = & \{ & q_j & \mid \quad j \in \Omega_2 & \}.
\end{array}$$

Here $B_s = \left\{ q_j + \sigma(p_{\Delta_{[s+1:m]}}) \;\middle|\; j \in \Delta_s \right\}$. We recall that $\Delta_{[s+1:m]} = \Delta_{s+1} \cup \ldots \cup \Delta_m \cup \Omega_1 \cup \Omega_2$. As usual $\sigma(\cdot)$ are linear forms in its inputs that we do not really care about. We "pull back" from the quotient space by adding $\sigma(v,u,w)$ to every vector yielding:

Lemma 5.3. *A partial basis for* $\mathsf{rowspan}(M_4)$ *from Lemma 5.2 can be further extended as* $A_8 \cup A_9 \cup A_{10} \cup A_{11}$ *where*

$$\begin{array}{lllll}
A_8 = & \{ \, p_i + \sigma(q) & +\sigma(v,u,w) & \mid i \in \Delta_1 & \} \\
A_9 = B_2 \cup \ldots \cup B_m & & & & \\
A_{10} = & \{ \, p_i + \sigma(q_{\Omega_1}) & +\sigma(v,u,w) & \mid \quad i \in \Omega_1 \cup \Omega_2 & \} \\
A_{11} = & \{ \quad q_j & +\sigma(v,u,w) & \mid \quad j \in \Omega_2 & \}.
\end{array}$$

Here $B_s = \left\{ q_j + \sigma(p_{\Delta_{[s+1:m]}}) + \sigma(v,u,w) \;\middle|\; j \in \Delta_s \right\}$. Finally, we can complete a basis for $\mathsf{rowspan}(M_4)$ by adding a basis for $W = \mathsf{rowspan}(M_4) \cap \mathsf{Span}(v,u,w)$. This can clearly be done by first pretending $v = w = 0$, writing the basis for $W \subseteq \mathsf{Span}(u)$, pulling it back by adding forms $\sigma(v,w)$, and then finally completing the basis by adding a basis for $\mathsf{rowspan}(M_4) \cap \mathsf{Span}(v,w)$. We summarize this as:

Lemma 5.4. *A partial basis for* $\mathsf{rowspan}(M_4)$ *from Lemmas 5.2 and 5.3 can be completed by adding* $A_{12} \cup A_{13}$ *where (for some index set* Γ*)*

$$\begin{aligned} A_{12} &= \{\, u_i + \sigma(v,w) \mid \; i \in \Gamma \} \\ A_{13} &= \mathsf{basis}(\mathsf{rowspan}(M_4) \cap \mathsf{Span}(v,w)). \end{aligned}$$

To summarize, basis for $\mathsf{rowspan}(M_4)$ can be assumed to be $\cup_{s=1}^{13} A_s$ where

$$\begin{array}{lllll|lll} A_1 = \{ & a_i + y_j + z_k & +\sigma(v,p,u,q,w) & & | & i \in \Sigma_1, & j \in \Phi_1, \quad k \in \Psi_1 \} \\ A_2 = \{ & a_i + y_j & +\sigma(v,p,u,q,w) & & | & i \in \Sigma_2, & j \in \Phi_2 \quad \} \\ A_3 = \{ & a_i \quad + z_k & +\sigma(v,p,u,q,w) & & | & i \in \Sigma_3, & k \in \Psi_3 \} \\ A_4 = \{ & a_i & +\sigma(v,p,u,q,w) & + \sigma(y_{\Phi_1,\Phi_5}) & | & i \in \Sigma_4 & \quad \} \\ A_5 = \{ & y_j + z_k & +\sigma(v,p,u,q,w) & & | & j \in \Phi_5, & k \in \Psi_5 \} \\ A_6 = \{ & y_j & +\sigma(v,p,u,q,w) & & | & j \in \Phi_6 & \quad \} \\ A_7 = \{ & z_k & +\sigma(v,p,u,q,w) & & | & & k \in \Psi_7 \} \end{array}$$

$$\begin{array}{lllll|l} A_8 = & \{\, p_i + \sigma(q) & & +\sigma(v,u,w) & | & i \in \Delta_1 \quad \} \\ A_9 = B_2 \cup \ldots \cup B_m & & & & & \\ A_{10} = & \{\, p_i + \sigma(q_{\Omega_1}) & & +\sigma(v,u,w) & | & i \in \Omega_1 \cup \Omega_2 \quad \} \\ A_{11} = & \{ & q_j & +\sigma(v,u,w) & | & j \in \Omega_2 \quad \} \end{array}$$

$$\begin{aligned} A_{12} &= \{\, u_i + \sigma(v,w) \mid \; i \in \Gamma \} \\ A_{13} &= \mathsf{basis}(\mathsf{rowspan}(M_4) \cap \mathsf{Span}(v,w)). \end{aligned} \tag{19}$$

Here $B_s = \left\{ q_j + \sigma(p_{\Delta_{[s+1:m]}}) + \sigma(v,u,w) \;\middle|\; j \in \Delta_s \right\}$. The variables $\{a_i, y_j, z_k\}$ appearing in $A_1, \ldots, A_7$, the variables $\{p_i, q_j\}$ appearing in $A_8, \ldots, A_{11}$ and the variables $\{u_i\}$ appearing in A_{12} will be called pivots (the reader should ignore the $\sigma(\cdot)$ forms to clearly understand which variables we are referring to as pivots).

5.5 Getting M_4 into Form: Part III

In this section, we make further changes to the basis for $\mathsf{rowspan}(M_4)$ that are needed towards our final proof. We recommend however that the reader skips this section and jumps to the next section where we present a proof in a special but instructive case.

Step 1

We start with the basis in (19). We observe that:

- (v,p,u) variables can be "absorbed into" the pivot y-variables,
- (v,q,w) variables can be absorbed into the pivot z-variables,

- u-variables can be absorbed into the pivot p-variables, and
- w-variables can be absorbed into the pivot q-variables.

Therefore, if we have a y-variable as a pivot, there is no need to include (v, p, u)-variables in the corresponding $\sigma(\cdot)$ form (and similarly for the z, p, q pivots). This leads to the simplified $\sigma(\cdot)$ forms as shown below.

$$
\begin{array}{lllll}
A_1 = \{ & a_i + y_j + z_k & & \mid & i \in \Sigma_1, \quad j \in \Phi_1, \quad k \in \Psi_1\} \\
A_2 = \{ & a_i + y_j \quad +\sigma(\quad q, w) & & \mid & i \in \Sigma_2, \quad j \in \Phi_2 \quad\quad \} \\
A_3 = \{ & a_i \quad + z_k \quad +\sigma(\quad p, u \quad) & & \mid & i \in \Sigma_3, \quad\quad k \in \Psi_3\} \\
A_4 = \{ & a_i \quad\quad +\sigma(v, p, u, q, w) + \sigma(y_{\Phi_1,\Phi_5}) & & \mid & i \in \Sigma_4 \quad\quad\quad\quad \} \\
A_5 = \{ & \quad y_j + z_k & & \mid & \quad\quad j \in \Phi_5, \quad k \in \Psi_5\} \\
A_6 = \{ & \quad y_j \quad +\sigma(\quad q, w) & & \mid & \quad\quad j \in \Phi_6 \quad\quad \} \\
A_7 = \{ & \quad\quad z_k \quad +\sigma(\quad p, u \quad) & & \mid & \quad\quad\quad\quad k \in \Psi_7\}
\end{array}
$$

$$
\begin{array}{lllll}
A_8 = & \{\, p_i + \sigma(q) & +\sigma(v, \quad w) & \mid & i \in \Delta_1 \quad\quad \} \\
A_9 = B_2 \cup \ldots \cup B_m & & & & \\
A_{10} = & \{\, p_i + \sigma(q_{\Omega_1}) & +\sigma(v, \quad w) & \mid & i \in \Omega_1 \cup \Omega_2 \quad \} \\
A_{11} = & \{ \quad\quad\quad q_j & +\sigma(v, u \quad) & \mid & j \in \Omega_2 \quad\quad \}
\end{array}
$$

$$
\begin{aligned}
A_{12} &= \{\, u_i + \sigma(v, w) \mid \quad i \in \Gamma\} \\
A_{13} &= \mathsf{basis}(\mathsf{rowspan}(M_4) \cap \mathsf{Span}(v, w)).
\end{aligned}
$$

Here $B_s = \left\{ q_j + \sigma(p_{\Delta_{[s+1:m]}}) + \sigma(v, u \quad) \;\middle|\; j \in \Delta_s \right\}$.

Step 2

Consider A_7 and its vectors $\{z_k + \sigma(p, u) \mid k \in \Psi_7\}$. By adding vectors from A_8 if necessary, we can assume that the form $\sigma(\cdot)$ does not depend on p_{Δ_1} (we may need re-absorption of v, q, w into z_k). We wish to make the dependence on p_{Δ_0} more restrictive. So our concern is with their $z_k + \sigma(p_{\Delta_0})$ component. We can change the matched basis for p_{Δ_0} so that for a partition $\Delta_0 = \Delta_0' \cup \Delta_0''$, $\Psi_7 = \Psi_{7a} \cup \Psi_{7b}$, these components turn into

$$\{z_k + p_s \mid k \in \Psi_{7a}, s \in \Delta_0'\} \;\cup\; \{z_k + \sigma(p_{\Delta_0'}) \mid k \in \Psi_{7b}\}.$$

Further, adding the former to the latter as necessary (which amounts to a change of basis for z_{Ψ_7}), the latter components can be made independent of p_{Δ_0} altogether. Additionally, for the former we may absorb u and $p_{\Delta_{[2:m]}}$ into p_s. Thus we may split A_7 into A_{7a} and A_{7b} as shown below. We emphasize that $|\Delta_0'| = |\Psi_{7a}|$. With these changes, the basis for $\mathsf{rowspan}(M_4)$ can be written as:

$$
\begin{array}{lllll}
A_1 = \{ & a_i + y_j + z_k & & \mid & i \in \Sigma_1, \quad j \in \Phi_1, \quad k \in \Psi_1\} \\
A_2 = \{ & a_i + y_j \quad +\sigma(\quad q, w) & & \mid & i \in \Sigma_2, \quad j \in \Phi_2 \quad\quad \} \\
A_3 = \{ & a_i \quad + z_k \quad +\sigma(\quad p, u \quad) & & \mid & i \in \Sigma_3, \quad\quad k \in \Psi_3\} \\
A_4 = \{ & a_i \quad\quad +\sigma(v, p, u, q, w) + \sigma(y_{\Phi_1,\Phi_5}) & & \mid & i \in \Sigma_4 \quad\quad\quad\quad \} \\
A_5 = \{ & \quad y_j + z_k & & \mid & \quad\quad j \in \Phi_5, \quad k \in \Psi_5\} \\
A_6 = \{ & \quad y_j \quad +\sigma(\quad q, w) & & \mid & \quad\quad j \in \Phi_6 \quad\quad \} \\
A_{7a} = \{ & \quad\quad z_k + p_s & & \mid & \quad\quad\quad\quad k \in \Psi_{7a}, s \in \Delta_0'\} \\
A_{7b} = \{ & \quad\quad z_k \quad +\sigma(\quad p_{\Delta_{[2:m]}}, u \quad) & & \mid & \quad\quad\quad\quad k \in \Psi_{7b}\}
\end{array}
$$

$$
\begin{array}{lllll}
A_8 = & \{\, p_i + \sigma(q) & +\sigma(v, \quad w) & \mid & i \in \Delta_1 \quad\quad \} \\
A_9 = B_2 \cup \ldots \cup B_m & & & & \\
A_{10} = & \{\, p_i + \sigma(q_{\Omega_1}) & +\sigma(v, \quad w) & \mid & i \in \Omega_1 \cup \Omega_2 \quad \} \\
A_{11} = & \{ \qquad\qquad q_j & +\sigma(v, u \quad) & \mid & j \in \Omega_2 \quad\quad \}
\end{array}
$$

$$
\begin{aligned}
A_{12} &= \{\, u_i + \sigma(v, w) \mid \quad i \in \Gamma\} \\
A_{13} &= \mathsf{basis}(\mathsf{rowspan}(M_4) \cap \mathsf{Span}(v, w)).
\end{aligned}
$$

Here $B_s = \left\{ q_j + \sigma(p_{\Delta_{[s+1:m]}}) + \sigma(v, u \quad) \;\middle|\; j \in \Delta_s \right\}$.

Step 3

Finally, we consider A_{11} and its vectors $\{q_j + \sigma(v, u \quad) | \; j \in \Omega_2\}$. By adding vectors from A_{12} if necessary, we can assume that the form $\sigma(\cdot)$ does not depend on u_Γ (we may need re-absorption of w into q_j). In other words, $\sigma(\cdot)$ depends only on the remaining variables of u denoted as $u_{\overline{\Gamma}}$. By a change of basis on $u_{\overline{\Gamma}}$ variables, we can write, for some $\overline{\Gamma}_0 \subseteq \overline{\Gamma}$,

$$
\mathsf{Span}(A_{11}) = \mathsf{Span}(\{u_i + \sigma(q_{\Omega_2}, v) \mid i \in \overline{\Gamma}_0\}) \oplus (\mathsf{Span}(A_{11}) \cap \mathsf{Span}(q_{\Omega_2}, v))\,.
$$

Out of these two component spaces, we retain only the latter as new A_{11} and merge the former with A_{12} (redefining new Γ as $\Gamma \cup \overline{\Gamma}_0$). Thus we reach our **final form:**

$$
\begin{array}{lllllll}
A_1 = \{ & a_i + y_j + z_k & & \mid & i \in \Sigma_1, & j \in \Phi_1, & k \in \Psi_1\} \\
A_2 = \{ & a_i + y_j & +\sigma(\quad\; q, w) & \mid & i \in \Sigma_2, & j \in \Phi_2 & \quad\} \\
A_3 = \{ & a_i \quad\; + z_k & +\sigma(\;\; p, u \quad) & \mid & i \in \Sigma_3, & & k \in \Psi_3\} \\
A_4 = \{ & a_i & +\sigma(v, p, u, q, w) + \sigma(y_{\Phi_1, \Phi_5}) & \mid & i \in \Sigma_4 & & \quad\} \\
A_5 = \{ & \quad y_j + z_k & & \mid & & j \in \Phi_5, & k \in \Psi_5\} \\
A_6 = \{ & \quad y_j & +\sigma(\quad\; q, w) & \mid & & j \in \Phi_6 & \quad\} \\
A_{7a} = \{ & \quad\; z_k + p_s & & \mid & & & k \in \Psi_{7a}, s \in \Delta_0'\} \\
A_{7b} = \{ & \quad\; z_k & +\sigma(\;\; p_{\Delta_{[2:m]}}, u \quad) & \mid & & & k \in \Psi_{7b}\}
\end{array}
$$

$$
\begin{array}{lllll}
A_8 = & \{\, p_i + \sigma(q) & +\sigma(v, \quad w) & \mid & i \in \Delta_1 \quad\quad \} \\
A_9 = \quad B_2 \cup \ldots \cup B_m & & & & \\
A_{10} = & \{\, p_i + \sigma(q_{\Omega_1}) & +\sigma(v, \quad w) & \mid & i \in \Omega_1 \cup \Omega_2 \quad \} \\
A_{11} \subseteq \quad \mathsf{Span}(q_{\Omega_2}, v). & & & &
\end{array}
$$

Here $B_s = \left\{ q_j + \sigma(p_{\Delta_{[s+1:m]}}) + \sigma(v, u \quad) \;\middle|\; j \in \Delta_s \right\}$.

$$
\begin{aligned}
A_{12} &= \{\, u_i + \sigma(q_{\Omega_2}, v, w) \mid \quad i \in \Gamma\} \\
A_{13} &= \mathsf{basis}(\mathsf{rowspan}(M_4) \cap \mathsf{Span}(v, w)).
\end{aligned} \tag{20}
$$

6 Four-wise Correlations: A (somewhat) Simplified Case

We now begin the proof of our main technical Theorem 2.15, i.e. to upper bound the expectation

$$\mathop{\mathbb{E}}_{x\in(\{0,1\}^k)^\ell}[f_{=i}(M_1x)f_{=i}(M_2x)f_{=i}(M_3x)f_{=i}(M_4x)]. \tag{21}$$

For the benefit of the reader, this section presents the proof in the special case where in the basis for $\mathsf{rowspan}(M_4)$ given in (19) (ignore Section 5.5 and modifications therein for now):

- All the linear forms $\sigma(\cdot)$ are zero.
- $A_{10} = A_{11} = A_{13} = \emptyset$. A_9 consists of just B_2.
- There is no further partition of A_7 into A_{7a} and A_{7b}.

Given matrices M_1, M_2, M_3, M_4, we note:

- Let $H_4 = g = \{g_1, \ldots, g_{h_4}\}$ be the rows that appeared in all four matrices (and were removed).
- Let $H_3 = r = \{r_1, \ldots, r_{h_3}\}$ be the rows that appeared in (exactly) three matrices (and were removed). Let $r(1), r(2), r(3), r(4) \subseteq H_3$ be the sets of rows that appeared in the four matrices respectively, so that $|r(1)| + |r(2)| + |r(3)| + |r(4)| = 3 \cdot h_3$.
- When we take expectation over $x \in (\{0,1\}^k)^\ell$, if w is a row of a matrix, we make the change of basis $w' = \langle w, x\rangle$ where w' is uniformly distributed over $\{0,1\}^k$ and moreover independently for rows that are linearly independent. For the ease of notation, we drop the prime from the superscript and call the new variable w as well.

Thus we assume that:

$$\begin{aligned}
M_1x &= g, r(1), v, p, u, y\\
M_2x &= g, r(2), v, q, w, z\\
M_3x &= g, r(3), a, p_1 + q_1, \ldots, p_n + q_n, u, w\\
M_4x &= g, r(4), A_1, \ldots, A_9, A_{12}.
\end{aligned} \tag{22}$$

We recall that (in the present special case):

$$\begin{array}{lllll}
A_1 = \{ & a_i + y_j + z_k & \mid & i \in \Sigma_1, \quad j \in \Phi_1, & k \in \Psi_1\}\\
A_2 = \{ & a_i + y_j & \mid & i \in \Sigma_2, \quad j \in \Phi_2 & \}\\
A_3 = \{ & a_i \quad + z_k & \mid & i \in \Sigma_3, & k \in \Psi_3\}\\
A_4 = \{ & a_i & \mid & i \in \Sigma_4 & \}\\
A_5 = \{ & y_j + z_k & \mid & j \in \Phi_5, & k \in \Psi_5\}\\
A_6 = \{ & y_j & \mid & j \in \Phi_6 & \}\\
A_7 = \{ & z_k & \mid & & k \in \Psi_7\}
\end{array}$$

$$\begin{array}{lll}
A_8 = & \{ p_i & \mid i \in \Delta_1 \quad \}\\
A_9 = & \{ \quad q_j & \mid \quad j \in \Delta_2 \quad \}
\end{array}$$

$$A_{12} = \{ u_i \mid \ i \in \Gamma\}.$$

Lemma 6.1. *The dimension d of $\oplus_{j=1}^{4}(\mathsf{rowspan}(M_j))$ is:*

$$\begin{aligned} d =& |g| + |r| + |v| + (|p| + |q|) + (|u|) + |w| + (|a| + |y| + |z|) \\ =& |g| + |r| + |v| + (2|\Delta_0| + 2|\Delta_1| + 2|\Delta_2|) + (|\Gamma| + |\overline{\Gamma}|) + |w| + \\ & (3|\Sigma_1| + 2|\Sigma_2| + 2|\Sigma_3| + |\Sigma_4| + 2|\Phi_5| + |\Phi_6| + |\Psi_7|). \end{aligned}$$

Proof. The number of all the variables appearing above are added up. It is noted that the p and q variables both equal in number to $|\Delta_0| + |\Delta_1| + |\Delta_2|$. Also, $|\Sigma_1| = |\Phi_1| = |\Psi_1|$ and similar equalities. We note that the input u is partitioned as $(u_\Gamma, u_{\overline{\Gamma}})$. □

We split inputs $M_1x, \ldots, M_4x$ in (22) into three parts: Fourier analysis will be applied on the third part, Cauchy-Schwartz on the second part, and the first part will be thought of as a "restriction". The splits are as below. To clarify the notation, p_{Δ_0} denotes, as before, the variables $\{p_i \mid i \in \Delta_0\}$, $(p+q)_{\Delta_0}$ denotes the variables $\{p_j + q_j \mid j \in \Delta_0\}$, and for the ease of notation, $(a+y+z)_1$ denotes the triples $\{a_i + y_j + z_k \in A_1\}$ (and similarly).

	L	J	K
$M_1x =:$	$\{g, r(1), u_\Gamma\}$	$\{v, p_{\Delta_1 \cup \Delta_2}, u_{\overline{\Gamma}}, y_{\Phi_6}\}$	$\{p_{\Delta_0}, y_{\Phi_1}, y_{\Phi_2}, y_{\Phi_5}\}$
$M_2x =:$	$\{g, r(2)\}$	$\{v, q_{\Delta_1 \cup \Delta_2}, w, z_{\Psi_7}\}$	$\{q_{\Delta_0}, z_{\Psi_1}, z_{\Psi_3}, z_{\Psi_5}\}$
$M_3x =:$	$\{g, r(3), u_\Gamma\}$	$\{a_{\Sigma_4}, (p+q)_{\Delta_1 \cup \Delta_2}, u_{\overline{\Gamma}}, w\}$	$\{a_{\Sigma_1}, a_{\Sigma_2}, a_{\Sigma_3}, (p+q)_{\Delta_0}\}$
$M_4x =:$	$\{g, r(4), u_\Gamma\}$	$\{a_{\Sigma_4}, y_{\Phi_6}, z_{\Psi_7}, p_{\Delta_1}, q_{\Delta_2}\}$	$\{(a+y+z)_1, (a+y)_2, (a+z)_3, (y+z)_5\}$

Denoting the parts in the splits as $(L_1, J_1, K_1), \ldots, (L_4, J_4, K_4)$ respectively, consider the restrictions:

$$\lambda_{1,L_1,J_1}(K_1) = f_{=i}(L_1, J_1, K_1), \ldots, \lambda_{4,L_4,J_4}(K_4) = f_{=i}(L_4, J_4, K_4).$$

Dropping the subscripts (but keeping in mind that they are always there), the goal is to upper bound

$$\mathbb{E}\left[\lambda_1(K_1)\lambda_2(K_2)\lambda_3(K_3)\lambda_4(K_4)\right], \tag{23}$$

where for notational ease, we did not write the long list of variables that the expectation is taken over. We do note that L_s, J_s, K_s all depend on the inputs. Writing the K_s explicitly:

$$\begin{aligned} \lambda_1(\quad & p_{\Delta_0}, && y_{\Phi_1}, && y_{\Phi_2}, && y_{\Phi_5}, &&) \\ \lambda_2(\quad & q_{\Delta_0}, && z_{\Psi_1}, && z_{\Psi_3}, && z_{\Psi_5}, &&) \\ \lambda_3(\quad & a_{\Sigma_1}, && a_{\Sigma_2}, && a_{\Sigma_3}, && (p+q)_{\Delta_0} &&) \\ \lambda_4(\quad & a_{\Sigma_1} + y_{\Phi_1} + z_{\Psi_1}, && a_{\Sigma_2} + y_{\Phi_2}, && a_{\Sigma_3} + z_{\Psi_3}, && y_{\Phi_5} + z_{\Psi_5} &&). \end{aligned}$$

The notation (and similar ones) $a_{\Sigma_1} + y_{\Phi_1} + z_{\Psi_1}$ is imprecise, but we use it for the ease. It really refers to $\{a_i + y_j + z_k \in A_1 \mid i \in \Sigma_1, j \in \Phi_1, k \in \Psi_1\}$. Now writing the λ_s in the Fourier representation and taking expectation over its inputs, we see that the expectation in (23) equals (there is a product of four terms that are written one below the other for visual ease)

$$\mathbb{E}\left[\sum_{W,Y,Z,S,B} \begin{array}{l} \widehat{\lambda}_1(S,W,Y,B) \\ \widehat{\lambda}_2(S,W,Z,B) \\ \widehat{\lambda}_3(W,Y,Z,S) \\ \widehat{\lambda}_4(W,Y,Z,B). \end{array}\right] \tag{24}$$

To explain the reasoning, we note that the Fourier expansion will have a term (as part of a larger product term)

$$\cdots\cdots \chi_W(a_{\Sigma_1} + y_{\Phi_1} + z_{\Psi_1})\, \chi_{W'}(y_{\Phi_1})\, \chi_{W''}(z_{\Psi_1})\, \chi_{W'''}(a_{\Sigma_1}) \cdots\cdots$$

and taking expectation over $a_{\Sigma_1}, y_{\Phi_1}, z_{\Psi_1}$, the term vanishes unless $W = W' = W'' = W'''$. Similar reasoning is applied above to "Fourier tuples" Y, Z, S, B.

For fixed $L_1, \ldots, L_4, W, Y, Z, S, B$, we consider the expectation over $J_1, \ldots, J_4$ (or rather inputs in those sets). The point here is that all inputs in $J_1, \ldots, J_4$ appear twice:

- Exactly twice, these being $\{v, u_{\overline{\Gamma}}, y_{\Phi_6}, w, z_{\Psi_7}, a_{\Sigma_4}\}$.
- Or "effectively" exactly twice, these being $p_{\Delta_1}, p_{\Delta_2}, q_{\Delta_1}, q_{\Delta_2}$. What we mean here is that for indices in Δ_1 (and similarly in Δ_2), we have inputs $p_{\Delta_1}, q_{\Delta_1}, (p+q)_{\Delta_1}, p_{\Delta_1}$ appearing in J_1, J_2, J_3, J_4 respectively. These can be paired as $(p_{\Delta_1}, q_{\Delta_1})$ and $((p+q)_{\Delta_1}, p_{\Delta_1})$. The latter pair is distributed same as the former and this is what matters for applying Cauchy-Schwarz.

Replacing the Fourier coefficients by their absolute values and using repeated Cauchy-Schwartz (see Lemma A.4), we see that (24) is upper bounded by

$$\mathop{\mathbb{E}}_{g,r,u_\Gamma}\left[\sum_{W,Y,Z,S,B} \sqrt{\mathop{\mathbb{E}}_{J_1}\left[\widehat{\lambda}^2_{1,J_1}(S,W,Y,B)\right]}\sqrt{\mathop{\mathbb{E}}_{J_2}\left[\widehat{\lambda}^2_{2,J_2}(S,W,Z,B)\right]}\right.$$
$$\left.\sqrt{\mathop{\mathbb{E}}_{J_3}\left[\widehat{\lambda}^2_{3,J_3}(W,Y,Z,S)\right]}\sqrt{\mathop{\mathbb{E}}_{J_4}\left[\widehat{\lambda}^2_{4,J_4}(W,Y,Z,B)\right]}\right].$$

Again applying Cauchy-Schwartz (note that the pairing is first-third and fourth-second factors) we get an upper bound $\sqrt{\mathsf{Term}_1} \cdot \sqrt{\mathsf{Term}_2}$ where

$$\mathsf{Term}_1 = \mathop{\mathbb{E}}_{g,r,u_\Gamma}\left[\sum_{W,Y,Z,S,B} \mathop{\mathbb{E}}_{J_1}\left[\widehat{\lambda}^2_{1,J_1}(S,W,Y,B)\right]\mathop{\mathbb{E}}_{J_3}\left[\widehat{\lambda}^2_{3,J_3}(W,Y,Z,S)\right]\right]$$
$$\mathsf{Term}_2 = \mathop{\mathbb{E}}_{g,r,u_\Gamma}\left[\sum_{W,Y,Z,S,B} \mathop{\mathbb{E}}_{J_4}\left[\widehat{\lambda}^2_{4,J_4}(W,Y,Z,B)\right]\mathop{\mathbb{E}}_{J_2}\left[\widehat{\lambda}^2_{2,J_2}(S,W,Z,B)\right]\right].$$

We consider Term_1. Noting that W, Y, S appear in both $\widehat{\lambda}_1(\cdot), \widehat{\lambda}_3(\cdot)$, B appears only in $\widehat{\lambda}_1(\cdot)$, Z appears only in $\widehat{\lambda}_3(\cdot)$, and that $\lambda_3(\cdot)$ does not depend on $r(1) \setminus r(3)$ (so expectation over it can be pushed inside),

we can rewrite Term_1 as:

$$\mathsf{Term}_1 = \mathop{\mathbb{E}}_{g,r,u_\Gamma}\left[\sum_{W,Y,S}\left(\mathop{\mathbb{E}}_{r(1)\setminus r(3),J_1}\left[\sum_B \widehat{\lambda}^2_{1,J_1}(S,W,Y,B)\right]\right)\left(\mathop{\mathbb{E}}_{J_3}\left[\sum_Z \widehat{\lambda}^2_{3,J_3}(W,Y,Z,S)\right]\right)\right].$$

Finally, using Lemma A.5, we have the upper bound:

$$\mathsf{Term}_1 \leqslant \left(\max_{\substack{g,r(1)\cap r(3),u_\Gamma,\\ W,Y,S}}\mathop{\mathbb{E}}_{r(1)\setminus r(3),J_1}\left[\sum_B \widehat{\lambda}^2_{1,J_1}(S,W,Y,B)\right]\right)\left(\mathop{\mathbb{E}}_{g,r,u_\Gamma,J_3}\left[\sum_{W,Y,S,Z} \widehat{\lambda}^2_{3,J_3}(W,Y,Z,S)\right]\right).$$

The second factor is $\mathbb{E}_{L_3,J_3,K_3}\left[\|\lambda_{3,L_3,J_3}(K_3)\|_2^2\right] = \|f_{=i}\|_2^2 \leqslant \frac{2^{i^2}}{2^{i\ell}}\,\eta$. The first factor is bounded by, using Lemma 3.20, $2^{6i^3}\frac{\varepsilon}{2^{d_1\cdot\ell}}$ where

$$\begin{aligned} d_1 &= (|J_1| + |r(1)\setminus r(3)|) + 2(|W|+|Y|+|S|) + |B| \\ &= |v| + |\Delta_1| + |\Delta_2| + |\overline{\Gamma}| + |\Phi_6| + |r(1)\setminus r(3)| + 2|\Sigma_1| + 2|\Sigma_2| + 2|\Delta_0| + |\Phi_5|. \end{aligned}$$

We similarly re-write Term_2 as:

$$\mathsf{Term}_2 = \mathop{\mathbb{E}}_{g,r,u_\Gamma}\left[\sum_{W,Z,B}\left(\mathop{\mathbb{E}}_{\substack{J_4,u_\Gamma,\\ r(4)\setminus r(2)}}\left[\sum_Y \widehat{\lambda}^2_{4,J_4}(W,Y,Z,B)\right]\right)\left(\mathop{\mathbb{E}}_{J_2}\left[\sum_S \widehat{\lambda}^2_{2,J_2}(S,W,Z,B)\right]\right)\right].$$

Here $\lambda_2(\cdot)$ does not depend on u_Γ and $r(4)\setminus r(2)$, so both are pushed inside. As before,

$$\mathsf{Term}_2 \leqslant \left(\max_{\substack{g,r(4)\cap r(2),\\ W,Z,B}}\mathop{\mathbb{E}}_{\substack{J_4,u_\Gamma,\\ r(4)\setminus r(2)}}\left[\sum_Y \widehat{\lambda}^2_{4,J_4}(W,Y,Z,B)\right]\right)\left(\mathop{\mathbb{E}}_{g,r,u_\Gamma,J_2}\left[\sum_{W,Z,B,S} \widehat{\lambda}^2_{2,J_2}(S,W,Z,B)\right]\right).$$

The second factor is $\mathbb{E}_{L_2,J_2,K_2}\left[\|\lambda_{2,L_2,J_2}(K_2)\|_2^2\right] = \|f_{=i}\|_2^2 \leqslant \frac{2^{i^2}}{2^{i\ell}}\,\eta$. The first factor is bounded by, using Lemma 3.20, $2^{6i^3}\frac{\varepsilon}{2^{d_2\cdot\ell}}$ where

$$\begin{aligned} d_2 &= (|J_4| + |r(4)\setminus r(2)| + |\Gamma|) + 2(|W|+|Z|+|B|) + |Y| \\ &= |\Sigma_4| + |\Phi_6| + |\Psi_7| + |\Delta_1| + |\Delta_2| + |r(4)\setminus r(2)| + |\Gamma| + 2|\Sigma_1| + 2|\Sigma_3| + 2|\Phi_5| + |\Sigma_2|. \end{aligned}$$

The proof of Theorem 2.15 (in the special case) is complete by recalling that we have an upper bound of $\sqrt{\mathsf{Term}_1}\sqrt{\mathsf{Term}_2}$ and that $i \leqslant r$ and $\frac{1}{2}((d_1+i)+(d_2+i)) = d$ as below. One gets an upper bound of $\frac{2^{7r^3}}{2^{d\ell}}\,\eta\varepsilon$ in Theorem 2.15.

Lemma 6.2. $d_1 + i + d_2 + i = 2d$.

Proof. We write down expressions for d_1, d_2 as above followed by expressions for i ($= |L_3 \cup J_3 \cup K_3|$) and i ($= |L_2 \cup J_2 \cup K_2|$):

$$\begin{aligned} d_1 = \quad & |v| + |\Delta_1| + |\Delta_2| + |\overline{\Gamma}| + |\Phi_6| + |r(1)\setminus r(3)| + 2|\Sigma_1| + 2|\Sigma_2| + 2|\Delta_0| + |\Phi_5|. \\ d_2 = \quad & |\Sigma_4| + |\Phi_6| + |\Psi_7| + |\Delta_1| + |\Delta_2| + |r(4)\setminus r(2)| + |\Gamma| + 2|\Sigma_1| + 2|\Sigma_3| + 2|\Phi_5| + |\Sigma_2|. \\ i = \quad & |g| + |r(3)| + |\Gamma| + |\Sigma_4| + |\Delta_1| + |\Delta_2| + |\overline{\Gamma}| + |w| + |\Sigma_1| + |\Sigma_2| + |\Sigma_3| + |\Delta_0|. \\ i = \quad & |g| + |r(2)| + |v| + |\Delta_1| + |\Delta_2| + |w| + |\Psi_7| + |\Delta_0| + |\Sigma_1| + |\Sigma_3| + |\Phi_5|. \end{aligned}$$

It can be verified that the overall sum is exactly $2d$ where as in Lemma 6.1,

$$\begin{aligned} d = \quad & |g| + |r| + |v| + 2|\Delta_0| + 2|\Delta_1| + 2|\Delta_2| + |\Gamma| + |\overline{\Gamma}| + |w| + \\ & 3|\Sigma_1| + 2|\Sigma_2| + 2|\Sigma_3| + |\Sigma_4| + 2|\Phi_5| + |\Phi_6| + |\Psi_7|. \end{aligned}$$

One notes that since every element of $r = r(1) \cup r(2) \cup r(3) \cup r(4)$ is contained in precisely three of these sets, $|r| = |r(3)| + |r(1) \setminus r(3)| = |r(2)| + |r(4) \setminus r(2)|$. □

7 Four-wise Correlations: the General Case

We now begin the full proof of our main technical Theorem 2.15, i.e. to upper bound the expectation

$$\mathop{\mathbb{E}}_{x \in (\{0,1\}^k)^\ell} [f_{=i}(M_1 x) f_{=i}(M_2 x) f_{=i}(M_3 x) f_{=i}(M_4 x)]. \tag{25}$$

Given matrices M_1, M_2, M_3, M_4, we recall:

- Let $H_4 = g = \{g_1, \ldots, g_{h_4}\}$ be the rows that appeared in all four matrices (and were removed).
- Let $H_3 = r = \{r_1, \ldots, r_{h_3}\}$ be the rows that appeared in (exactly) three matrices (and were removed). Let $r(1), r(2), r(3), r(4) \subseteq H_3$ be the sets of rows that appeared in the four matrices respectively, so that $|r(1)| + |r(2)| + |r(3)| + |r(4)| = 3 \cdot h_3$.
- When we take expectation over $x \in (\{0,1\}^k)^\ell$, if w is a row of a matrix, we make the change of basis $w' = \langle w, x \rangle$ where w' is uniformly distributed over $\{0,1\}^k$ and moreover independently for rows that are linearly independent. For the ease of notation, we drop the prime from the superscript and call the new variable w as well.

Thus we assume that (given the basis for $\mathsf{rowspan}(\mathsf{M_4})$ by (20), written again below for convenience):

$$\begin{aligned} M_1 x &= g, r(1), v, p, u, y \\ M_2 x &= g, r(2), v, q, w, z \\ M_3 x &= g, r(3), a, p_1 + q_1, \ldots, p_n + q_n, u, w \\ M_4 x &= g, r(4), A_1, \ldots, A_6, A_{7a}, A_{7b}, A_8, \ldots, A_{13}. \end{aligned} \tag{26}$$

Lemma 7.1. *The dimension d of $\oplus_{j=1}^4 (\mathsf{rowspan}(M_j))$ is:*

$$\begin{aligned} d =& |g| + |r| + |v| + (|p| + |q|) + (|u|) + |w| + (|a| + |y| + |z|) \\ =& |g| + |r| + |v| + (2|\Delta_0'| + 2|\Delta_0''| + 2|\Delta_1| + 2|\Delta_2| + \ldots + 2|\Delta_m| + 2|\Omega_1| + 2|\Omega_2|) + (|\Gamma| + |\overline{\Gamma}|) + |w| + \\ & (3|\Sigma_1| + 2|\Sigma_2| + 2|\Sigma_3| + |\Sigma_4| + 2|\Phi_5| + |\Phi_6| + |\Psi_{7a}| + |\Psi_{7b}|). \end{aligned}$$

Proof. The number of all the variables appearing above are added up. It is noted that the p and q variables both equal in number to $|\Delta_0'| + |\Delta_0''| + |\Delta_1| + |\Delta_2| + \ldots + |\Delta_m| + |\Omega_1| + |\Omega_2|$ and $\Delta_0 = \Delta_0' \cup \Delta_0''$. We have $|\Sigma_1| = |\Phi_1| = |\Psi_1|$ and similar equalities. We emphasize that $|\Delta_0'| = |\Psi_{7a}|$. □

We recall for convenience that:

$$
\begin{array}{lllll}
A_1 = \{ & a_i + y_j + z_k & & \mid & i \in \Sigma_1, \quad j \in \Phi_1, \quad k \in \Psi_1\} \\
A_2 = \{ & a_i + y_j & +\sigma(\quad q, w) & \mid & i \in \Sigma_2, \quad j \in \Phi_2 \quad \} \\
A_3 = \{ & a_i \quad + z_k & +\sigma(\quad p, u \quad) & \mid & i \in \Sigma_3, \quad k \in \Psi_3\} \\
A_4 = \{ & a_i & +\sigma(v,p,u,q,w) + \sigma(y_{\Phi_1,\Phi_5}) & \mid & i \in \Sigma_4 \quad \} \\
A_5 = \{ & \quad y_j + z_k & & \mid & j \in \Phi_5, \quad k \in \Psi_5\} \\
A_6 = \{ & \quad y_j & +\sigma(\quad q, w) & \mid & j \in \Phi_6 \quad \} \\
A_{7a} = \{ & \quad z_k + p_s & & \mid & k \in \Psi_{7a}, s \in \Delta_0'\} \\
A_{7b} = \{ & \quad z_k & +\sigma(\quad p_{\Delta_{[2:m]}}, u \quad) & \mid & k \in \Psi_{7b}\}
\end{array}
$$

$$
\begin{array}{lllll}
A_8 = & \{\, p_i + \sigma(q) & +\sigma(v, \quad w) & \mid & i \in \Delta_1 \quad \} \\
A_9 = & B_2 \cup \ldots \cup B_m & & & \\
A_{10} = & \{\, p_i + \sigma(q_{\Omega_1}) & +\sigma(v, \quad w) & \mid & i \in \Omega_1 \cup \Omega_2 \quad \} \\
A_{11} \subseteq & \mathsf{Span}(q_{\Omega_2}, v). & & &
\end{array}
$$

Here $B_s = \left\{ q_j + \sigma(p_{\Delta_{[s+1:m]}}) + \sigma(v, u \quad) \;\middle|\; j \in \Delta_s \right\}$.

$$
\begin{aligned}
A_{12} &= \{\, u_i + \sigma(q_{\Omega_2}, v, w) \mid \quad i \in \Gamma\} \\
A_{13} &= \mathsf{basis}(\mathsf{rowspan}(M_4) \cap \mathsf{Span}(v, w)).
\end{aligned}
$$

We split each input in (26) into two parts. The mix of Fourier analysis and Cauchy-Schwartz will not be very clean. The splits are as below.

$$
\begin{array}{lll}
 & J & K \\
M_1 x =: & \{g, r(1), v, p_{\Delta_{[2:m]}}, u_\Gamma, u_{\overline{\Gamma}}\} & \{p_{\Delta_0 \cup \Delta_1}, y_{\Phi_1}, y_{\Phi_2}, y_{\Phi_5}, y_{\Phi_6}\} \\
M_2 x =: & \{g, r(2), v, q_{\Delta_{[2:m]}}, w, z_{\Psi_{7b}}\} & \{q_{\Delta_0 \cup \Delta_1}, z_{\Psi_1}, z_{\Psi_3}, z_{\Psi_5}, z_{\Psi_{7a}}\} \\
M_3 x =: & \{g, r(3), (p+q)_{\Delta_{[2:m]}}, u_\Gamma, u_{\overline{\Gamma}}, w\} & \{a, (p+q)_{\Delta_0 \cup \Delta_1}\} \\
M_4 x =: & \{g, r(4), z_{\Psi_{7b}}, A_9, A_{10}, A_{11}, A_{12}, A_{13}\} & \{(a+y+z)_1, (a+y)_2, (a+z)_3, \\
 & & a_{\Sigma_4}, (y+z)_5, y_{\Phi_6}, z_{\Psi_{7a}} + p_{\Delta_0'}, p_{\Delta_1}\Big\}.
\end{array}
$$

We are using an imprecise notation: inputs for $M_4 x$ (except for $g, r(4), (a+y+z)_1, (y+z)_5$) have the additional $\sigma(\cdot)$ terms that are omitted from the notation for ease. Denoting the parts in the splits as $(J_1, K_1), \ldots, (J_4, K_4)$ respectively, consider the restrictions:

$$\lambda_{1,J_1}(K_1) = f_{=i}(J_1, K_1), \ldots, \lambda_{4,J_4}(K_4) = f_{=i}(J_4, K_4).$$

Dropping the subscripts (but keeping in mind that they are always there), the goal is to upper bound

$$\mathbb{E}\left[\lambda_1(K_1)\lambda_2(K_2)\lambda_3(K_3)\lambda_4(K_4)\right], \tag{27}$$

where for notational ease, we did not write the long list of variables that the expectation is taken over. We do note that J_s, K_s all depend on the inputs. Writing the K_s explicitly:

$$\begin{array}{llllllll}
\lambda_1(& p_{\Delta_0'}, & p_{\Delta_0''}, & p_{\Delta_1}, & y_{\Phi_1}, & y_{\Phi_2}, & y_{\Phi_5}, & y_{\Phi_6} &) \\
\lambda_2(& q_{\Delta_0'}, & q_{\Delta_0''}, & q_{\Delta_1}, & z_{\Psi_1}, & z_{\Psi_3}, & z_{\Psi_5}, & z_{\Psi_{7a}} &) \\
\lambda_3(& a_{\Sigma_1}, & a_{\Sigma_2}, & a_{\Sigma_3}, & a_{\Sigma_4}, & (p+q)_{\Delta_0'}, & (p+q)_{\Delta_0''}, & (p+q)_{\Delta_1} &) \\
\lambda_4(& a_{\Sigma_1}+y_{\Phi_1}+z_{\Psi_1}, & a_{\Sigma_2}+y_{\Phi_2}, & a_{\Sigma_3}+z_{\Psi_3}, & a_{\Sigma_4}, & y_{\Phi_5}+z_{\Psi_5}, & y_{\Phi_6}, & z_{\Psi_{7a}}+p_{\Delta_0'}, \quad p_{\Delta_1} &).
\end{array}$$

Now writing the λ_s in the Fourier representation and taking expectation over their inputs, we see that the expectation in (27) equals, up to a caveat to be fixed shortly, (there is a product of four terms that are written one below the other for visual ease)

$$\mathop{\mathbb{E}}_{\substack{g,r,v,u_\Gamma,u_{\overline{\Gamma}},w \\ p_{\Delta_{[2:m]}},q_{\Delta_{[2:m]}},z_{\Psi_{7b}}}} \left[\sum_{\substack{W,Y,Z,B,P \\ T,Q,N,S,D,X}} \mathsf{sign} \cdot \begin{array}{llllllll}
\widehat{\lambda}_1(& S+Q, & D, & X+N, & W, & Y, & P, & T &) \\
\widehat{\lambda}_2(& S, & D, & X, & W, & Z, & P, & Q &) \\
\widehat{\lambda}_3(& W, & Y, & Z, & B, & S, & D, & X &) \\
\widehat{\lambda}_4(& W, & Y, & Z, & B, & P, & T, & Q, \quad N &).
\end{array} \right] \tag{28}$$

A remark: there are $\sigma(\cdot)$ terms that were omitted from the notation. They have a two-fold effect. Firstly, there is a $\mathsf{sign} \in \{-1, 1\}$ that depends on $(Y, Z, B, T, N, S, D, X; v, u, w, p, q)$. We will take absolute values immediately next, so this sign does not really matter. Secondly, there are additional $\sigma(\cdot)$ terms now in the Fourier domain, and the form of the Fourier coefficients is not quite as in (28), but actually as below:

$$\begin{array}{llllllll}
\widehat{\lambda}_1(& S+Q+\sigma(B,Z), & D+\sigma(B,Z), & X+N+\sigma(B,Z), & W+\sigma(B), & Y, & P+\sigma(B), & T &) \\
\widehat{\lambda}_2(& S+\sigma(Y,B,T,N), & D+\sigma(Y,B,T,N), & X+\sigma(Y,B,T,N), & W, & Z, & P, & Q &) \\
\widehat{\lambda}_3(& W, & Y, & Z, & B, & S, & D, & X &) \\
\widehat{\lambda}_4(& W, & Y, & Z, & B, & P, & T, & Q, \quad N &).
\end{array} \tag{29}$$

Denoting the Fourier coefficients as $\widehat{\lambda}_1(V_1), \widehat{\lambda}_2(V_2), \widehat{\lambda}_3(V_3), \widehat{\lambda}_4(V_4)$, an upper bound on the desired expectation is:

$$\mathop{\mathbb{E}}_{\substack{g,r,v,u_\Gamma,u_{\overline{\Gamma}},w,z_{\Psi_{7b}}\\ p_{\Delta_{[2:m]}},q_{\Delta_{[2:m]}}}} \sum_{\substack{W,Y,Z,B,P\\ T,Q,N,S,D,X}} \left[|\widehat{\lambda}_1(V_1)| \cdot |\widehat{\lambda}_2(V_2)| \cdot |\widehat{\lambda}_3(V_3)| \cdot |\widehat{\lambda}_4(V_4)|\right]$$

$$= \mathop{\mathbb{E}}_{\substack{g,r,v,u_\Gamma,u_{\overline{\Gamma}},w\\ p_{\Delta_{[2:m]}},q_{\Delta_{[2:m]}}}} \sum_{\substack{W,Y,Z,B,P\\ T,Q,N,S,D,X}} \mathop{\mathbb{E}}_{z_{\Psi_{7b}}} \left[|\widehat{\lambda}_1(V_1)| \cdot |\widehat{\lambda}_2(V_2)| \cdot |\widehat{\lambda}_3(V_3)| \cdot |\widehat{\lambda}_4(V_4)|\right].$$

We note that $z_{\Psi_{7a}}$ appears only in J_2, J_4. Using Cauchy-Schwartz, we get an upper bound

$$\mathop{\mathbb{E}}_{\substack{g,r,v,u_\Gamma,u_{\overline{\Gamma}},w\\ p_{\Delta_{[2:m]}},q_{\Delta_{[2:m]}}}} \sum_{\substack{W,Y,Z,B,P\\ T,Q,N,S,D,X}} \left[|\widehat{\lambda}_1(V_1)| \sqrt{\mathop{\mathbb{E}}_{z_{\Psi_{7b}}} \left[\widehat{\lambda}_2^2(V_2)\right]} |\widehat{\lambda}_3(V_3)| \sqrt{\mathop{\mathbb{E}}_{z_{\Psi_{7b}}} \left[\widehat{\lambda}_4^2(V_4)\right]}\right].$$

A point to note here is as follows: in λ_4, the variables $z_{\Psi_{7b}}$ actually appear along with additional $\sigma(p_{\Delta_{[2:m]}}, u)$ terms. However the expectation over these additional variables is still not considered and is still at the "outer" level. Hence the Cauchy-Schwartz over $z_{\Psi_{7b}}$ can be safely applied. Moreover, once Cauchy-Schwartz, i.e. expectation over $z_{\Psi_{7b}}$, is applied, we can ignore these $\sigma(\cdot)$ terms henceforth.[3] We will use this trick repeatedly.

Next, we consider the variables $(p_{\Delta_2}, q_{\Delta_2}), \ldots, (p_{\Delta_m}, q_{\Delta_m})$, one pair at a time. Let's consider $(p_{\Delta_2}, q_{\Delta_2})$ as an illustration. We note that p_{Δ_2} appears in J_1, q_{Δ_2} appears in J_2, $(p+q)_{\Delta_2}$ appears in J_3 and q_{Δ_2} appears in J_4. We note two points. In J_3, the distribution of $(p+q)_{\Delta_2}$ is same as that of p_{Δ_2}. In J_4, there are additional $\sigma(p_{\Delta_{[3:m]}}, v, u)$ terms but the expectation over these variables is still at the outer level. Thus we may safely apply Cauchy-Schwartz over $(p_{\Delta_2}, q_{\Delta_2})$, pairing the first-second and third-fourth factors, ignore the $\sigma(\cdot)$ terms henceforth, and get the upper bound

$$\mathop{\mathbb{E}}_{\substack{g,r,v,u_\Gamma,u_{\overline{\Gamma}},w\\ p_{\Delta_{[3:m]}},q_{\Delta_{[3:m]}}}} \sum_{\substack{W,Y,Z,B,P\\ T,Q,N,S,D,X}} \left[\sqrt{\mathop{\mathbb{E}}_{p_{\Delta_2}} \left[\widehat{\lambda}_1^2(V_1)\right]} \sqrt{\mathop{\mathbb{E}}_{z_{\Psi_{7b}},q_{\Delta_2}} \left[\widehat{\lambda}_2^2(V_2)\right]} \sqrt{\mathop{\mathbb{E}}_{p_{\Delta_2}} \left[\widehat{\lambda}_3^2(V_3)\right]} \sqrt{\mathop{\mathbb{E}}_{z_{\Psi_{7b}},q_{\Delta_2}} \left[\widehat{\lambda}_4^2(V_4)\right]}\right].$$

We apply the same argument iteratively to get an upper bound

$$\mathop{\mathbb{E}}_{\substack{g,r,v,u_\Gamma,u_{\overline{\Gamma}},w\\ p_{\Omega_1\cup\Omega_2},q_{\Omega_1\cup\Omega_2}}} \sum_{\substack{W,Y,Z,B,P\\ T,Q,N,S,D,X}} \left[\sqrt{\mathop{\mathbb{E}}_{p_{\Delta_{2,\ldots,m}}} \left[\widehat{\lambda}_1^2(V_1)\right]} \sqrt{\mathop{\mathbb{E}}_{\substack{z_{\Psi_{7b}}\\ q_{\Delta_{2,\ldots,m}}}} \left[\widehat{\lambda}_2^2(V_2)\right]} \sqrt{\mathop{\mathbb{E}}_{p_{\Delta_{2,\ldots,m}}} \left[\widehat{\lambda}_3^2(V_3)\right]} \sqrt{\mathop{\mathbb{E}}_{\substack{z_{\Psi_{7b}}\\ q_{\Delta_{2,\ldots,m}}}} \left[\widehat{\lambda}_4^2(V_4)\right]}\right].$$

Next, we Cauchy-Schwartz over $u_{\overline{\Gamma}}$. This is possible since it appears explicitly only in J_1, J_3. It appears in J_4 implicitly as part of several $\sigma(\cdot)$ terms, but all those terms got "ignored" or "eliminated" in prior steps! Hence we get an upper bound

$$\mathop{\mathbb{E}}_{\substack{g,r,v,u_\Gamma,w\\ p_{\Omega_1\cup\Omega_2},q_{\Omega_1\cup\Omega_2}}} \sum_{\substack{W,Y,Z,B,P\\ T,Q,N,S,D,X}} \left[\sqrt{\mathop{\mathbb{E}}_{\substack{u_{\overline{\Gamma}}\\ p_{\Delta_{2,\ldots,m}}}} \left[\widehat{\lambda}_1^2(V_1)\right]} \sqrt{\mathop{\mathbb{E}}_{\substack{z_{\Psi_{7b}}\\ q_{\Delta_{2,\ldots,m}}}} \left[\widehat{\lambda}_2^2(V_2)\right]} \sqrt{\mathop{\mathbb{E}}_{\substack{u_{\overline{\Gamma}}\\ p_{\Delta_{2,\ldots,m}}}} \left[\widehat{\lambda}_3^2(V_3)\right]} \sqrt{\mathop{\mathbb{E}}_{\substack{z_{\Psi_{7b}}\\ q_{\Delta_{2,\ldots,m}}}} \left[\widehat{\lambda}_4^2(V_4)\right]}\right].$$

[3] Formally, if one wishes to, by change of variables $z_{\Psi_{7b}} \leftarrow z_{\Psi_{7b}} + \sigma(p_{\Delta_{[2:m]}}, u)$.

Finally, we apply Cauchy-Schwartz twice (the pairing is first-third and fourth-second factors) to get an upper bound $\sqrt{\mathsf{Term}_1} \cdot \sqrt{\mathsf{Term}_2}$ where

$$\mathsf{Term}_1 = \mathop{\mathbb{E}}_{\substack{g,r,v,u_\Gamma,w \\ p_{\Omega_1\cup\Omega_2},q_{\Omega_1\cup\Omega_2}}} \sum_{\substack{W,Y,Z,B,P \\ T,Q,N,S,D,X}} \left[\mathop{\mathbb{E}}_{\substack{u_{\overline{\Gamma}} \\ p_{\Delta_{2,\ldots,m}}}} \left[\widehat{\lambda}_1^2(V_1)\right] \cdot \mathop{\mathbb{E}}_{\substack{u_{\overline{\Gamma}} \\ p_{\Delta_{2,\ldots,m}}}} \left[\widehat{\lambda}_3^2(V_3)\right] \right]$$

$$\mathsf{Term}_2 = \mathop{\mathbb{E}}_{\substack{g,r,v,u_\Gamma,w \\ p_{\Omega_1\cup\Omega_2},q_{\Omega_1\cup\Omega_2}}} \sum_{\substack{W,Y,Z,B,P, \\ T,Q,N,S,D,X}} \left[\mathop{\mathbb{E}}_{\substack{z_{\Psi_{7b}} \\ q_{\Delta_{2,\ldots,m}}}} \left[\widehat{\lambda}_4^2(V_4)\right] \cdot \mathop{\mathbb{E}}_{\substack{z_{\Psi_{7b}} \\ q_{\Delta_{2,\ldots,m}}}} \left[\widehat{\lambda}_2^2(V_2)\right] \right]. \tag{30}$$

Lemma 7.2. *We have the upper bound* $\mathsf{Term}_1 \leqslant 2^{7i^3} \frac{\eta\varepsilon}{2^{(d_1+i)\cdot\ell}}$ *where*

$$d_1 = |r(1)\setminus r(3)| + |\Delta_{2,\ldots,m}| + |\Omega_1| + |\Omega_2| + |v| + |\overline{\Gamma}| + 2|\Sigma_1| + 2|\Sigma_2| + 2|\Delta_0''| + |\Psi_{7a}| + |\Delta_1| + |\Phi_5| + |\Phi_6|.$$

Proof. Let us recall the the definitions of V_1, V_3:

$$V_1 = (S + Q + \sigma(B,Z),\ D + \sigma(B,Z),\ X + N + \sigma(B,Z),\ W + \sigma(B),\ P + \sigma(B),\ Y,\ T),$$

$$V_3 = (W, Y, Z, B, S, D, X).$$

Since P, Q, N do not appear in V_3 and we will only be concerned about summing over all possibilities, we might as well take V_1 as

$$V_1 = (Q,\ D + \sigma(B,Z),\ N,\ W + \sigma(B),\ P, Y, T).$$

Further in V_1, we may replace $D + \sigma(B,Z)$ by D and $W + \sigma(B)$ by W. This will induce a change in V_3, but since B, Z are present therein and the Fourier coefficients are basis invariant, the $\sigma(B,Z), \sigma(Z)$ terms there can be cleared. To summarize, we may assume that V_1 and V_3 are:

$$V_1 = (Q, D, N, W, P, Y, T), \qquad V_3 = (W, Y, Z, B, S, D, X).$$

Noting that W, Y, D are common to V_1 and V_3, we may thus write

$$\mathsf{Term}_1 = \mathop{\mathbb{E}}_{\substack{g,u_\Gamma \\ r(3)}} \sum_{W,Y,D} \left[\left(\mathop{\mathbb{E}}_{\substack{v,r(1)\setminus r(3),p_{\Omega_1\cup\Omega_2} \\ p_{\Delta_{2,\ldots,m}},u_{\overline{\Gamma}}}} \left[\sum_{Q,N,P,T} \widehat{\lambda}_1^2(W,Y,D,Q,N,P,T) \right] \right) \left(\mathop{\mathbb{E}}_{\substack{q_{\Omega_1\cup\Omega_2},w \\ p_{\Delta_{2,\ldots,m}},u_{\overline{\Gamma}}}} \left[\sum_{B,Z,S,X} \widehat{\lambda}_3^2(W,Y,D,B,Z,S,X) \right] \right) \right].$$

The vigilant reader must have noticed that we have pushed several expectations "inside". This is justified as follows. λ_1 does not depend on $q_{\Omega_1\cup\Omega_2}$ and w. λ_3 does not depend on $v, r(1)\setminus r(3)$, and since it depends only on $(p+q)_{\Omega_1\cup\Omega_2}$, it is "effectively" independent of $p_{\Omega_1\cup\Omega_2}$. Using Lemma A.5, Term_1 is bounded by

$$\left(\max_{\substack{g,r(1)\cap r(3),u_\Gamma \\ W,Y,D}} \mathop{\mathbb{E}}_{\substack{v,r(1)\setminus r(3),p_{\Omega_1\cup\Omega_2} \\ p_{\Delta_{2,\ldots,m}},u_{\overline{\Gamma}}}} \left[\sum_{\substack{Q,N, \\ P,T}} \lambda_1^2(W,Y,D,Q,N,P,T) \right] \right) \left(\mathbb{E} \left[\sum_{\substack{W,Y,D, \\ B,Z,S,X}} \widehat{\lambda}_3^2(W,Y,D,B,Z,S,X) \right] \right).$$

The second factor equals (as usual) $\|f_{=i}\|_2^2 \leqslant \frac{2^{i^2}}{2^{i\ell}}\,\eta$. The first factor is bounded, using Lemma 3.20, by $2^{6i^3}\frac{\varepsilon}{2^{d_1\cdot\ell}}$ where

$$\begin{aligned} d_1 &= |r(1)\setminus r(3)| + (|\Delta_{2,\dots,m}| + |\Omega_1| + |\Omega_2|) + |v| + |\overline{\Gamma}| + (2|W| + 2|Y| + 2|D|) + |Q| + |N| + |P| + |T| \\ &= |r(1)\setminus r(3)| + |\Delta_{2,\dots,m}| + |\Omega_1| + |\Omega_2| + |v| + |\overline{\Gamma}| + 2|\Sigma_1| + 2|\Sigma_2| + 2|\Delta_0''| + |\Psi_{7a}| + |\Delta_1| + |\Phi_5| + |\Phi_6|. \end{aligned}$$

□

Lemma 7.3. *We have the upper bound* $\mathsf{Term}_2 \leqslant 2^{7i^3}\frac{\eta\varepsilon}{2^{d_2\cdot\ell}}$ *where*

$$\begin{aligned} d_2 = {} & |r(4)\setminus r(2)| + |\Delta_{2,\dots,m}| + |\Omega_1| + |\Omega_2| + |\Gamma| + |\Psi_{7b}| + 2|\Sigma_1| + 2|\Sigma_3| + 2|\Phi_5| + 2|\Psi_{7a}| + \\ & |\Sigma_2| + |\Sigma_4| + |\Phi_6| + |\Delta_1|. \end{aligned}$$

Proof. Let us recall the the definitions of V_2, V_4.

$$V_2 = (S + \sigma(Y,B,T,N),\ \ D + \sigma(Y,B,T,N),\ \ X + \sigma(Y,B,T,N),\ \ W,\ Z,\ P,\ Q),$$

$$V_4 = (W,Y,Z,B,P,T,Q,N).$$

Since S, D, X do not appear in V_4, we might as well write $V_2 = (S,D,X,W,Z,P,Q)$. Noting that W, Z, P, Q are common to V_2 and V_4, we may thus write

$$\mathsf{Term}_2 = \mathop{\mathbb{E}}_{\substack{g,v,r(2),w\\ q_{\Omega_1\cup\Omega_2}}} \sum_{W,Z,P,Q} \left[\left(\mathop{\mathbb{E}}_{\substack{r(4)\setminus r(2),u_\Gamma,p_{\Omega_1\cup\Omega_2}\\ z_{\Psi_{7b}},q_{\Delta_{2,\dots,m}}}} \left[\sum_{Y,B,T,N} \widehat{\lambda}_4^2(W,Z,P,Q,\ \ Y,B,T,N) \right] \right) \left(\mathop{\mathbb{E}}_{z_{\Psi_{7b}},q_{\Delta_{2,\dots,m}}} \left[\sum_{S,D,X} \widehat{\lambda}_2^2(W,Z,P,Q,\ \ S,D,X) \right] \right) \right].$$

We have pushed the expectation over $r(4)\setminus r(2), u_\Gamma, p_{\Omega_1\cup\Omega_2}$ inside as λ_2 does not depend on them. Using Lemma A.5, we upper bound Term_2 by

$$\left(\max_{\substack{g,v,r(4)\cap r(2),w\\ q_{\Omega_1\cup\Omega_2},W,Z,P,Q}} \mathop{\mathbb{E}}_{\substack{r(4)\setminus r(2),u_\Gamma,p_{\Delta_{\Omega_1\cup\Omega_2}}\\ z_{\Psi_{7b}},q_{\Delta_{2,\dots,m}}}} \left[\sum_{\substack{Y,B,\\ T,N}} \widehat{\lambda}_4^2 \begin{pmatrix} W,Z,P,Q,\\ Y,B,T,N \end{pmatrix} \right] \right) \left(\mathbb{E} \left[\sum_{\substack{W,Z,P,Q\\ S,D,X}} \widehat{\lambda}_2^2 \begin{pmatrix} W,Z,P,Q,\\ S,D,X \end{pmatrix} \right] \right).$$

As before, the second factor equals $\|f_{=i}\|_2^2 \leqslant \frac{2^{i^2}}{2^{i\ell}}\,\eta$. The first factor is bounded, using Lemma 3.20, by $2^{6i^3}\frac{\varepsilon}{2^{d_2\cdot\ell}}$ where

$$\begin{aligned} d_2 = {} & |r(4)\setminus r(2)| + |\Delta_{2,\dots,m}| + |\Omega_1| + |\Omega_2| + |\Gamma| + |\Psi_{7b}| + (2|W| + 2|Z| + 2|P| + 2|Q|) + \\ & |Y| + |B| + |T| + |N| \\ = {} & |r(4)\setminus r(2)| + |\Delta_{2,\dots,m}| + |\Omega_1| + |\Omega_2| + |\Gamma| + |\Psi_{7b}| + 2|\Sigma_1| + 2|\Sigma_3| + 2|\Phi_5| + 2|\Psi_{7a}| + \\ & |\Sigma_2| + |\Sigma_4| + |\Phi_6| + |\Delta_1|. \end{aligned}$$

□

We get the overall upper bound $\sqrt{\mathsf{Term}_1}\sqrt{\mathsf{Term}_2}$, which is at most $2^{7r^3}\frac{\eta\varepsilon}{2^{d\ell}}$ (noting $i \leqslant r$) provided $\frac{1}{2}((d_1 + i) + (d_2 + i)) = d$. This is proved below completing the proof of Theorem 2.15.

Lemma 7.4. $d_1 + i + d_2 + i = 2d$.

Proof. We write down expressions for d_1, d_2 as above followed by expressions for i (from M_3) and i (from M_2):

$$\begin{aligned}
d_1 = \quad & |r(1) \setminus r(3)| + |\Delta_{2,\ldots,m}| + |\Omega_1| + |\Omega_2| + |v| + |\overline{\Gamma}| + 2|\Sigma_1| + 2|\Sigma_2| + 2|\Delta_0''| + |\Psi_{7a}| + |\Delta_1| + \\
& |\Phi_5| + |\Phi_6|. \\
d_2 = \quad & |r(4) \setminus r(2)| + |\Delta_{2,\ldots,m}| + |\Omega_1| + |\Omega_2| + |\Gamma| + |\Psi_{7b}| + 2|\Sigma_1| + 2|\Sigma_3| + 2|\Phi_5| + 2|\Psi_{7a}| + \\
& |\Sigma_2| + |\Sigma_4| + |\Phi_6| + |\Delta_1|. \\
i = \quad & |g| + |r(3)| + |\Sigma_1| + |\Sigma_2| + |\Sigma_3| + |\Sigma_4| + |\Delta_0'| + |\Delta_0''| + |\Delta_1| + |\Delta_{2,\ldots,m}| + |\Omega_1| + |\Omega_2| + \\
& |\Gamma| + |\overline{\Gamma}| + |w|. \\
i = \quad & |g| + |r(2)| + |v| + |\Delta_0'| + |\Delta_0''| + |\Delta_1| + |\Delta_{2,\ldots,m}| + |\Omega_1| + |\Omega_2| + |w| + |\Sigma_1| + \\
& |\Sigma_3| + |\Phi_5| + |\Psi_{7a}| + |\Psi_{7b}|.
\end{aligned}$$

It can be verified that the overall sum is exactly $2d$ where as in Lemma 7.1,

$$\begin{aligned}
d = \quad & |g| + |r| + |v| + (2|\Delta_0'| + 2|\Delta_0''| + 2|\Delta_1| + 2|\Delta_{2,\ldots,m}| + 2|\Omega_1| + 2|\Omega_2|) + (|\Gamma| + |\overline{\Gamma}|) + |w| + \\
& (3|\Sigma_1| + 2|\Sigma_2| + 2|\Sigma_3| + |\Sigma_4| + 2|\Phi_5| + |\Phi_6| + |\Psi_{7a}| + |\Psi_{7b}|).
\end{aligned}$$

One notes that since every element of $r = r(1) \cup r(2) \cup r(3) \cup r(4)$ is contained in precisely three of these sets, $|r| = |r(3)| + |r(1) \setminus r(3)| = |r(2)| + |r(4) \setminus r(2)|$. Also as emphasized before $|\Psi_{7a}| = |\Delta_0'|$. □

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A Auxiliary Lemmas

Lemma A.1. *Suppose A, B, C are three spaces such that $A \cap B = \{0\}$ and $C \subseteq A \oplus B$. Then sets of vectors can be chosen in the following manner:*

- $a_1, \ldots, a_p, a'_1, \ldots, a'_r$ *are in* A *and are linearly independent.*
- $b_1, \ldots, b_q, b'_1, \ldots, b'_r$ *are in* B *and are linearly independent.*
- $(a_1, \ldots, a_p, b_1, \ldots, b_q, a'_1 + b'_1, \ldots, a'_r + b'_r)$ *is a basis for* C.

Moreover:

- *If in addition,* $A \subseteq B \oplus C, B \subseteq A \oplus C$,
  - $a_1, \ldots, a_p, a'_1, \ldots, a'_r$ *is already a basis for* A.
  - $b_1, \ldots, b_q, b'_1, \ldots, b'_r$ *is already a basis for* B.
- *Otherwise, the sets can (clearly) be extended further so that*
  - $a_1, \ldots, a_p, a'_1, \ldots, a'_r, a''_1, \ldots, a''_m$ *is a basis for* A.
  - $b_1, \ldots, b_q, b'_1, \ldots, b'_r, b''_1, \ldots, b''_n$ *is a basis for* B.

Proof. Let $(a_1, \ldots, a_p)$ be a basis for $A \cap C$ and $(b_1, \ldots, b_q)$ be a basis for $B \cap C$. Let $c_1, \ldots, c_r \in C$ be such that $(a_1, \ldots, a_p, b_1, \ldots, b_q, c_1, \ldots, c_r)$ is a basis for C. Since $C \subseteq A \oplus B$, $c_j = a'_j + b'_j$ for some $a'_j \in A, b'_j \in B$.

We now prove that $a_1, \ldots, a_p, a'_1, \ldots, a'_r$ are linearly independent. Suppose (on the contrary) that for some index sets $\Phi \subseteq \{1, \ldots, p\}$ and $\Psi \subseteq \{1, \ldots, r\}$, we have $\bigoplus_{j\in\Phi} a_j \bigoplus_{j\in\Psi} a'_j = 0$. Consider

$$v = \bigoplus_{j\in\Phi} a_j \bigoplus_{j\in\Psi} (a'_j + b'_j) = \left(\bigoplus_{j\in\Phi} a_j \bigoplus_{j\in\Psi} a'_j \right) \bigoplus_{j\in\Psi} b'_j = \bigoplus_{j\in\Psi} b'_j.$$

Thus we have $v \in C$ as well as $v \in B$ and hence $v \in B \cap C$. Therefore $v = \bigoplus_{j\in\Sigma} b_j$ for some index set $\Sigma \subseteq \{1, \ldots, q\}$ and substituting above

$$\bigoplus_{j\in\Phi} a_j \bigoplus_{j\in\Sigma} b_j \bigoplus_{j\in\Psi} (a'_j + b'_j) = 0.$$

This contradicts, unless $\Phi = \Sigma = \Psi = \emptyset$, the assumption that $a_1, \ldots, a_p, b_1, \ldots, b_q, a'_1 + b'_1, \ldots, a'_r + b'_r$ is a basis for C and hence linearly independent.

Finally, we show that if $A \subseteq B \oplus C$, then $a_1, \ldots, a_p, a'_1, \ldots, a'_r$ is in fact a basis for A. Indeed, consider any $a \in A$. Since $A \subseteq B \oplus C$, $a = b + c$ for some $b \in B, c \in C$. We write c in the basis for C as $\bigoplus_{j\in\Phi} a_j \bigoplus_{j\in\Sigma} b_j \bigoplus_{j\in\Psi} (a'_j + b'_j)$ and hence

$$a = b \bigoplus_{j\in\Phi} a_j \bigoplus_{j\in\Sigma} b_j \bigoplus_{j\in\Psi} (a'_j + b'_j).$$

Since $A \cap B = \{0\}$, it follows that $a = \bigoplus_{j\in\Phi} a_j \bigoplus_{j\in\Psi} a'_j$. □

Lemma A.2. *Suppose A, Y, Z are independent spaces and $W \subseteq A \oplus Y \oplus Z$. Then there is a basis for W of the following form $\cup_{s=1}^{7} A_s$ where*

$$\begin{array}{lllll}
A_1 = \{ & a_i + y_j + z_k & \mid & i \in \Sigma_1, \quad j \in \Phi_1, & k \in \Psi_1 \} \\
A_2 = \{ & a_i + y_j & \mid & i \in \Sigma_2, \quad j \in \Phi_2 & \} \\
A_3 = \{ & a_i \quad + z_k & \mid & i \in \Sigma_3, & k \in \Psi_3 \} \\
A_4 = \{ & a_i \quad + \sigma & \mid & i \in \Sigma_4 & \} \\
A_5 = \{ & y_j + z_k & \mid & j \in \Phi_5, & k \in \Psi_5 \} \\
A_6 = \{ & y_j & \mid & j \in \Phi_6 & \} \\
A_7 = \{ & z_k & \mid & & k \in \Psi_7 \}
\end{array}$$

and we have

- $\{a_i \mid i \in \Sigma_1 \cup \Sigma_2 \cup \Sigma_3 \cup \Sigma_4\}$ *are linearly independent vectors in A.*
- $\{y_j \mid j \in \Phi_1 \cup \Phi_2 \cup \Phi_5 \cup \Phi_6\}$ *are linearly independent vectors in Y.*
- $\{z_k \mid k \in \Psi_1 \cup \Psi_3 \cup \Psi_5 \cup \Psi_7\}$ *are linearly independent vectors in Z.*
- *The σ are arbitrary linear forms in $\{y_j \mid j \in \Phi_1 \cup \Phi_5\}$, not necessarily all same.*

Proof. We start choosing a basis for $W \subseteq A \oplus Y \oplus Z$, picking one vector at a time, and adding it to $S = \cup_{s=1}^{3} A_s \cup_{s=5}^{7} A_s$ as below. We note that we do not add vectors to A_4 yet (this will be done after the process below ends):

Initialize $A_1 = A_2 = A_3 = A_5 = A_6 = A_7 = \emptyset$. $\quad S = \cup_{s=1}^{3} A_s \cup_{s=5}^{7} A_s = \emptyset$.

Initialize $\Sigma = \Sigma_1 \cup \Sigma_2 \cup \Sigma_3 \cup \Sigma_4 = \emptyset$. $\Phi = \Phi_1 \cup \Phi_2 \cup \Phi_5 \cup \Phi_6 = \emptyset$. $\Psi = \Psi_1 \cup \Psi_3 \cup \Psi_5 \cup \Psi_7 = \emptyset$.

Initialize $i^* = j^* = k^* = 1$.

Repeat as long as possible:

- Pick a vector $w \in W$, if possible, that fits any of the six cases below.
- If $w = a + y + z$ where $a \notin \mathsf{Span}\{a_i | i \in \Sigma\}, y \notin \mathsf{Span}\{y_j | j \in \Phi\}, z \notin \mathsf{Span}\{z_k | k \in \Psi\}$, then
  - Let $a_{i^*} = a,\ y_{j^*} = y,\ z_{k^*} = z$.
  - Add $w = a_{i^*} + y_{j^*} + z_{k^*}$ to A_1 as well as S.
  - add i^* to Σ and Σ_1, j^* to Φ and Φ_1, k^* to Ψ and Ψ_1.
  - Increment i^*, j^*, k^* each.
- $\cdots\cdots$ 5 **more similar cases** $\cdots\cdots$

We hope that the process is clear to the reader. The sets of vectors and indices grow as the process continues. The six cases correspond to the six types of $w : a + y + z, a + y, a + z, y + z, y, z$, which are added to $A_1, A_2, A_3, A_5, A_6, A_7$ respectively. In each case, we pick the vector w only if each of its components is linearly independent of vectors of the same "kind" that have already been "used" before (i.e. those indexed in Σ, Φ, Ψ respectively). The indices i^*, j^*, k^* are the next available indices. The sets Σ, Φ, Ψ maintain all the indices used so far (of the three kinds respectively).

We assume henceforth that the process above has ended. Let $\mathsf{Span}(S) \subseteq W$ be the span of all the vectors chosen so far. A small modification of the process above ensures that $(W \cap (Y \oplus Z)) \subseteq \mathsf{Span}(S)$. This is simply by considering the vectors in W in the order

$$W \cap Y, \quad W \cap Z, \quad W \cap (Y \oplus Z), \quad \mathsf{rest},$$

and using Lemma A.1. Hence we may assume henceforth that $(W \cap (Y \oplus Z)) \subseteq \mathsf{Span}(S)$.

We now finish the argument by completing the basis for W and showing that every vector remaining in $W \setminus \mathsf{Span}(S)$ is of A_4-type (possibly after adding a vector in $\mathsf{Span}(S)$). Indeed let $w = a + y + z$ be any "remaining" vector in $W \setminus \mathsf{Span}(S)$. We observe that:

- It must be that $a \notin \mathsf{Span}\{a_i | i \in \Sigma\}$. This is because, otherwise we can cancel out a by adding back appropriate vectors in S. This would result in a vector in $W \cap (Y \oplus Z) \subseteq \mathsf{Span}(S)$, a contradiction.
- It must be that $y \in \mathsf{Span}\{y_j | j \in \Phi\}$ as well as $z \in \mathsf{Span}\{z_k | k \in \Psi\}$. This is because, otherwise we can keep the one (or both) for which this condition fails and cancel out the other (if any) by adding back appropriate vectors in S. This would result in a vector of the type $a + y + z$ or $a + y$ or $a + z$, contradicting the end of the above process.
- Finally, we can cancel out z as well as "part of y that occurs in $\{y_j | j \in \Phi_2 \cup \Phi_6\}$" by adding back appropriate vectors in S.

□

Lemma A.3. *Suppose P, Q are independent spaces, $\mathsf{dim}(P) = \mathsf{dim}(Q) = n$, and $W \subseteq P \oplus Q$. Suppose moreover that $p_1, \ldots, p_n$ and $q_1, \ldots, q_n$ are given as bases of P, Q respectively. Then there is an $n \times n$ invertible matrix M such that after a change of basis (reusing the names)*

$$(p_1, \ldots, p_n) \leftarrow M(p_1, \ldots, p_n), \qquad (q_1, \ldots, q_n) \leftarrow M(q_1, \ldots, q_n),$$

there is a partition of the index set $\{1, \ldots, n\} = \Delta_0 \cup \Delta_1 \cup \Delta_2 \cup \ldots \cup \Delta_m \cup \Omega_1 \cup \Omega_2$ and a basis for W of the form:

$$\{p_i + \sigma(q) \mid i \in \Delta_1\} \cup B_2 \cup \ldots \cup B_m \cup C_1 \cup C_2,$$

$$B_k = \left\{ q_j + \sigma(p_{\Delta_{[k+1:m]}}) \;\middle|\; j \in \Delta_k \right\}$$

where $\Delta_{[k+1:m]} = \Delta_{k+1} \cup \ldots \cup \Delta_m \cup \Omega_1 \cup \Omega_2$,

$$C_1 = \{ p_i + \sigma(q_{\Omega_1}) \mid i \in \Omega_1 \cup \Omega_2 \},$$

$$C_2 = \{ q_j \mid j \in \Omega_2 \}.$$

We recall that $\sigma(\cdot)$ are arbitrary linear forms in respective variables (not necessarily the same).

Proof. The proof is iterative. Let W_P, W_Q denote the projections of W onto P and Q respectively, i.e.

$$W_P = \{p \in P \mid \exists q \in Q,\ p + q \in W\},$$
$$W_Q = \{q \in Q \mid \exists p \in P,\ p + q \in W\}.$$

It is clearly possible to choose matched bases $(p_1, \ldots, p_d, p_{d+1}, \ldots, p_n)$ and $(q_1, \ldots, q_d, q_{d+1}, \ldots, q_n)$ for P and Q respectively such that

$$\begin{array}{lllll} W_P = \mathsf{Span\ of} & p_1, \ldots, p_s, & & & p_{t+1}, \ldots, p_d \\ W_Q = \mathsf{Span\ of} & q_1, \ldots, q_s, & q_{s+1}, \ldots, q_t & & \end{array}$$

The "unused" indices $\{d+1, \ldots, n\}$ are placed in Δ_0. We choose arbitrary linear forms $\sigma_i(q)$ so that

$$p_{t+1} + \sigma_{t+1}(q), \ldots, p_d + \sigma_d(q) \in W.$$

These vectors are added to a partial basis for W and the indices $\{t+1, \ldots, d\}$ are added to Δ_1. Letting $W' = W \cap \mathsf{Span}(p_1, \ldots, p_t, q_1, \ldots, q_t)$, clearly

$$W = W' \oplus \mathsf{Span}(p_{t+1} + \sigma_{t+1}(q), \ldots, p_d + \sigma_d(q)).$$

Since the latter are already added to a partial basis, we only need to find a further basis for W'. Moreover, our index-space is now reduced to $\{1, \ldots, t\}$ and we can continue iteratively. This process makes progress unless $p_{t+1}, \ldots, p_d$ are absent, i.e. if

$$\begin{array}{lll} W_P = \mathsf{Span\ of} & p_1, \ldots, p_s, & \\ W_Q = \mathsf{Span\ of} & q_1, \ldots, q_s, & q_{s+1}, \ldots, q_t. \end{array}$$

In this case, the iterative process is stopped, and we begin a new iterative process. We set $m = 2$ and choose

$$q_{s+1} + \sigma(p_{\{1,\ldots,s\}}), \ldots, q_t + \sigma(p_{\{1,\ldots,s\}}) \in W.$$

These vectors are added to a partial basis for W letting $\Delta_m = \{s+1, \ldots, t\}$. We increment m by one and iterate the process on $W' = W \cap \mathsf{Span}(p_1, \ldots, p_s, q_1, \ldots, q_s)$. We note that $W'_Q = \mathsf{Span}(q_1, \ldots, q_s)$ while $W'_P \subseteq \mathsf{Span}(p_1, \ldots, p_s)$ (it may shrink to a proper subspace). After appropriate change of matched basis, $W'_P = \mathsf{Span}(p_1, \ldots, p_r)$ for $r \leqslant s$. This process makes progress unless we have

$$\begin{array}{lll} W_P = \mathsf{Span\ of} & p_1, \ldots, p_s, & \\ W_Q = \mathsf{Span\ of} & q_1, \ldots, q_s. & \end{array}$$

At this point, a basis for W is completed by first taking elements

$$p_1 + \sigma_1(q_{\{1,\ldots,s\}}), \ldots, p_s + \sigma_s(q_{\{1,\ldots,s\}}) \in W,$$

and adding to it $q_{r+1}, \ldots, q_s \in W \cap Q$. We can eliminate dependency of the former on the latter by elimination. The proof is completed by setting $\Omega_1 = \{1, \ldots, r\}$ and $\Omega_2 = \{r+1, \ldots, s\}$. □

Lemma A.4. *Let $X_1, \ldots, X_n$ be uniformly and independently distributed variables over $\{0,1\}^k$. Let*

$$\lambda_i \left(Y_{ij} \mid 1 \leqslant j \leqslant s_i \right), \ \ 1 \leqslant i \leqslant m,$$

be real-valued functions of its arguments where:

- *Each $Y_{ij} = X_r$ for some $r \in \{1, \ldots, n\}$.*
- *In the collection $\{Y_{ij} \mid 1 \leqslant i \leqslant m, 1 \leqslant j \leqslant s_i\}$, each X_r appears exactly twice, as $Y_{i'j'}$ and $Y_{i''j''}$ for $i' \neq i''$.*

- *It (therefore) holds that* $\sum_{i=1}^{m} s_i = 2n$.

Then

$$\mathop{\mathbb{E}}_{X_1,\ldots,X_n}\left[\prod_{i=1}^{m}|\lambda_i(Y_{i1},\ldots,Y_{is_i})|\right] \leqslant \prod_{i=1}^{m}\sqrt{\mathop{\mathbb{E}}_{Y_{i1},\ldots,Y_{is_i}}\left[\lambda_i^2(Y_{i1},\ldots,Y_{is_i})\right]}.$$

Proof. By induction. For $n = 1$, the only scenario and its proof is Cauchy-Schwartz:

$$\mathop{\mathbb{E}}_{X_1}\left[|\lambda_1(X_1)\lambda_2(X_1)|\right] \leqslant \sqrt{\mathop{\mathbb{E}}_{X_1}\left[\lambda_1^2(X_1)\right]}\sqrt{\mathop{\mathbb{E}}_{X_1}\left[\lambda_2^2(X_1)\right]}.$$

Otherwise, we assume w.l.o.g. that $n \geqslant 2$ and $Y_{11} = Y_{21} = X_n$. Applying Cauchy-Schwartz on X_n,

$$\mathop{\mathbb{E}}_{X_1,\ldots,X_n}\left[\prod_{i=1}^{m}|\lambda_i(Y_{i1},\ldots,Y_{is_i})|\right] \leqslant \mathop{\mathbb{E}}_{X_1,\ldots,X_{n-1}}\left[\sqrt{\mathop{\mathbb{E}}_{X_n}\left[\lambda_1^2(X_n,\cdot)\right]}\sqrt{\mathop{\mathbb{E}}_{X_n}\left[\lambda_2^2(X_n,\cdot)\right]}\prod_{i=3}^{m}|\lambda_i(Y_{i1},\ldots,Y_{is_i})|\right].$$

We get a further desired upper bound by induction hypothesis applied to functions $\lambda_1', \lambda_2', \lambda_3, \ldots, \lambda_m$ where

$$\lambda_1'(Y_{12},\ldots,Y_{1s_1}) = \mathop{\mathbb{E}}_{X_n}\left[\lambda_1^2(X_n, Y_{12},\ldots,Y_{1s_1})\right], \qquad \lambda_2'(Y_{22},\ldots,Y_{2s_2}) = \mathop{\mathbb{E}}_{X_n}\left[\lambda_2^2(X_n, Y_{12},\ldots,Y_{1s_1})\right].$$

□

Lemma A.5. *Suppose* $x \in \{0,1\}^k$ *is a uniformly distributed input and* $A, B \subseteq \{1,\ldots,k\}$ *such that* $A \cup B = \{1,\ldots,k\}$. *Let* x_A, x_B *denote the restricted input to* A *and* B *respectively. Suppose* λ, ψ *are non-negative functions of* x_A *and* x_B *respectively. Then*

$$\mathop{\mathbb{E}}_{x}\left[\lambda(x_A)\psi(x_B)\right] \leqslant \left(\max_{x_{A\cap B}} \mathop{\mathbb{E}}_{x_{A\setminus B}}\left[\lambda(x_A)\right]\right) \cdot \mathop{\mathbb{E}}_{x_B}\left[\psi(x_B)\right].$$

Proof. This is self-evident. Suppose β is the maximum above. Then

$$\mathop{\mathbb{E}}_{x}\left[\lambda(x_A)\psi(x_B)\right] = \mathop{\mathbb{E}}_{x_{A\cap B}}\left[\mathop{\mathbb{E}}_{x_{A\setminus B}}\left[\lambda(x_A)\right] \cdot \mathop{\mathbb{E}}_{x_{B\setminus A}}\left[\psi(x_B)\right]\right] \leqslant \beta \cdot \mathop{\mathbb{E}}_{x_{A\cap B}}\left[\mathop{\mathbb{E}}_{x_{B\setminus A}}\left[\psi(x_B)\right]\right] = \beta \cdot \mathop{\mathbb{E}}_{x_B}\left[\psi(x_B)\right].$$

□

B Significance of the 2-to-2 Games Theorem

In this section, we briefly summarize the main implications of the 2-to-2 Games Theorem (with imperfect completeness; some of these implications depend on its specific proof obtained in the present and previous works). We denote by ε a constant that can be taken as arbitrarily small.

- **Hardness Results:**

  The following results were already known based on the 2-to-2 Games Conjecture (as indicated in the reference; perfect completeness in the last two results if 2-to-2 Games Conjecture holds with perfect completeness):

- [16]: Gap Max Cut$\left(\frac{1}{2}+\Omega(\varepsilon), \frac{1}{2}+\frac{\varepsilon}{\log(1/\varepsilon)}\right)$ is NP-hard. This is optimal up to the constant in the Ω-notation.
- [10]: Gap Independent Set$\left(1-\frac{1}{\sqrt{2}}-\varepsilon,\ \varepsilon\right)$ is NP-hard and as a corollary, Vertex Cover is NP-hard to approximate within a factor strictly less than $\sqrt{2}$ (the latter is an improvement over the 1.36 hardness in [6]). Between these two implications, the "correct gap-location" (arbitrarily low soundness) for the Independent Set problem is more interesting and fundamental.
- [5]: It is NP-hard to distinguish whether a graph has four disjoint independent sets of (relative) size $\frac{1}{4}-\varepsilon$ each (and hence is almost 4-colorable) or whether there is no independent set of (relative) size ε (and hence is not almost $\left(\frac{1}{\varepsilon}\right)$-colorable).
- [8]: It is NP-hard to properly color (using k colors) more than a fraction $1-\frac{1}{k}+O(\frac{\ln k}{k^2})$ edges of an almost k–colorable graph. This is optimal up to the constant in the O-notation.

- **(Lasserre) Integrality Gaps with Perfect Completeness:**
  - If one concerns integrality gap (say up to a polynomial number of rounds of the Lasserre relaxation), the previous result for graph coloring holds with perfect completeness. I.e. there is a graph along with an SDP solution such that (a) the SDP solution pretends as if the graph is 4-colorable whereas (b) in actuality, the graph has no independent set of size ε.
  - Integrality gap (say up to a polynomial number of rounds of the Lasserre relaxation) for the 2-to-2 Games problem holds with perfect completeness and soundness ε.
  - These results are a consequence of the integrality gap known for the 3Lin problem with perfect completeness [7, 18] and the fact that the proof of the 2-to-2 Games Theorem is a reduction from 3Lin. The integrality gap instance for 3Lin can be "translated" via the reduction.

- **Evidence towards the Unique Games Conjecture:**

  GapUG$(\frac{1}{2}-\varepsilon,\varepsilon)$ is NP-hard, i.e. a weaker form of the Unique Games Conjecture holds with completeness $\approx \frac{1}{2}$. As far as the authors know (and we have consulted the algorithmic experts), the known algorithmic attacks on the Unique Game problem work equally well whether the completeness is ≈ 1 or whether it is $\approx \frac{1}{2}$. Thus, the implication that GapUG$(\ \cdot\ ,\ \varepsilon)$ is NP-hard with completeness $\approx \frac{1}{2}$ is a compelling evidence, in our opinion, that the known algorithmic attacks are (far) short of disproving the Unique Games Conjecture.

- **Unique Games Conjecture versus the Small Set Expansion Conjecture:**

  Raghavendra and Steurer [17] proposed the Small Set Expansion Conjecture and showed that it implies the Unique Games Conjecture. Roughly speaking, it states that GapSSE$(\varepsilon, 1-\varepsilon)$, the problem of distinguishing whether a graph has a small set of expansion at most ε or whether every small set has expansion at least $1-\varepsilon$, is NP-hard.

  The 2-to-2 Games Theorem arguably supports the (first author's) suspicion that the Unique Games Conjecture may be correct while the Small Set Expansion Conjecture may be incorrect. An informal reasoning is as follows.

  Raghavendra and Steurer give a reduction from GapSSE$[\varepsilon, 1-\varepsilon]$ to GapUG$[1-\varepsilon', \varepsilon']$. The same reduction also shows that GapSSE$[\beta, 1-\varepsilon]$ reduces to GapUG$[\approx \frac{1}{2}, \varepsilon']$ for some absolute constant β (say $\beta = \frac{3}{4}$). If one were to show that the latter problem is NP-hard without concluding anything

about the former, that may support the (first author's) suspicion. Indeed, this is precisely what happens in the proof of the 2-to-2 Games Theorem. One gets a reduction to $\mathsf{GapUG}[\approx \frac{1}{2}, \varepsilon']$ without getting a reduction to Gap SSE; the graphs in the reduction *always* have small non-expanding sets.

C Grassmann Graphs to the 2-to-2 Games Theorem

We summarize the chain of implications from the Grassmann graphs to the 2-to-2 Games Theorem. The chain is roughly:

$$\text{Grassmann Expansion Hypothesis} \overset{[4]}{\implies} \text{Linearity Testing Hypothesis} \overset{[2]}{\implies} \text{2 to 2 Games Conjecture.} \overset{[15,\ 3]}{}$$

- The Grassmann graphs and their potential application to the 2-to-2 Games problem were proposed in [15]. A key ingredient therein was a certain linearity testing primitive based on the Grassmann graph. Roughly speaking, in [15], the authors proposed a Weak Linearity Testing Hypothesis and showed that it implied a Weak 2-to-2 Games Conjecture. We do not elaborate on the qualifier "weak" here. It refers to seemingly unnatural variants that are nevertheless quite natural as far as application to Independent Set and Vertex Cover is concerned, which was the main motivation in [15].

- In [3], the authors formulated a Linearity Testing Hypothesis and showed that it implied the 2-to-2 Games Conjecture (with imperfect completeness).

- In [15, 3], it was already clear that the connectivity and expansion properties of the Grassmann graph would be crucial towards proving the Linearity Testing Hypotheses therein. In [4], the authors proposed (let's call it) Grassmann Expansion Hypothesis (stated as Theorem 1.8 in the present paper), and argued that it would at least be necessary towards proving the Linearity Testing Hypothesis. The authors presented a Fourier analytic framework and a preliminary set of results (for the first and second Fourier levels) towards proving the Grassmann Expansion Hypothesis.

- Barak, Kothari, and Steurer [2] proved that the Grassmann Expansion Hypothesis (almost immediately) implies the Linearity Testing hypothesis. While simple, this link is nevertheless important and was missed by the authors of [4].

- Finally, the Grassmann Expansion Hypothesis is proved in the present paper, stated as Theorem 1.8.

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