

Locally Testable Non-Malleable Codes*

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Abstract

In this work we adapt the notion of non-malleability for codes of Dziembowski, Pietrzak and Wichs (ICS 2010) to locally testable codes. Roughly speaking, a locally testable code is non-malleable if any tampered codeword which passes the local test with good probability is close to a valid codeword which either encodes the original, or an unrelated message.

We instantiate our definition by proving that a Reed-Muller-type code is non-malleable in the following sense: any adversary who independently tampers the coordinates of the code so that the tampered code passes the test with good probability, is tampering the underlying polynomial according to an affine transformation.

To the best of our knowledge, prior to this work, polynomial codes were not known to possess any non-malleability guarantees. Our analysis builds on the sampler-based decoding techniques common to several recent works.

1 Introduction

1.1 Locally Testable Codes and Non-Malleable Codes

A *coding scheme* is a pair (Enc, Dec) of functions $\text{Enc} : \Gamma^k \rightarrow \Gamma^n$ (possibly randomized) and $\text{Dec} : \Gamma^n \rightarrow \Gamma^k \cup \{\perp\}$ such that $\text{Dec}(\text{Enc}(m)) = m$ holds with probability 1 for all $m \in \Gamma^k$. We say $\mathbf{x} \in \Gamma^n$ is a *valid codeword* if $\mathbf{x} = \text{Enc}(m)$ for some $m \in \Gamma^k$ (and some choice of randomness for Enc). The quantity k/n is called the *rate* of the code. Given $\mathbf{x}, \mathbf{y} \in \Gamma^n$, the *distance* between \mathbf{x} and \mathbf{y} is $\Pr_{i \sim [n]}[\mathbf{x}_i \neq \mathbf{y}_i]$. The *distance of the code* is the minimum distance between any two distinct valid codewords. When a code's distance is bounded away from zero, one can try to design decoding-type algorithms with extra features such as error-correction or local-decoding/testing capabilities. In this paper, we will work with codes which admit a local testing algorithm.

Definition 1 (Local Testing). *Given a code (Enc, Dec) , $q \in \mathbb{N}$, and $\varepsilon_0, c > 0$, a (q, ε_0, c) -local tester, Test , is a randomized algorithm which takes $\mathbf{x} \in \Gamma^n$ as input, chooses $I = (i_1, \dots, i_q)$, a q -tuple of elements in $[n]$, reads the q symbols $\mathbf{x}_I = (\mathbf{x}_{i_1}, \dots, \mathbf{x}_{i_q})$ from \mathbf{x} and outputs a bit. Moreover,*

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- **Correctness:** if $\mathbf{x} \in \Gamma^n$ is a valid codeword then $\text{Test}(\mathbf{x}) = 1$ with probability 1;
- **Soundness:** for all $\mathbf{x} \in \Gamma^n$ such that $\varepsilon := \Pr_I[\text{Test}(\mathbf{x}; I) = 1] \geq \varepsilon_0$, there exists a valid codeword $\mathbf{y} \in \Gamma^n$ within distance $c \cdot \varepsilon$ of \mathbf{x} .

We say that (Enc, Dec) is a (q, ε_0) -locally testable code (LTC) if it has a (q, ε_0, c) -local tester for a constant $c > 0$. The intuitive interpretation of soundness is: if $\mathbf{x} \in \Gamma^n$ is such that $\text{Test}(\mathbf{x}) = 1$ with good probability, then \mathbf{x} is close to a valid codeword. List decoding for LTCs refers to the stronger guarantee: for any $\mathbf{x} \in \Gamma^n$, there is a short list of valid codewords which explain nearly all of $\text{Test}(\mathbf{x})$'s acceptance probability.

Definition 2 (List-Decoding for LTCs). Fix $\ell \in \mathbb{N}$ and $\varepsilon > 0$. A locally testable code is said to be (ℓ, ε) -list-decodable if for all $\mathbf{x} \in \Gamma^n$ there exists a set $L_{\mathbf{x}} \subset \Gamma^n$ of valid codewords such that $|L_{\mathbf{x}}| \leq \ell$ and

$$\Pr_I \left[\text{Test}(\mathbf{x}; I) = 1 \ \& \ \mathbf{x}_I \notin \{\mathbf{y}_I : \mathbf{y} \in L_{\mathbf{x}}\} \right] < \varepsilon.$$

Non-malleable codes [DPW18] (NMCs) provide meaningful security guarantees even in situations where error correction is impossible. Roughly speaking, (Enc, Dec) is *non-malleable* against the tampering function family $\mathcal{F} \subset \{f : \Gamma^n \rightarrow \Gamma^n\}$ if for all $f \in \mathcal{F}$ and $m \in \Gamma^k$, the distribution $(\text{Dec} \circ f \circ \text{Enc})(m)$ (over the randomness of Enc) is either equal to m (such is the case when f is the identity), or else is statistically independent of m . Non-malleable reductions [ADKO15] are useful relaxations which allow constructing non-malleable codes via concatenation. Intuitively, a non-malleable reduction from \mathcal{F} to \mathcal{G} guarantees that the tampering of codewords by functions in \mathcal{F} is captured by tampering messages by functions in \mathcal{G} . The key feature of non-malleable reductions is that they compose well. For example, if $(\text{Enc}_{\mathcal{F}}, \text{Dec}_{\mathcal{F}})$ is a non-malleable reduction from \mathcal{F} to \mathcal{G} and $(\text{Enc}_{\mathcal{G}}, \text{Dec}_{\mathcal{G}})$ is a non-malleable code against \mathcal{G} , then $(\text{Enc}_{\mathcal{F}} \circ \text{Enc}_{\mathcal{G}}, \text{Dec}_{\mathcal{G}} \circ \text{Dec}_{\mathcal{F}})$ is a non-malleable code against \mathcal{F} .

Definition 3 (Non-Malleable Reductions). Fix $\varepsilon > 0$ and tampering function families

$$\mathcal{F} \subset \{f : \Gamma^n \rightarrow \Gamma^n\} \text{ and } \mathcal{G} \subset \{g : \Gamma^k \rightarrow \Gamma^k \cup \{\perp\}\}.$$

We say that a coding scheme (Enc, Dec) is an ε -non-malleable reduction from \mathcal{F} to \mathcal{G} if for all $f \in \mathcal{F}$ there exists a distribution G_f on \mathcal{G} such that $\Delta((\text{Dec} \circ f \circ \text{Enc})(m), G_f(m)) \leq \varepsilon$ for all $m \in \Gamma^k$, where $G_f(m)$ is the distribution which draws $g \sim G_f$ and outputs $g(m)$ (Δ denotes statistical distance). A non-malleable code is a non-malleable reduction to the family of “trivial” tampering functions, containing only the identity and constants.

1.2 Non-Malleable, Locally-Testable Codes

Syntax for the Definition. Given a function $h : \Gamma^n \rightarrow (\Gamma \cup \{\perp\})^n$, and a distribution \mathcal{R} on Γ^n , we let $\mathcal{D}_h^{\mathcal{R}}$ denote the distribution which 1) draws $\mathbf{x} \sim \mathcal{R}$ and I according to Test ; 2) tampers to obtain $\tilde{\mathbf{x}} = h(\mathbf{x})$; 3) outputs $\tilde{\mathbf{x}}_I$ if $\text{Test}(\tilde{\mathbf{x}}; I) = 1$, \perp otherwise. When \mathcal{R} is the encoding distribution $\text{Enc}(m)$ for some $m \in \Gamma^k$, we denote the distribution by $\mathcal{D}_h(m)$ instead of $\mathcal{D}_h^{\text{Enc}(m)}$.

Definition 4 (Non-Malleability for LTCs). Fix parameters $\ell \in \mathbb{N}$, $\varepsilon > 0$, and function families

$$\mathcal{F} \subset \{f : \Gamma^n \rightarrow \Gamma^n\} \text{ and } \mathcal{G} \subset \{g : \Gamma^n \rightarrow (\Gamma \cup \{\perp\})^n\}.$$

We say that a LTC $(\text{Enc}, \text{Dec}, \text{Test})$ is an (ℓ, ε) -locally-testable non-malleable reduction from \mathcal{F} to \mathcal{G} if for all $f \in \mathcal{F}$, there exists $L_f = \{g^{(1)}, \dots, g^{(\ell)}\} \subset \mathcal{G}$ of size $|L_f| = \ell$ and a function $g : \Gamma^n \rightarrow (\Gamma \cup \{\perp\})^n$ such that:

1. for all $i \in [n]$ and $\mathbf{x} \in \Gamma^n$, $g(\mathbf{x})_i \in \{g^{(j)}(\mathbf{x})_i : g^{(j)} \in L_f\}$; and
2. there exists a distribution SIM on Γ^n such that for all $m \in \Gamma^k$, $\Delta(\mathcal{D}_f(m), \mathcal{D}_g^{\text{SIM}}) \leq \varepsilon$.

As before, if \mathcal{G} is the family of trivial tampering functions consisting just of the identity and constants, then (Enc, Dec) is called an (ℓ, ε) -locally-testable non-malleable code.

Remark. Some remarks are in order.

1. The list-decoding intuition is captured by the shortness of L_f : nearly all of the test passing probability of an f -tampered codeword is explained by f 's agreement with a short list of functions in \mathcal{G} . Note that each coordinate of g is a (possibly different) convex combination of the corresponding coordinates of the $g^{(j)}$. Non-malleability is captured by the fact that $\mathcal{D}_g^{\text{SIM}}$ does not depend on m .
2. Unlike Definition 3, the functions in \mathcal{G} in Definition 4 map codewords to codewords, rather than messages to messages. This modification is so that we can meaningfully compare $f(\mathbf{x})_I$ with $g(\mathbf{x})_I$, an important feature of local-testing definitions. The family $\text{Dec} \circ \mathcal{G} \circ \text{Enc}$ would be the corresponding distribution on message-to-message functions. In this work, \mathcal{G} will always be either the family of trivial tampering functions, or the family of affine tampering functions. In either case, $\text{Dec} \circ \mathcal{G} \circ \text{Enc}$ is also trivial or affine. The distribution G_f of Definition 3 outputs $g^{(j)}$ with probability proportional to the probability that $\mathcal{D}_f^{\text{SIM}}$ agrees with $g^{(j)}$.
3. Composing two standard non-malleable reductions – one from \mathcal{F} to \mathcal{G} , one from \mathcal{G} to \mathcal{H} – yields a non-malleable reduction from \mathcal{F} to \mathcal{H} . The same composition theorem does not hold generically for locally testable, non-malleable reductions. We use a non-generic composition theorem to combine a locally testable, non-malleable reduction from \mathcal{F} to \mathcal{G} with a non-malleable code against \mathcal{G} (for specific \mathcal{F} and \mathcal{G}) to obtain a locally testable, non-malleable code against \mathcal{F} . The test of our composed code involves locally decoding a symbol of the outer code so it can be checked for validity by the inner code. This idea is often used to compose locally testable codes and PCPs.

The following claim gives a useful set of sufficient conditions for a LTC being non-malleable. The simple proof is given in Appendix A. For the codes used in this work, the first three conditions will be more or less trivial to establish. Thus, Claim 1 essentially reduces proving non-malleability to the task of establishing condition 4. This will simplify our proofs considerably.

Claim 1 (Sufficient Conditions for Non-Malleability in LTCs). *Let $(\text{Enc}, \text{Dec}, \text{Test})$ be a LTC with $\text{Enc} : \Gamma^k \rightarrow \Gamma^n$, and let*

$$\mathcal{F} \subset \{f : \Gamma^n \rightarrow \Gamma^n\} \text{ and } \mathcal{G} \subset \{g : \Gamma^n \rightarrow (\Gamma \cup \{\perp\})^n\}$$

be function families. Suppose the following four conditions hold.

1. \mathcal{G} contains the constant “all \perp ” function; for all other $g \in \mathcal{G}$, and all valid codewords $\mathbf{x} \in \Gamma^n$, $\text{Test}(g(\mathbf{x}); I) = 1$ occurs with probability 1;
2. For all distinct $g, g' \in \mathcal{G}$ and $m \in \Gamma^k$, $\Pr_{\mathbf{x} \sim \text{Enc}(m), I} [g(\mathbf{x})_I = g'(\mathbf{x})_I] \leq \varepsilon/\ell^2$.
3. For all $f \in \mathcal{F}$, and $m, m' \in \Gamma^k$: $\Delta\left(\{f(\mathbf{x})_I\}_{\mathbf{x} \sim \text{Enc}(m), I}, \{f(\mathbf{x})_I\}_{\mathbf{x} \sim \text{Enc}(m'), I}\right) \leq \varepsilon$.
4. For all $f \in \mathcal{F}$ there exists a list $\mathcal{L}_f = \{g^{(1)}, \dots, g^{(\ell)}\} \subset \mathcal{G}$ of size $|\mathcal{L}_f| = \ell$ such that for all $m \in \Gamma^k$,

$$\Pr_{\mathbf{x} \sim \text{Enc}(m), I} \left[\text{Test}(f(\mathbf{x}); I) = 1 \ \& \ f(\mathbf{x})_I \notin \{g^{(j)}(\mathbf{x})_I : g^{(j)} \in \mathcal{L}_f\} \right] \leq \varepsilon.$$

Then $(\text{Enc}, \text{Dec}, \text{Test})$ is an $(\ell + 1, 3\varepsilon)$ -locally testable, non-malleable reduction from \mathcal{F} to \mathcal{G} .

1.3 Sampler Graphs

Notations. For a finite set S , $s \sim S$ indicates that s is drawn uniformly from S . For a bipartite graph $(A \cup B, E)$ and $a \in A$, $B(a)$ denotes the uniform distribution on the neighborhood of a in B : $\{b \in B : (a, b) \in E\}$. The neighborhood distribution $A(b)$ for $b \in B$ is defined analogously. For all bipartite graphs used in this work, the edge relations are natural. For example, A might be the set of lines in \mathbb{F}^k (\mathbb{F} a finite field), B the set of points in \mathbb{F}^k , and the edge relation captures incidence: $(a, b) \in E$ iff $b \in a$. For this reason, we simplify notations by supressing E and denoting bipartite graphs as A/B instead of $(A \cup B, E)$, and writing $a \sim b$ instead of $(a, b) \in E$.

Definition 5 (Biregularity). *Let A/B be a bipartite graph and fix $\eta > 0$. We say that A/B is η -biregular if the distribution which draws $a \sim A$, $b \sim B(a)$, and outputs (a, b) is within statistical distance η of the distribution which gives the same output by drawing $b \sim B$, $a \sim A(b)$.¹*

Biregularity as defined above ensures that for any $B' \subset B$ of size $|B'| = \lambda \cdot |B|$, the expectation (over $a \sim A$) of $\Pr_{b \sim B(a)} [b \in B']$ is close to λ . We say that A/B is *sampling* if, in addition, a concentration bound holds.

Definition 6 (Sampler Graph [Zuc97]). *Fix $\varepsilon, \delta > 0$. We say that the bipartite graph A/B is (ε, δ) -sampling if for all subsets $B' \subset B$ of size $|B'| = \lambda \cdot |B|$,*

$$\Pr_{a \sim A} \left[\left| \Pr_{b \sim B(a)} [b \in B'] - \lambda \right| > \varepsilon \right] \leq \delta.$$

¹This is related to the usual notion of biregularity; specifically, if A/B is biregular in the usual sense, then it is 0-biregular in the sense of Definition 5.

Double Samplers. A triple (A, B, C) is called a *double sampler* if B/C is sampling and for all $c \in C$, $A(c)/B(c)$ is sampling. Double samplers have been used implicitly in several works prior to their formalization in [DK17]. We use them implicitly in this work as well. The construction in [DK17] is of a double sampler of linear size (*i.e.*, $|A| \approx |B| \approx |C|$) based on high-dimensional expanders. The double samplers used in this work are built from elementary means and are not linear size (our double samplers have $|A| \gg |B| \gg |C|$). Importantly, a random object in our parameter regime is a double sampler with good probability, while this is not true in the linear size regime.

1.4 Our Code and Main Theorem

Notation. Let \mathbb{F} be a finite field, and let $k \geq 4$ and $d \geq 2$ be dimension and degree parameters, respectively. Let A be the set of affine 3-planes in \mathbb{F}^k , $C = \mathbb{F}^k$ and let the edge relation be incidence: $a \sim c$ iff $c \in a$. Let Γ and Γ_A be the sets of k -variate and 3-variate polynomials of degree at most d over \mathbb{F} , respectively.

Main Construction. With notations as above:

- **Enc(m):** For $m \in \mathbb{F}$, draw $\Phi \sim \Gamma$ such that $\Phi(\mathbf{0}) = m$ and output $\{\Phi|_a\}_{a \in A} \in \Gamma_A^{|A|}$. We will often write codewords as $\{(a, \alpha)\}_{a \in A}$ with the understanding that $\alpha = \Phi|_a$.
- **Dec($\{(a, \alpha)\}_{a \in A}$):** Given $\{(a, \alpha)\}_{a \in A}$, find $\Phi \in \Gamma$ such that $(a, \alpha) = (a, \Phi|_a)$ for all $a \in A$.² If such Φ exists, output $m = \Phi(\mathbf{0})$, otherwise output \perp .³
- **Test($\{(a, \alpha)\}_{a \in A}$):** Draw $c \sim C$, $a, a' \sim A(c)$; read (a, α) and (a', α') , and output 1 if $\alpha|_c = \alpha'|_c$ ($\alpha|_c$ denotes the evaluation of α at c), 0 otherwise.

The above code is known to be a $(2, |\mathbb{F}|^{-\Omega(1)})$ -locally testable code. This was proven originally in the influential works [AS97, RS97]. Our main theorem is that this code also possesses non-malleability guarantees. Before we state this formally, we introduce the tampering function families.

Tampering Function Families. We identify three types of tampering.

- **Coordinate-Wise:** $\mathcal{F} := \{\{f_a\}_{a \in A} \mid f_a : \Gamma_A \rightarrow \Gamma_A\}$ tampers codewords via

$$\{f_a\}_a : \{(a, \alpha)\}_a \mapsto \{(a, f_a(\alpha))\}_a.$$

- **Affine:** We say that $T : \Gamma \rightarrow \Gamma$ is *affine* if $\exists (s, \Phi_0) \in \mathbb{F} \times \Gamma$ such that $T(\Phi) = s \cdot \Phi + \Phi_0$. We define \mathcal{G} to be the family of coordinate-wise restrictions of global affine maps:

$$\mathcal{G} := \left\{ \{g_a\}_{a \in A} \mid \exists (s, \Phi_0) \in \mathbb{F} \times \Gamma \text{ st } g_a(\alpha) = s \cdot \alpha + \Phi_0|_a \forall a \in A \right\} \subset \mathcal{F}.$$

²Such Φ , if it exists, can be found in time $\text{poly}(|\mathbb{F}|)$ by interpolation.

³As written, decoding runs in time $\text{poly}(|\mathbb{F}|)$, which is exponential in the message length. However, local decoding algorithms exist which run in time $\text{poly}(\lambda, \log |\mathbb{F}|, 1/\delta)$ and output m (or a list containing m) with probability $1 - 2^{-\lambda}$ whenever the input is within distance δ of a valid encoding of m . See for example [Sud97].

- **Trivial:** We say that $T : \Gamma \rightarrow \Gamma$ is *trivial* if either $T(\Phi) = \Phi$ or if $\exists \Phi_0 \in \Gamma$ such that $T(\Phi) = \Phi_0$. We define \mathcal{H} to be the family of coordinate-wise restrictions of trivial maps:

$$\mathcal{H} := \left\{ \{h_a\}_{a \in A} \mid T(\Phi) = \Phi, \text{ or } \exists \Phi_0 \in \Gamma \text{ s.t. } T(\Phi) = \Phi_0|_a \forall a \in A \right\} \subset \mathcal{G}.$$

We also include the “all \perp function” (maps every coordinate to \perp) in \mathcal{G} and \mathcal{H} .

Theorem 1 (Main). *The code above is an (ℓ, ε) –locally-testable, non-malleable reduction from \mathcal{F} to \mathcal{G} where $\varepsilon = |\mathbb{F}|^{-\Omega(1)}$ and $\ell = 4/\varepsilon$.*

We use this locally testable, non-malleable reduction to build a locally testable, non-malleable code against \mathcal{F} . The explicit construction is given in section 6.

Theorem 2. *There exists an explicit (ℓ, ε) –locally testable, non-malleable code against \mathcal{F} , the family of coordinate-wise tampering functions where $\varepsilon = |\mathbb{F}|^{-\Omega(1)}$ and $\ell = 4/\varepsilon$.*

1.5 Prior Work

Sampler-Based Decoding. Our work fits into a recent line of work on sampler-based decoding [IKW12, Mos17, BDN17, DHK⁺19, DHKR19] (and many more). In these works, sampling properties of a code’s index set are exploited in order to give non-trivial decoding algorithms. Our work builds on techniques developed in these papers in order to “decode” a coordinate-wise tampering function which respects codeword proximity, to a small list of affine functions.

Non-Malleable Codes. Since the introduction of non-malleable codes in 2010 [DPW18], an immense research effort has focused on giving constructions which are secure against richer classes of tampering functions, and with better rate [DKO13, ADL14, ADKO15, CGL16, Li16, BDG⁺18] (and many, many more). Our work adapts some of the techniques developed in this area to the local-testing regime. We expect constructions with security against other types of tampering, and constructions with better rate are possible.

Locally Decodable Non-Malleable Codes. A few works combine the notions of local decodability with non-malleability [DLSZ15, CKR15]. These works give constructions of non-malleable codes which admit local decode/update subroutines. Our work differs in several ways from these. Most telling, is the fact that the codes in these works achieve high rate with super-constant locality, whereas our main construction achieves optimal locality with very poor rate. Moreover, our techniques differ significantly. The construction of [DLSZ15] had computational security; the second work [CKR15] showed how to replace the cryptographic primitives with information-theoretic variants in order to obtain statistical security. Our techniques on the other hand, are similar to those used in the LTC literature. We believe the following statement is fair and summarizes these differences: our work is aimed at inserting non-malleability into LTCs, while [DLSZ15, CKR15] are focused on inserting locality into NMCs.

2 The Affine Agreement Theorem

In this section we state the affine agreement theorem, which is at the core of the proof of Theorem 1. Theorem 1 follows from our affine agreement theorem in much the same way as list-decoding theorems often follow from agreement theorems. We begin by recalling, and adding some new notations.

Notation. Recall \mathbb{F} is a finite field, $k \geq 4$, $d \geq 2$, A is the set of affine 3-planes in \mathbb{F}^k , $C = \mathbb{F}^k$, forming the bipartite graph A/C where the edge relation is incidence: $a \sim c$ iff $c \in a$. Additionally, Γ and Γ_A are the sets of k -variate and 3-variate polynomials of degree at most d over \mathbb{F} , respectively. Also recall that \mathcal{F} denotes the family of coordinate-wise functions from $\Gamma_A^{|\mathcal{A}|} \rightarrow \Gamma_A^{|\mathcal{A}|}$; and we saw that $\{f_a\}_{a \in A} \in \mathcal{F}$ maps $\{(a, \alpha)\}_a \mapsto \{(a, \tilde{\alpha})\}_a$, where $\tilde{\alpha} = f_a(\alpha)$. We always use tildes to indicate images under various tampering functions. Now, let $\Gamma_C = \mathbb{F}$, and let $\bar{A} = A \times \Gamma_A$ and $\bar{C} = C \times \Gamma_C$. This forms another bipartite graph \bar{A}/\bar{C} where the edge relation is incidence and agreement: $(a, \alpha) \sim (c, \gamma)$ iff $c \in a$ and $\alpha|_c = \gamma$. We write \bar{a} and \bar{c} instead of (a, α) and (c, γ) .

Theorem 3 (Affine Agreement). *There exists $\varepsilon = |\mathbb{F}|^{-\Omega(1)}$ such that for all $\{f_a\} \in \mathcal{F}$, the following holds. If*

$$\Pr_{\Phi, (c, a, a')} [\tilde{\alpha}|_c = \tilde{\alpha}'|_c] \geq 6\varepsilon,$$

where the probability is over $\Phi \sim \Gamma$, $c \sim C$, $a, a' \sim A(c)$ and where $(\tilde{\alpha}, \tilde{\alpha}') = (f_a(\Phi|_a), f_{a'}(\Phi|_{a'}))$, then there exists an affine $T : \Gamma \rightarrow \Gamma$ such that $\Pr_{(\Phi, a) \sim \Gamma \times A} [\tilde{\alpha} = T(\Phi)|_a] \geq \varepsilon$.

Proof of Theorem 1 Assuming Theorem 3. Let ε be as in Theorem 3 above and fix $f = \{f_a\}_a \in \mathcal{F}$. We prove that the code is non-malleable by establishing the sufficient conditions of Claim 1. The first three are immediate.

1. Clearly \mathcal{G} contains the constant ‘all- \perp ’ function, and for all other $\{g_a\}_a \in \mathcal{G}$, the equality $g_a(\Phi|_a)|_c = g_{a'}(\Phi|_{a'})|_c$ holds for all $\Phi \in \Gamma$, and (c, a, a') .
2. For all distinct $\{g_a\}_a, \{g'_a\}_a \in \mathcal{G}$, $g_a(\Phi|_a) = g'_a(\Phi|_a)$ holds only if either $g_a = g'_a$ (occurs with probability $\mathcal{O}(|\mathbb{F}|^{-1})$) when $\{g_a\}_a \neq \{g'_a\}_a$, or if $g_a \neq g'_a$ but $g_a(\Phi|_a) = g'_a(\Phi|_a)$ (also probability $\mathcal{O}(|\mathbb{F}|^{-1})$). Thus, $\Pr_{\Phi, (c, a, a')} [g_a(\Phi|_a) = g'_a(\Phi|_a)] = \mathcal{O}(|\mathbb{F}|^{-1}) \ll \varepsilon/\ell^2$.
3. Note that for all (a, a') such that $\mathbf{0} \notin a \cup a'$, the distribution which draws $\Phi \sim \Gamma$ such that $\Phi(\mathbf{0}) = m$, and outputs $(\Phi|_a, \Phi|_{a'})$ is identical to the one which draws $\Phi \sim \Gamma$ and gives the same output. It follows that for all $m, m' \in \Gamma^k$,

$$\Delta \left(\left\{ (f_a(\Phi|_a), f_{a'}(\Phi|_{a'})) \right\}_{\substack{\Phi: \Phi(\mathbf{0})=m \\ (c, a, a')}} \right), \left\{ (f_a(\Phi|_a), f_{a'}(\Phi|_{a'})) \right\}_{\substack{\Phi: \Phi(\mathbf{0})=m' \\ (c, a, a')}} \right) \leq \Pr_{(c, a, a')} [\mathbf{0} \in a \cup a'],$$

which is $\mathcal{O}(|\mathbb{F}|^{-1}) \ll \varepsilon$.

For the last condition, we show that there exists $L_f \subset \mathcal{G}$ of size at most ℓ such that

$$\Pr_{\Phi, (c, a, a')} [\tilde{\alpha}|_c = \tilde{\alpha}'|_c \ \& \ (\tilde{\alpha}, \tilde{\alpha}') \notin \{(g_a(\alpha), g_{a'}(\alpha')) : \{g_a\}_a \in L_f\}] < 6\varepsilon, \quad (1)$$

where $(\tilde{\alpha}, \tilde{\alpha}') = (f_a(\alpha), f_{a'}(\alpha'))$ for $(\alpha, \alpha') = (\Phi|_a, \Phi|_{a'})$, and where $\Phi \sim \Gamma$.⁴ Towards this end, let $L_f := \{\{g_a\}_a \in \mathcal{G} : \Pr_{(\Phi, a) \sim \Gamma \times A}[\tilde{\alpha} = g_a(\alpha)] \geq \varepsilon/2\}$.

Small List Size. Assume for contradiction that $|L_f| \geq \ell = 4/\varepsilon + 1$, and so contains a set $\{\{g_a^1\}_a, \dots, \{g_a^\ell\}_a\}$. By inclusion-exclusion,

$$\begin{aligned} 1 &\geq \Pr_{(\Phi, a) \sim \Gamma \times A}[\tilde{\alpha} \in \{g_a^i(\alpha) : i = 1, \dots, \ell\}] \\ &\geq \frac{\ell \cdot \varepsilon}{2} - \sum_{1 \leq i < j \leq \ell} \Pr_{\Phi, a} [g_a^i(\alpha) = g_a^j(\alpha)] > 2 - \binom{\ell}{2} \cdot \left(\frac{1}{|\Gamma|} + \frac{d}{|\mathbb{F}|} \right). \end{aligned}$$

The last inequality used $\ell\varepsilon > 4$, and the bound on $\Pr_{\Phi, a} [g_a^i(\Phi|_a) = g_a^j(\Phi|_a)]$ from point 2 above. The right hand side simplifies to $2 - o(1) > 1$, a contradiction.

Proximity Implies List Decoding. Suppose $\{f_a\}$ is such that (1) does not hold. Define $\{f'_a\}_a \in \mathcal{F}$ as follows: $f'_a(\alpha) = f_a(\alpha)$, unless $f_a(\alpha) = g_a(\alpha)$ for some $\{g_a\}_a \in L_f$ in which case $f'_a(\alpha)$ outputs a random $\tilde{\alpha} \notin \{g_a(\alpha) : \{g_a\}_a \in L_f\}$. Note

$$\Pr_{\Phi, (c, a, a')} [f'_a(\alpha)|_c = f'_{a'}(\alpha')|_c] \geq 6\varepsilon$$

since (1) does not hold. Therefore, by Theorem 3, there exists an affine $\mathbb{T} : \Gamma \rightarrow \Gamma$ such that $\Pr_{\Phi, a} [f'_a(\Phi|_a) = \mathbb{T}(\Phi)|_a] \geq \varepsilon$. Thus $\Pr_{\Phi, a} [f_a(\Phi|_a) = \mathbb{T}(\Phi)|_a] \geq \varepsilon - \ell/|\Gamma_A| \geq \varepsilon/2$, and so the coordinate-wise version of \mathbb{T} is in L_f . This is a contradiction since by construction, for every $\{g_a\}_a \in L_f$, $f'_a(\alpha) \neq g_a(\alpha)$ holds for all $a \in A$ and $\alpha \in \Gamma_A$. \square

2.1 Reducing the NM Agreement Theorem to Two Lemmas

The proof of Theorem 3 will occupy much of the rest of this paper. In this section, we separate the proof into two parts by stating two lemmas which combine to immediately prove the theorem.

Proof of Theorem 3. Suppose $\varepsilon = |\mathbb{F}|^{-\Omega(1)}$ is chosen so it satisfies Lemmas 1 and 2, below. Let $\{f_a\}_a \in \mathcal{F}$ be such that

$$\Pr_{\Phi, (c, a, a')} [\tilde{\alpha}|_c = \tilde{\alpha}'|_c] \geq 6\varepsilon. \quad (2)$$

By Lemma 1 below, there exists a function $h : \bar{C} \rightarrow \Gamma_C$ such that

$$\Pr_{(a, \Phi) \sim A \times \Gamma} \left[\Pr_{\bar{c} \sim \bar{C}(\bar{a})} [\tilde{\alpha}|_c = \tilde{\gamma}] \geq 1 - \zeta \right] \geq 2\varepsilon, \quad (3)$$

where $\tilde{\gamma} = h(\bar{c})$, $\bar{a} = (a, \Phi|_a)$, and where $\zeta = |\mathbb{F}|^{-\Omega(1)}$ is specified precisely in Section 4. By Lemma 2, there exists an affine map $\mathbb{T} : \Gamma \rightarrow \Gamma$ such that

$$\Pr_{(a, \Phi) \sim A \times \Gamma} [\tilde{\alpha} = \mathbb{T}(\Phi)|_a] \geq \varepsilon. \quad (4)$$

\square

⁴as noted in point 3 above, the difference in probability caused by drawing $\Phi \sim \Gamma$ such that $\Phi(\mathbf{0}) = m$ instead is negligible.

Lemma 1 (Global Agreement). *There exists $\varepsilon = |\mathbb{F}|^{-\Omega(1)}$ such that whenever $\{f_a\}_a \in \mathcal{F}$ is such that (2) holds, there exists $h : \bar{C} \rightarrow \Gamma_C$ such that (3) holds.*

Lemma 2 (Affine Agreement). *There exists $\varepsilon = |\mathbb{F}|^{-\Omega(1)}$ such that whenever $\{f_a\}_a \in \mathcal{F}$ and $h : \bar{C} \rightarrow \Gamma_C$ are such that (3) holds, there exists an affine $T : \Gamma \rightarrow \Gamma$ such that (4) holds.*

Lemma 1 is proved in Section 4 using a sampler-based decoding argument similar to ones which have appeared in several recent works, for example [BDN17]. The linearity test analyzed in the proof of Lemma 2 in Section 5 is new to this work.

3 Sampler Graph Preliminaries

Sampler graphs play a big role in the proofs in the following sections. In this section we introduce the graphs whose sampling will be used, and various properties of sampler graphs. All of the graphs are what we call “incidence \times agreement” graphs, such as \bar{A}/\bar{C} from last section. We begin with some notation.

Notation. Recall \mathbb{F} is a finite field, $k \geq 4$, $d \geq 2$, A is the set of 3-planes in \mathbb{F}^k , $C = \mathbb{F}^k$, Γ and Γ_A are the sets of k -variate and 3-variate polynomials of degree at most d over \mathbb{F} , respectively, $\Gamma_C = \mathbb{F}$. This defines an incidence \times agreement bipartite graph \bar{A}/\bar{C} where $\bar{A} = A \times \Gamma_A$, $\bar{C} = C \times \Gamma_C$ and the edge relation is “incidence \times agreement”: $\bar{a} = (a, \alpha) \sim (c, \gamma) = \bar{c}$ iff $c \in a$ and $\alpha|_c = \gamma$. For $r = 1, 2$, let B_r denote the set of affine r -dimensional planes in \mathbb{F}^k , let Γ_{B_r} be the set of r -variate polynomials of degree at most d over \mathbb{F} , and let $\bar{B}_r = B_r \times \Gamma_{B_r}$. At various points during the proof, we will use that $\bar{A}/\bar{B}_r/\bar{C}$ is a double sampler. The incidence \times agreement edge relation extends naturally to \bar{A}/\bar{B}_r , \bar{B}_r/\bar{C} , and \bar{B}_2/\bar{B}_1 . For example, if $\bar{a} = (a, \alpha) \in \bar{A}$ and $\bar{b} = (b, \beta) \in \bar{B}_2$, then $\bar{a} \sim \bar{b}$ iff $b \subset a$ and $\alpha|_b = \beta$.

3.1 Incidence \times Agreement Samplers

We begin by listing the incidence \times agreement samplers we will need in the remainder of the paper and proving they are sampling. In the claim statement below, $\bar{A}(\bar{c})$, for $\bar{c} \in \bar{C}$, denotes the set of $\bar{a} \in \bar{A}$ such that $\bar{a} \sim \bar{c}$. In the proof which follows, we use $\bar{A}(\bar{c})$ to mean either this set, or the uniform distribution on this set; in all cases, our intention should be clear from the context.

Claim 2. *The following graphs are all $\mathcal{O}(|\mathbb{F}|^{-1})$ -biregular and $(12 \cdot |\mathbb{F}|^{-1/15}, |\mathbb{F}|^{-1/15})$ -sampling:*

- | | | |
|--|-------------------------------|--|
| (1) \bar{B}_1/\bar{C} | (2) \bar{A}/\bar{C} | (3) $\bar{A}(\bar{c}, \bar{c}')/\bar{C} \forall \bar{c}, \bar{c}' \in \bar{C}$ |
| (4) $\bar{A}(\bar{c})/\bar{C}^2 \forall \bar{c} \in \bar{C}$ | (5) \bar{A}/\bar{C}^2 | (6) $\bar{B}_2(\bar{c})/\bar{C} \forall \bar{c} \in \bar{C}$ |
| (7) $\bar{A}(\bar{b})/\bar{C} \forall \bar{b} \in \bar{B}_1$ | (8) $A \times \Gamma/\bar{C}$ | (9) $A \times \Gamma/\bar{B}_1$ |

Proof. It is easy to see that all of the graphs in the Claim statement are $\mathcal{O}(|\mathbb{F}|^{-1})$ -biregular, as per Definition 5. By symmetry, graphs (1), (2), (5), (8), (9) are actually 0-biregular. The others have a slight error introduced by the fact, for example, that the distribution which draws $\bar{a} \sim \bar{A}(\bar{c})$ and outputs a random element of $\bar{C}(\bar{a})$ is more likely to output \bar{c} than $\bar{c}' \neq \bar{c}$. However, an easy

calculation shows that the statistical distance between the required distributions is $\mathcal{O}(|\mathbb{F}|^{-1})$; the same is true for all examples in the list. The rest of the proof is divided into two stages. First, we use a pairwise independence argument to show that $\overline{B}_1/\overline{C}$, $\overline{B}_2/\overline{C}$, and $\overline{B}_2(\overline{c})/\overline{B}_1(\overline{c})$ for all $\overline{c} \in \overline{C}$ are $(|\mathbb{F}|^{-1/5}, |\mathbb{F}|^{-3/5})$ -sampling. Then we reduce the sampling of every graph above to the sampling of these three.

We phrase the pairwise independence argument for a generic bipartite graph A/B . The key feature we need involves a set X which parametrizes the neighborhoods $B(a)$ for all $a \in A$. Given $x \in X$ and $a \in A$, we write the x -th neighbor of a as $a(x) \in B$, so X parametrizes neighborhoods as $B(a) = \{a(x) : x \in X\}$ for all $a \in A$. The property we require is that for all $x_1 \neq x_2 \in X$, the random variable $(a(x_1), a(x_2))$ (randomness over $a \sim A$) is uniform on B^2 . For $\overline{B}_1/\overline{C}$, $X = \mathbb{F}$ since $\overline{C}(\overline{b})$ is parametrized by the points on the line b . Likewise, for $\overline{B}_2/\overline{C}$, $X = \mathbb{F}^2$. Finally, for $\overline{B}_2(\overline{c})/\overline{B}_1(\overline{c})$, $X = \mathbb{F} \cup \{\infty\}$, since $\overline{B}_1(\overline{c}, \overline{b}_2)$ is parametrized by all possible slopes of a line in b_2 through c . In all cases, independence follows from the fact that for every $b_1 \in B$, the distribution which draws $a \sim A(b_1)$ and outputs $b_2 \sim B(a) \setminus \{b_1\}$ is the uniform distribution on B .

So now, let A/B be a bipartite graph which satisfies the pairwise independent parametrized neighborhood property described above. Let $B' \subset B$ be a subset of size $|B'| = \lambda \cdot |B|$. For $b \in B$, let $\mathbb{1}_{B'}(b)$ indicate whether $b \in B'$ or not, and let $\hat{\mathbb{1}}_{B'}(b) := \mathbb{1}_{B'}(b) - \lambda$. Note $\mathbb{E}_{b \sim B}[\hat{\mathbb{1}}_{B'}(b)] = 0$. Finally, define $f : A \rightarrow [0, 1]$ by $f(a) := \mathbb{E}_{b \sim B(a)}[\hat{\mathbb{1}}_{B'}(b)]$. We will show $\mathbb{E}_{a \sim A}[f(a)^2] \leq |\mathbb{F}|^{-1}$. This suffices by Markov's inequality:

$$\Pr_{a \sim A} \left[\left| \Pr_{b \sim B(a)}(b \in B') - \lambda \right| > |\mathbb{F}|^{-1/5} \right] \leq \Pr_{a \sim A} \left[f(a)^2 > |\mathbb{F}|^{-2/5} \right] \leq |\mathbb{F}|^{2/5} \cdot \mathbb{E}_{a \sim A} [f(a)^2].$$

We use the pairwise independence property to conclude:

$$\begin{aligned} \mathbb{E}_{a \sim A} [f(a)^2] &= \mathbb{E}_{a \sim A} \left[\mathbb{E}_{x_1, x_2 \sim X} [\hat{\mathbb{1}}_{B'}(a(x_1)) \cdot \hat{\mathbb{1}}_{B'}(a(x_2))] \right] \\ &\leq \frac{1}{|X|} + \mathbb{E}_{b_1, b_2 \sim B} [\hat{\mathbb{1}}_{B'}(b_1) \cdot \hat{\mathbb{1}}_{B'}(b_2)] = \frac{1}{|X|}. \end{aligned}$$

For the reductions in the second phase, we use the following generic facts about samplers.

Fact 1 (Extending Sampling via Biregularity). Fix $\varepsilon, \varepsilon', \delta, \delta', \eta > 0$. Suppose $A/B/C$ are such that $B(a)/C(a)$ is η -biregular and $C(a, b) = C(b)$ for all $a \in A$ and $b \in B(a)$. The following hold.

1. If B/C is (ε', δ') -sampling and A/B is η -biregular, then A/C is (ε, δ) -sampling, where $\delta \geq \varepsilon^{-1} \cdot (2\eta + \varepsilon' + \delta')$.
2. If A/B is (ε', δ') -sampling and B/C is η -biregular, then A/C is (ε, δ) -sampling, where $\varepsilon \geq 3\varepsilon' + 2\eta$ and $\delta \geq \delta'/\varepsilon'$.

Fact 2 (Replacement Product). Let $\varepsilon, \varepsilon', \delta, \delta' > 0$ be such that $\delta \cdot (\varepsilon - 5\varepsilon') \geq 2\delta'/\varepsilon'$. Suppose $A/B/C$ is such that:

- A/C , B/C and $B(a)/C(a)$ are 0-biregular for all $a \in A$; and
- A/C and $A(c)/B(c)$ are (ε', δ') -sampling for all $c \in C$.

Then A/B is (ε, δ) -sampling.

We prove Fact 1 below, outside the current proof. Fact 2 is a version of the replacement product, proved originally in [WZ93] in the context of seeded randomness extractors (which are equivalent to sampler graphs). We include a proof in Appendix B for completeness. For now, we use these facts to complete the proof of Claim 2.

We have shown that $\overline{B}_2/\overline{C}$ and $\overline{B}_2(\overline{c})/\overline{B}_1(\overline{c})$ for all $\overline{c} \in \overline{C}$ are each $(|\mathbb{F}|^{-1/5}, |\mathbb{F}|^{-3/5})$ -sampling. Therefore, $\overline{B}_1/\overline{C}$ and $\overline{B}_2/\overline{B}_1$ are both $(7 \cdot |\mathbb{F}|^{-1/5}, |\mathbb{F}|^{-1/5})$ -sampling (we have already shown sampling of $\overline{B}_1/\overline{C}$ with better parameters, sampling of $\overline{B}_2/\overline{B}_1$ follows from Fact 2). The first point of Fact 1 says that any time we have Z such that Z/\overline{B}_1 or Z/\overline{B}_2 is $\mathcal{O}(|\mathbb{F}|^{-1})$ -biregular, then Z/\overline{C} or Z/\overline{B}_1 is $(3 \cdot |\mathbb{F}|^{-1/15}, 3 \cdot |\mathbb{F}|^{-2/15})$ -sampling. This proves the sampling of all graphs except for (4) and (5): $\overline{A}(\overline{c})/\overline{C}^2$ for all $\overline{c} \in \overline{C}$, and $\overline{A}/\overline{C}^2$, so it remains to prove sampling of these. Note $\overline{A}(\overline{c})/\overline{B}_1$ for all $\overline{c} \in \overline{C}$ and $\overline{A}/\overline{B}_1$ are $(3 \cdot |\mathbb{F}|^{-1/15}, 3 \cdot |\mathbb{F}|^{-2/15})$ -samplers, since $\overline{A}(\overline{c})/\overline{B}_2$ and $\overline{A}/\overline{B}_2$ are $\mathcal{O}(|\mathbb{F}|^{-1})$ -biregular. Thus we can use the second point Fact 1 to get $(12 \cdot |\mathbb{F}|^{-1/15}, |\mathbb{F}|^{-1/15})$ -sampling of graphs (4) and (5) because $\overline{B}_1/\overline{C}^2$ is $\mathcal{O}(|\mathbb{F}|^{-1})$ -biregular. \square

Proof of Fact 1. Assume $A/B/C$ are such that for all $a \in A$, $B(a)/C(a)$ is η -biregular, and $C(a, b) = C(b)$. Let $C' \subset C$ be a subset of size $|C'| = \lambda \cdot |C|$. The key observation in both cases is that for all $a \in A$,

$$\left| \Pr_{c \sim C(a)}(c \in C') - \lambda \right| \leq \left| \mathbb{E}_{b \sim B(a)}[\Pr_{c \sim C(b)}(c \in C')] - \lambda \right| + \eta.$$

Now, let $\text{val} := \Pr_{a \sim A} \left[\left| \Pr_{c \sim C(a)}(c \in C') - \lambda \right| > \varepsilon \right]$ be the quantity we have to bound. For the first point we have

$$\text{val} \leq \varepsilon^{-1} \cdot \left(\mathbb{E}_{\substack{a \sim A \\ b \sim B(a)}} \left[\left| \Pr_{c \sim C(b)}(c \in C') - \lambda \right| \right] + \eta \right) \leq \varepsilon^{-1} \cdot (2\eta + \varepsilon' + \delta'),$$

by Markov's inequality, the η -biregularity of A/B and the (ε', δ') -sampling of B/C . For the second point we have

$$\text{val} \leq \Pr_{a \sim A} \left[\left| \mathbb{E}_{b \sim B(a)}[\lambda(b)] - \mathbb{E}_{b \sim B}[\lambda(b)] \right| > \varepsilon - 2\eta \geq 3\varepsilon' \right] \leq \delta'/\varepsilon',$$

where $\lambda(b) := \Pr_{c \sim C(b)}(c \in C')$. We have used the η -biregularity of B/C to say that $\mathbb{E}_{b \sim B}[\lambda(b)]$ is in $\lambda \pm \eta$, and the (ε', δ') -sampling of A/B combined with the first point of Fact 3, stated in the next section. \square

3.2 Using Sampler Graphs

Fact 3 (Properties of Samplers). *Suppose A/B is η -biregular and (ε, δ) -sampling. We have the following.*

1. For any $\rho > 0$ and $f : B \rightarrow [0, 1]$,

$$\Pr_{a \sim A} \left[\left| \mathbb{E}_{b \sim B(a)}[f(b)] - \mathbb{E}_{b \sim B}[f(b)] \right| > \varepsilon + 2\rho \right] \leq \delta/\rho.$$

2. For any $\rho > 0$, B/A is $(\rho, 2(\varepsilon + \delta + \eta)/\rho)$ -sampling.
3. For any $B' \subset B$ of size $|B'| = \lambda \cdot |B|$ with $\lambda > \varepsilon$,

$$\Delta\left(\{(a, b) : \begin{smallmatrix} b \sim B' \\ a \sim A(b) \end{smallmatrix}\}, \{(a, b) : \begin{smallmatrix} a \sim A \\ b \sim B'(a) \end{smallmatrix}\}\right) \leq \delta + \eta/\varepsilon,$$

where $B'(a)$ denotes the distribution which draws $b \sim B(a)$ and outputs if $b \in B'$, else resamples (or if $B(a) \cap B' = \emptyset$, $B'(a)$ outputs an arbitrary $b \in B$).

The facts above are all well-known. See, for example, [Zuc97, IKW12, BDN17] for proofs of points 1, 2, and 3, respectively.

Notational Conventions and Example Use. Our proofs in the next sections relies heavily, and often implicitly, on the fact that the graphs of Claim 2 are samplers, and on the properties of sampler graphs stated in Fact 3. To facilitate readability, from here on, we reserve the quantity $\delta > 0$ for the loss introduced any time a sampling argument is used. As an example of how looks in the body of the paper, let $\bar{C}' \subset \bar{C}$ be a set with $|\bar{C}'| \geq \lambda \cdot |\bar{C}|$, and let \mathbf{E} be some event. Then we might deduce: $\mathbb{E}_{\bar{c}, \bar{c}' \sim \bar{C}'}[\Pr_{\bar{a} \sim \bar{A}(\bar{c}, \bar{c}')}(\mathbf{E})] \geq \mathbb{E}_{\bar{a} \sim \bar{A}}[\Pr_{\bar{c}, \bar{c}' \sim \bar{C}'(\bar{a})}(\mathbf{E})] - \delta$, “because of the sampling of \bar{A}/\bar{C}^2 .” Formally, we are using the third point of Fact 3, the fact that \bar{A}/\bar{C}^2 is η' -biregular, (ε', δ') -sampling with $\lambda > \varepsilon'$ and that $\delta \geq \delta' + \eta'/\varepsilon'$.

Setting the Sampling Parameter. In the example use mentioned above, $\eta' = \mathcal{O}(|\mathbb{F}|^{-1})$ and $\varepsilon', \delta' = \mathcal{O}(|\mathbb{F}|^{-1/15})$. Thus, $\delta = \mathcal{O}(|\mathbb{F}|^{-1/15})$ is sufficient for $\delta \geq \delta' + \eta'/\varepsilon'$ to hold. In general, each sampler property use will put a lower bound on δ , and so we simply set δ large enough so that they all hold. Explicitly, $\delta = 3 \cdot |\mathbb{F}|^{-1/60}$ is sufficient for our purposes.

We conclude this section with a corollary listing two sampler-based facts which will be useful in the calculations in the next section.

Claim 3. *Let the notations be as above, and let $\delta = 3 \cdot |\mathbb{F}|^{-1/60}$ and $\eta = \mathcal{O}(|\mathbb{F}|^{-1})$. Let $\bar{C}' \subset \bar{C}$ be a subset of size $|\bar{C}'|/|\bar{C}| \geq 12 \cdot |\mathbb{F}|^{-1/15}$. We have the following.*

1.

$$\left\{ \left(\bar{c}, \bar{b}, \bar{c}' \right) \middle| \begin{array}{l} \bar{a} \sim \bar{A} \\ \bar{c} \sim \bar{C}(\bar{a}) \\ \bar{c}' \sim \bar{C}'(\bar{a}) \\ \mathbf{b} \sim \mathbf{B}_2(\mathbf{a}, \mathbf{c}, \mathbf{c}') \end{array} \right\} \approx_{\delta} \left\{ \left(\bar{c}, \bar{b}, \bar{c}' \right) \middle| \begin{array}{l} \bar{c} \sim \bar{C} \\ \bar{b} \sim \bar{B}_2(\bar{c}) \\ \bar{c}' \sim \bar{C}'(\bar{b}) \end{array} \right\},$$

where in the first distribution $\bar{\mathbf{b}} = (\mathbf{b}, \alpha|_{\mathbf{b}})$, where $\bar{\mathbf{a}} = (\mathbf{a}, \alpha)$.

2.

$$\left\{ \left(\bar{\mathbf{a}}, \bar{\mathbf{b}}, \bar{\mathbf{c}}' \right) \middle| \begin{array}{l} \bar{\mathbf{a}} \sim \bar{\mathbf{A}} \\ \bar{\mathbf{c}} \sim \bar{\mathbf{C}}(\bar{\mathbf{a}}) \\ \bar{\mathbf{c}}' \sim \bar{\mathbf{C}}'(\bar{\mathbf{a}}) \\ \mathbf{b} \sim \mathbf{B}_2(\mathbf{c}, \mathbf{c}') \end{array} \right\} \approx_{\delta} \left\{ \left(\bar{\mathbf{a}}, \bar{\mathbf{b}}, \bar{\mathbf{c}}' \right) \middle| \begin{array}{l} \bar{\mathbf{c}}' \sim \bar{\mathbf{C}}' \\ \bar{\mathbf{b}} \sim \bar{\mathbf{B}}_2(\bar{\mathbf{c}}') \\ \bar{\mathbf{a}} \sim \bar{\mathbf{A}}(\bar{\mathbf{b}}) \end{array} \right\},$$

where in the first distribution $\bar{\mathbf{b}} = \bar{\mathbf{a}}|_{\mathbf{b}}$.

In both (1) and (2) above, \approx_δ denotes that the two distributions are within statistical distance δ of one another.

Proof. For the first part, we have

$$\left\{ \begin{array}{l} \bar{a} \sim \bar{A} \\ \bar{c} \sim \bar{C}(\bar{a}) \\ \bar{c}' \sim \bar{C}'(\bar{a}) \end{array} \right\} \approx_{\delta/3} \left\{ \begin{array}{l} \bar{c}' \sim \bar{C}' \\ \bar{a} \sim \bar{A}(\bar{c}') \\ \bar{c} \sim \bar{C}(\bar{a}) \end{array} \right\} \approx_\eta \left\{ \begin{array}{l} \bar{c}' \sim \bar{C}' \\ \bar{b} \sim \bar{B}_2(\bar{c}') \\ \bar{c} \sim \bar{C}(\bar{b}) \end{array} \right\} \approx_{\delta/3} \left\{ \begin{array}{l} \bar{b} \sim \bar{B}_2 \\ \bar{c}' \sim \bar{C}'(\bar{b}) \\ \bar{c} \sim \bar{C}(\bar{b}) \end{array} \right\},$$

where each distribution outputs $(\bar{c}, \bar{b}, \bar{c}')$ and where $\bar{b} = \bar{a}|_{\bar{b}}$ for $\bar{b} \sim B_2(a, c, c')$ is implied in the first two distributions. The first relation follows from sampling of \bar{A}/\bar{C} ; the second follows from the η -biregularity of $B_2(\bar{a}, \bar{c}')/\bar{C}(\bar{a})$ for all $\bar{a} \in \bar{A}$ and $\bar{c}' \in \bar{C}(\bar{a})$, and the 0-biregularity of $\bar{A}(\bar{c}')/\bar{B}(\bar{c}')$ for all $\bar{c}' \in \bar{C}$; the third follows from the sampling of \bar{B}_2/\bar{C} . Finally, the last distribution is identical to the desired distribution on the right of point 1 because of the 0-biregularity of \bar{B}_2/\bar{C} . We work similarly for the second point:

$$\left\{ \begin{array}{l} \bar{a} \sim \bar{A} \\ \bar{c} \sim \bar{C}(\bar{a}) \\ \bar{c}' \sim \bar{C}'(\bar{a}) \end{array} \right\} \approx_{\delta/2} \left\{ \begin{array}{l} \bar{c}' \sim \bar{C}' \\ \bar{a} \sim \bar{A}(\bar{c}') \\ \bar{c} \sim \bar{C}(\bar{a}) \end{array} \right\} \approx_\eta \left\{ \begin{array}{l} \bar{c}' \sim \bar{C}' \\ \bar{a} \sim \bar{A}(\bar{c}') \\ \bar{b} \sim \bar{B}_2(\bar{a}, \bar{c}') \end{array} \right\} \equiv \left\{ \begin{array}{l} \bar{c}' \sim \bar{C}' \\ \bar{b} \sim \bar{B}_2(\bar{c}') \\ \bar{a} \sim \bar{A}(\bar{b}) \end{array} \right\},$$

where each distribution outputs $(\bar{a}, \bar{b}, \bar{c}')$ and where $\bar{b} = \bar{a}|_{\bar{b}}$ (as above, $\bar{b} \sim B_2(a, c, c')$ is implicit in the first two distributions). We have used the sampling of \bar{A}/\bar{C} , η -biregularity of $\bar{B}_2(\bar{a}, \bar{c}')/\bar{C}(\bar{a})$ for all $\bar{a} \in \bar{A}$ and $\bar{c}' \in \bar{C}(\bar{a})$, and 0-biregularity of $\bar{A}(\bar{c}')/\bar{B}_2(\bar{c}')$ for all $\bar{c}' \in \bar{C}$. \square

4 Global Agreement

In this section we prove Lemma 1, restated below in a quantitative form.

Lemma 1 (Restated). *Suppose $\varepsilon \geq |\mathbb{F}|^{-1/1000}$, and fix parameters $\eta = |\mathbb{F}|^{-9/10}$, $\delta = 3 \cdot |\mathbb{F}|^{-1/60}$, and $\tau = \mathcal{O}(\delta/\varepsilon^6 + \eta/\varepsilon^{11})$. Suppose $\{f_a\}_a \subset \{f : \Gamma_A \rightarrow \Gamma_A\}$ is such that*

$$\Pr_{\Phi, (c, a, a')} [\tilde{\alpha}|_c = \tilde{\alpha}'|_c] = 6\varepsilon \tag{5}$$

where the probability is over $\Phi \sim \Gamma$, $c \sim C$, $a, a' \sim A(c)$, and where $(\tilde{\alpha}, \tilde{\alpha}') = (f_a(\Phi|_a), f_{a'}(\Phi|_{a'}))$. Then there exists a set $G \subset A \times \Gamma$ of size at least $|G| \geq 2\varepsilon \cdot |A \times \Gamma|$ and a function $h : \bar{C} \rightarrow \Gamma_C$ such that: $\Pr_{\substack{(a, \Phi) \sim G \\ c \sim C(a)}} [\tilde{\gamma} \sim \tilde{\alpha}] \geq 1 - \zeta$, where $\tilde{\gamma} = h(c, \Phi|_c)$ and $\zeta := \varepsilon^{-2} \cdot (\tau + \delta) + \varepsilon^{-1} \cdot (\eta + \delta)$.

Remark. Many different $\mathcal{O}(|\mathbb{F}|^{-1})$ parameters are introduced during the course of our analysis which are all $\mathcal{O}(|\mathbb{F}|^{-1})$. We encourage the reader to think of two levels of parameters: level one consists of ε only; all other parameters are in level 2 and are much smaller. The level two parameters are each defined to be smaller than ε^c for some constant $c = \mathcal{O}(1)$ which arises during our analysis. So in the above theorem, for example, in order for τ to be level 2, it must be that

$\delta \ll \varepsilon^6$ and $\eta \ll \varepsilon^{11}$; additionally, for ζ to be level 2, $\tau \ll \varepsilon^2$ is required. We remark that the analysis prioritizes modularity and succinctness, rather than optimizing constants. As a result, the small constant $1/1000$ is suboptimal.

We begin by introducing the notation and ideas needed to prove Lemma 1 in Section 4.1. The actual proof appears in Section 4.2, conditioned on two claims which we state in Section 4.1 and prove in Section 4.3.

4.1 Proof Setup.

Notations. In this section \mathbb{B} denotes the set of 2-dimensional planes in \mathbb{F}^m , and $\Gamma_{\mathbb{B}}$ is the set of 2-variate polynomials over \mathbb{F} of degree at most d , and $\overline{\mathbb{B}} = \mathbb{B} \times \Gamma_{\mathbb{B}}$. The sets $\overline{\mathbb{A}}, \overline{\mathbb{C}}, \Gamma$ are as usual. We will take advantage of the sampling properties of the triple $\overline{\mathbb{A}}/\overline{\mathbb{B}}/\overline{\mathbb{C}}$. When considering two polynomials whose domains intersect, we write \sim to indicate that they agree on the intersection. For example, given $\tilde{\alpha}, \tilde{\alpha}' \in \Gamma_{\mathbb{A}}$ defined on $\mathbf{a}, \mathbf{a}' \in \mathbb{A}(\overline{\mathbb{C}})$ we write $\tilde{\alpha} \sim \tilde{\alpha}'$ if $\tilde{\alpha}|_{\mathbf{c}} = \tilde{\alpha}'|_{\mathbf{c}}$.

We say that $(\mathbf{c}, \gamma, \tilde{\gamma})$ is *good* if $\Pr_{(\mathbf{a}, \Phi)}[\tilde{\alpha} \sim \tilde{\gamma}] \geq 4\varepsilon$, where the probability is over $\mathbf{a} \sim \mathbb{A}(\mathbf{c})$ and $\Phi \sim \Gamma(\overline{\mathbb{C}})$. We say $\overline{\mathbf{c}} = (\mathbf{c}, \gamma)$ is *good* if there exists $\tilde{\gamma}$ such that $(\mathbf{c}, \gamma, \tilde{\gamma})$ is. Note that $\Pr_{\overline{\mathbf{c}} \sim \overline{\mathbb{C}}}[\overline{\mathbf{c}} \text{ good}] \geq 2\varepsilon$. To see this, let $\mathbf{p}_{\mathbf{c}, \gamma, \tilde{\gamma}} = \Pr_{(\mathbf{a}, \Phi)}[\tilde{\alpha} \sim \tilde{\gamma}]$. Then (5) gives

$$6\varepsilon = \mathbb{E}_{\overline{\mathbf{c}} \sim \overline{\mathbb{C}}} \left[\sum_{\tilde{\gamma}} \mathbf{p}_{\mathbf{c}, \gamma, \tilde{\gamma}} \cdot \Pr_{\mathbf{a}' \sim \mathbb{A}(\mathbf{c})}[\tilde{\alpha}' \sim \tilde{\gamma}] \right] \leq \mathbb{E}_{\overline{\mathbf{c}} \sim \overline{\mathbb{C}}} \left[\max_{\tilde{\gamma}} \{ \mathbf{p}_{\mathbf{c}, \gamma, \tilde{\gamma}} \} \right].$$

We have used that $\sum_{\tilde{\gamma}} \Pr_{\mathbf{a}' \sim \mathbb{A}(\mathbf{c})}[\tilde{\alpha}' \sim \tilde{\gamma}] = 1$ for all $\overline{\mathbf{c}}$.

Local Functions. Let $h_0 : \overline{\mathbb{C}} \rightarrow \Gamma_{\mathbb{C}}$ be the randomized function which sends $\overline{\mathbf{c}} = (\mathbf{c}, \gamma)$ to a random $\tilde{\gamma}$ such that $(\mathbf{c}, \gamma, \tilde{\gamma})$ is good if such $\tilde{\gamma}$ exists, and to an arbitrary $\tilde{\gamma} \in \Gamma_{\mathbb{C}}$ if not. For $\overline{\mathbf{c}} \in \overline{\mathbb{C}}$, let $g_{\overline{\mathbf{c}}} : \overline{\mathbb{B}}(\overline{\mathbf{c}}) \rightarrow \Gamma_{\mathbb{B}}$ be the randomized function where $g_{\overline{\mathbf{c}}}(\overline{\mathbf{b}})$ is the distribution on $\Gamma_{\mathbb{B}}$ which draws $\overline{\mathbf{a}} \sim \overline{\mathbb{A}}(\overline{\mathbf{b}})$ such that $\tilde{\alpha} \sim h_0(\overline{\mathbf{c}})$, and outputs $\tilde{\beta} = \tilde{\alpha}|_{\mathbf{b}}$.

Definition 7 (Well-Defined). Let $\eta = |\mathbb{F}|^{-9/10}$. We say that $g_{\overline{\mathbf{c}}}$ is well-defined if

$$\Pr_{\substack{\overline{\mathbf{b}} \sim \overline{\mathbb{B}}(\overline{\mathbf{c}}) \\ \overline{\mathbf{a}}, \overline{\mathbf{a}}' \sim \overline{\mathbb{A}}(\overline{\mathbf{b}})}} \left[\tilde{\alpha} \approx \tilde{\alpha}' \mid \tilde{\alpha} \sim h_0(\overline{\mathbf{c}}) \sim \tilde{\alpha}' \right] \geq 1 - \eta,$$

where $\tilde{\alpha} \approx \tilde{\alpha}'$ indicates that $\tilde{\alpha}|_{\mathbf{b}} = \tilde{\alpha}'|_{\mathbf{b}}$.

Previous work [IKW12, BDN17] refers to the good $\overline{\mathbf{c}} \in \overline{\mathbb{C}}$ for which $g_{\overline{\mathbf{c}}}$ is well-defined as *excellent*; the fact that the excellent points comprise a non-negligible fraction of $\overline{\mathbb{C}}$ is a crucial component of the proofs in these papers. We require one extra property from our specialized subset of $\overline{\mathbb{C}}$ which simplifies the remainder of our proof greatly. The following is proved in Section 4.3.

Claim 4. There exists a set $\overline{\mathbb{C}}' \subset \overline{\mathbb{C}}$ such that the following hold: 1) $|\overline{\mathbb{C}}'| \geq \varepsilon^3 |\overline{\mathbb{C}}|$; 2) every $\overline{\mathbf{c}} \in \overline{\mathbb{C}}'$ is good and such that $g_{\overline{\mathbf{c}}}$ is well-defined; 3)

$$\Pr_{\overline{\mathbf{c}}, \overline{\mathbf{c}}' \sim \overline{\mathbb{C}}'} \left[\Pr_{\overline{\mathbf{a}} \sim \overline{\mathbb{A}}(\overline{\mathbf{c}}, \overline{\mathbf{c}}')} [h_0(\overline{\mathbf{c}}) \sim \tilde{\alpha} \sim h_0(\overline{\mathbf{c}}')] \geq \varepsilon^5 \right] \geq 1 - \sigma,$$

where $\sigma := \delta/\varepsilon^3 + \delta/\varepsilon^6 + \eta/\varepsilon^{11}$.

Intuitively, the extra property captured by (3) demands that the set of excellent points can be partitioned into large sets of *mutually compatible* points; the set \bar{C}' is any member of this partition.

The Global Function. Let $h : \bar{C} \rightarrow \Gamma_C$ be the randomized function where $h(\bar{c})$ draws $\bar{b} \sim \bar{B}(\bar{c})$, $\bar{c}' \sim \bar{C}'(\bar{b})$ and outputs $\tilde{\beta}|_c$ where $\tilde{\beta} = g_{\bar{c}'}(\bar{b})$. The following is also proved in Section 4.3.

Claim 5. We have $\Pr_{(\bar{c}, \bar{b}, \bar{c}')} [h(\bar{c}) \sim \tilde{\beta}] \geq 1 - \tau$, where $\tau := (\sigma + 2\varepsilon^{-5}(\eta + \delta) + 2\delta)$, $\tilde{\beta} = g_{\bar{c}'}(\bar{b})$ and the probability is over $\bar{c} \sim \bar{C}$, $\bar{b} \sim \bar{B}(\bar{c})$, $\bar{c}' \sim \bar{C}'(\bar{b})$.

4.2 Proof of Lemma 1

Notational Convention. Let $h_0, h : \bar{C} \rightarrow \Gamma_C$ be the functions defined in Section 4.1. In this section if we write $\tilde{\gamma}$ when working with $\bar{c} \in \bar{C}$, it should be understood that $\tilde{\gamma} = h(\bar{c})$. We will always refer to $h_0(\bar{c})$ explicitly.

Proof. Suppose $(\varepsilon, \{f_a\})$ are such that (5) holds; let $\bar{C}' \subset \bar{C}$ be the set guaranteed by Claim 4. We define G to be the set of $(a, \Phi) \in A \times \Gamma$ such that $\Pr_{\bar{c} \sim \bar{C}'(\bar{a})} [\tilde{\alpha} \sim h_0(\bar{c})] \geq \varepsilon$. We have,

$$\mathbb{E}_{(a, \Phi) \sim A \times \Gamma} \left[\Pr_{\bar{c} \sim \bar{C}'(\bar{a})} [\tilde{\alpha} \sim h_0(\bar{c})] \right] \geq \mathbb{E}_{\bar{c} \sim \bar{C}'} \left[\Pr_{\substack{a \sim A(c) \\ \Phi \sim \Gamma(\bar{c})}} [\tilde{\alpha} \sim h_0(\bar{c})] \right] - \delta \geq 3\varepsilon$$

We have used the sampling of $A \times \Gamma / \bar{C}$ for the first inequality, and that all $\bar{c} \in \bar{C}'$ are good for the second (and $4\varepsilon - \delta \geq 3\varepsilon$). It follows that $|G| \geq 2\varepsilon|A \times \Gamma|$. Thus, it remains to prove that $\Pr_{(a, \Phi), c} [\tilde{\gamma} \sim \tilde{\alpha}] \geq 1 - \zeta$, where the probability is over $(a, \Phi) \sim G$, $c \sim C(a)$ and where $\tilde{\gamma} = h(c, \Phi|_c)$, where h is the global function defined in Section 4.1.

So let $p := \Pr_{(a, \Phi), c} [\tilde{\gamma} \sim \tilde{\alpha}]$ be the probability we are trying to bound. We have

$$p \geq \Pr_{\substack{(a, \Phi) \\ b, c, \bar{c}'}} [\tilde{\gamma} \sim \tilde{\beta} \sim \tilde{\alpha} | \tilde{\alpha} \sim h_0(\bar{c}')] \geq \Pr_{\substack{(a, \Phi) \\ b, c, \bar{c}'}} [\tilde{\gamma} \sim \tilde{\beta} | \tilde{\alpha} \sim h_0(\bar{c}')] - \Pr_{\substack{(a, \Phi) \\ b, c, \bar{c}'}} [\tilde{\beta} \not\sim \tilde{\alpha} | \tilde{\alpha} \sim h_0(\bar{c}')],$$

where the probabilities are over $(a, \Phi) \sim G$, $c \sim C(a)$, $\bar{c}' \sim \bar{C}'(\bar{a})$, $b \sim B(a, c, c')$, and where $\tilde{\beta} = g_{\bar{c}'}(\bar{b})$, for $\bar{b} = (b, \Phi|_b)$. We conclude by bounding both probabilities on the right; denoted RHS_1 and RHS_2 , respectively. We have

$$\begin{aligned} 1 - \text{RHS}_1 &= \Pr_{\substack{(a, \Phi) \sim G \\ b, c, \bar{c}'}} [\tilde{\gamma} \not\sim \tilde{\beta} | \tilde{\alpha} \sim h_0(\bar{c}')] \leq \frac{\Pr_{(a, \Phi), b, c, \bar{c}'} [\tilde{\gamma} \not\sim \tilde{\beta}]}{\min_{(a, \Phi) \in G} \left\{ \Pr_{\bar{c}' \sim \bar{C}'(\bar{a})} [\tilde{\alpha} \sim h_0(\bar{c}')] \right\}} \\ &\leq \frac{\varepsilon^{-2}}{2} \cdot \Pr_{\substack{\bar{a} \sim \bar{A} \\ b, c, \bar{c}'}} [\tilde{\gamma} \not\sim \tilde{\beta}] < \varepsilon^{-2} \cdot \left(\Pr_{\substack{\bar{c} \sim \bar{C} \\ \bar{b} \sim \bar{B}(\bar{c}) \\ \bar{c}' \sim \bar{C}'(\bar{b})}} [\tilde{\gamma} \not\sim \tilde{\beta}] + \delta \right) \leq \varepsilon^{-2} \cdot (\tau + \delta). \end{aligned}$$

The first inequality on the second line used the definition of G and that $|G| \geq 2\varepsilon \cdot |A \times \Gamma|$; the second used Claim 3, point 1; and the last used Claim 5. Finally,

$$\text{RHS}_2 \leq \frac{\varepsilon^{-1}}{2} \cdot \Pr_{\substack{\bar{a} \sim \bar{A} \\ \bar{c}' \sim \bar{C}'(\bar{a}) \\ \bar{b} \sim \bar{B}(\bar{c}', \bar{a})}} [\tilde{\beta} \not\sim \tilde{\alpha} | \tilde{\alpha} \sim h_0(\bar{c}')]$$

$$\leq \varepsilon^{-1} \cdot \left(\max_{\bar{c}' \in \bar{\mathcal{C}}'} \left\{ \Pr_{\substack{\bar{b} \sim \bar{B}(\bar{c}') \\ \bar{a} \sim \bar{A}(\bar{b})}} [\tilde{\beta} \not\sim \tilde{\alpha} | \tilde{\alpha} \sim \mathbf{h}_0(\bar{c}')] \right\} + \delta \right) \leq \varepsilon^{-1} (\eta + \delta).$$

We have used Claim 3 point 2 and the fact that $\mathbf{g}_{\bar{c}'}$ is well-defined for all $\bar{c}' \in \bar{\mathcal{C}}'$. The result follows. \square

4.3 Proving the Claims

Starting Assumption and Notational Conventions. Throughout this section, we assume the hypotheses of Lemma 1, namely $(\varepsilon, \{f_a\})$ are such that $\Pr_{\Phi, (c, a, a')} [\tilde{\alpha} \sim \tilde{\alpha}'] = 6\varepsilon$ (i.e., such that (5) holds). Let $\mathbf{h}_0, \mathbf{h} : \bar{\mathcal{C}} \rightarrow \Gamma_{\mathcal{C}}$ be the functions defined in Section 4.1. In this section if we write $\tilde{\gamma}$ when working with $\bar{c} \in \bar{\mathcal{C}}$, it should be understood that $\tilde{\gamma} = \mathbf{h}_0(\bar{c})$. We will refer to $\mathbf{h}(\bar{c})$ explicitly (note, this is opposite to the convention of Section 4.2). Given $\bar{c}, \bar{c}' \in \bar{\mathcal{C}}$ set $\mu_{\bar{c}}, \rho(\bar{c})$ and $\mathbf{q}(\bar{c}, \bar{c}')$ to:

$$\Pr_{\substack{\bar{a} \sim A(\bar{c}) \\ \Phi \sim \Gamma(\bar{c})}} [\tilde{\gamma} \sim \tilde{\alpha}]; \Pr_{\substack{\bar{b} \sim \bar{B}(\bar{c}) \\ \bar{a} \sim \bar{A}(\bar{b})}} [\tilde{\beta} \sim \tilde{\alpha} | \tilde{\gamma} \sim \tilde{\alpha}]; \Pr_{\bar{a} \sim \bar{A}(\bar{c}, \bar{c}')} [\tilde{\gamma} \sim \tilde{\alpha} \sim \tilde{\gamma}'].$$

In Section 4.1 we called $\bar{c} \in \bar{\mathcal{C}}$ such that $\mu_{\bar{c}} \geq 4\varepsilon$ *good*. Also for $\bar{c} \in \bar{\mathcal{C}}$ we defined local functions $\mathbf{g}_{\bar{c}} : \bar{B}(\bar{c}) \rightarrow \Gamma_B$ and said that $\mathbf{g}_{\bar{c}}$ was *well-defined* if $\rho(\bar{c}) \geq 1 - \eta$, where $\eta = |\mathbb{F}|^{-9/10}$. In the remainder of this section we prove three claims; the first two combine to prove Claim 4, the last is Claim 5.

Claim 6. *There exists a set $\bar{\mathcal{C}}'_0 \subset \bar{\mathcal{C}}$ such that the following hold: 1) $|\bar{\mathcal{C}}'_0| \geq \varepsilon |\bar{\mathcal{C}}|$; 2) $\mu_{\bar{c}} \geq 4\varepsilon$ for every $\bar{c} \in \bar{\mathcal{C}}'_0$; 3) $\rho(\bar{c}) \geq 1 - \eta$ for every $\bar{c} \in \bar{\mathcal{C}}'_0$.*

Proof. Let $\bar{\mathcal{C}}'_0 \subset \bar{\mathcal{C}}$ be the set of $\bar{c} \in \bar{\mathcal{C}}$ for which $\mu_{\bar{c}} \geq 4\varepsilon$ and $\rho(\bar{c}) \geq 1 - \eta$ (i.e., $\bar{c} \in \bar{\mathcal{C}}'_0$ if \bar{c} is good and such that $\mathbf{g}_{\bar{c}}$ is well-defined). We bound $|\bar{\mathcal{C}}'_0|$ using three observations. First, as noted in Section 4.1, $\Pr_{\bar{c} \sim \bar{\mathcal{C}}} [\mu_{\bar{c}} \geq 4\varepsilon] \geq 2\varepsilon$. Second, for all $\bar{c} \in \bar{\mathcal{C}}$ such that $\mu_{\bar{c}} \geq 4\varepsilon$:

$$\Pr_{\substack{\bar{b} \sim \bar{B}(\bar{c}) \\ \bar{a}, \bar{a}' \sim \bar{A}(\bar{b})}} [\tilde{\alpha} \sim \tilde{\gamma} \sim \tilde{\alpha}'] = \mathbb{E}_{\bar{b} \sim \bar{B}(\bar{c})} [\mu_{\bar{c}}(\bar{b})^2] \geq \Pr_{\bar{b} \sim \bar{B}(\bar{c})} [|\mu_{\bar{c}}(\bar{b}) - \mu_{\bar{c}}| \leq \varepsilon] \cdot 9\varepsilon^2 \geq \varepsilon^2,$$

where $\mu_{\bar{c}}(\bar{b}) := \Pr_{\bar{a}, \bar{a}' \sim \bar{A}(\bar{b})} [\tilde{\alpha} \sim \tilde{\gamma}]$ is shorthand. We have used the sampling of $\bar{A}(\bar{c})/\bar{B}(\bar{c})$ to (crudely) lower bound $\Pr_{\bar{b} \sim \bar{B}(\bar{c})} [|\mu_{\bar{c}}(\bar{b}) - \mu_{\bar{c}}| \leq \varepsilon]$. Finally, by Markov's inequality and Schwartz-Zippel:

$$\Pr_{\bar{c} \sim \bar{\mathcal{C}}} \left[\Pr_{\substack{\bar{b} \sim \bar{B}(\bar{c}) \\ \bar{a}, \bar{a}' \sim \bar{A}(\bar{b})}} [\tilde{\alpha} \not\sim \tilde{\alpha}' \ \& \ \tilde{\alpha} \sim \tilde{\gamma} \sim \tilde{\alpha}'] > \eta \varepsilon^2 \right] \leq \frac{d}{\eta \varepsilon^2 |\mathbb{F}|}.$$

Putting these together gives

$$\begin{aligned} \frac{|\bar{\mathcal{C}}'_0|}{|\bar{\mathcal{C}}|} &= \Pr_{\bar{c} \sim \bar{\mathcal{C}}} \left[\mu_{\bar{c}} \geq 4\varepsilon \ \& \ \Pr_{\substack{\bar{b} \sim \bar{B}(\bar{c}) \\ \bar{a}, \bar{a}' \sim \bar{A}(\bar{b})}} [\tilde{\alpha} \not\sim \tilde{\alpha}' | \tilde{\alpha} \sim \tilde{\gamma} \sim \tilde{\alpha}'] \leq \eta \right] \\ &\geq \Pr_{\bar{c} \sim \bar{\mathcal{C}}} [\mu_{\bar{c}} \geq 4\varepsilon] - \Pr_{\bar{c} \sim \bar{\mathcal{C}}} \left[\Pr_{(\bar{b}, \bar{a}, \bar{a}')} [\tilde{\alpha} \not\sim \tilde{\alpha}' \ \& \ \tilde{\alpha} \sim \tilde{\gamma} \sim \tilde{\alpha}'] > \eta \varepsilon^2 \right] \\ &\geq 2\varepsilon - \frac{d}{\eta \varepsilon^2 |\mathbb{F}|} \geq \varepsilon. \end{aligned}$$

\square

Claim 4 (Restated). There exists a set $\bar{C}' \subset \bar{C}$ such that the following hold: 1) $|\bar{C}'| \geq \varepsilon^3 |\bar{C}|$; 2) $\mu_{\bar{c}} \geq 4\varepsilon$ for every $\bar{c} \in \bar{C}'$; 3) $p(\bar{c}) \geq 1 - \eta$ for every $\bar{c} \in \bar{C}'$; 4) $\Pr_{\bar{c}, \bar{c}' \sim \bar{C}'} [q(\bar{c}, \bar{c}') \geq \varepsilon^5] \geq 1 - \sigma$, where $\sigma := \delta/\varepsilon^3 + \delta/\varepsilon^6 + \eta/\varepsilon^{11}$.

Proof. By Claim 6 it suffices to construct a large subset of \bar{C}'_0 such that the fourth property holds. For this purpose, we equip \bar{C}'_0 with a graph structure: $\bar{c}, \bar{c}' \in \bar{C}'_0$ are adjacent if $q(\bar{c}, \bar{c}') \geq \varepsilon^2$. Our final set \bar{C}' will be the neighborhood, $N(\bar{c}') := \{\bar{c} \in \bar{C}'_0 : q(\bar{c}, \bar{c}') \geq \varepsilon^2\}$ of some $\bar{c}' \in \bar{C}'_0$. In order for this to work, \bar{c}' should satisfy: 1) $|N(\bar{c}')|$ must be large; 2) $\Pr_{\bar{c}, \bar{c}'' \sim N(\bar{c}')} [q(\bar{c}, \bar{c}'') < \varepsilon^5]$ must be small. We show there exists such a $\bar{c}' \in \bar{C}'_0$. Specifically we prove

1. $\mathbb{E}_{\bar{c}, \bar{c}' \sim \bar{C}'_0} [q(\bar{c}, \bar{c}')] \geq 3\varepsilon^2$; and
2. $\Pr_{\substack{\bar{c}' \sim \bar{C}'_0 \\ \bar{c}, \bar{c}'' \sim N(\bar{c}')}} [q(\bar{c}, \bar{c}'') \geq \varepsilon^5 \mid |N(\bar{c}')| > \varepsilon^3 |\bar{C}|] \geq 1 - \sigma$.

It follows from the first point that $\Pr_{\bar{c}' \sim \bar{C}'_0} [|N(\bar{c}')| \geq \varepsilon^3 |\bar{C}|] > \varepsilon^2$ (using $|\bar{C}'_0| \geq \varepsilon |\bar{C}|$). Thus, the two points together guarantee the existence of some $\bar{c}' \in \bar{C}'_0$ such that $|N(\bar{c}')| \geq \varepsilon^3 |\bar{C}|$ and $\Pr_{\bar{c}, \bar{c}'' \sim N(\bar{c}')} [q(\bar{c}, \bar{c}'') \geq \varepsilon^5] \geq 1 - \sigma$. Setting $\bar{C}' = N(\bar{c}')$ for such a $\bar{c}' \in \bar{C}'_0$ completes the proof. So it remains to establish the above two bounds.

For the first, we have

$$\begin{aligned} \mathbb{E}_{\bar{c}, \bar{c}' \sim \bar{C}'_0} [q(\bar{c}, \bar{c}')] &\geq \mathbb{E}_{\bar{a} \sim \bar{A}} \left[\Pr_{\bar{c} \sim \bar{C}'_0(\bar{a})} [\tilde{\gamma} \sim \tilde{\alpha}]^2 \right] - \delta \geq \mathbb{E}_{\bar{a} \sim \bar{A}} \left[\Pr_{\bar{c} \sim \bar{C}'_0(\bar{a})} [\tilde{\gamma} \sim \tilde{\alpha}] \right]^2 - \delta \\ &\geq \mathbb{E}_{\bar{c} \sim \bar{C}'_0} [\mu_{\bar{c}}]^2 - 3\delta \geq 16\varepsilon^2 - 3\delta \geq 3\varepsilon^2. \end{aligned}$$

We have used the sampling of \bar{A}/\bar{C}^2 , Jensen's inequality, the sampling of \bar{A}/\bar{C} , and the fact that $\mu_{\bar{c}} \geq 4\varepsilon$ for all $\bar{c} \in \bar{C}'_0$. Establishing the second bound is more involved. Towards this end, we define three quantities, shorthanded as $\text{val}_1, \text{val}_2, \text{val}_3$; each is a function of $(\bar{c}, \bar{c}', \bar{c}'')$:

- $\text{val}_1 := \left| \Pr_{\bar{a}' \sim \bar{A}(\bar{c}, \bar{c}', \bar{c}'')} [\tilde{\gamma} \sim \tilde{\alpha}' \sim \tilde{\gamma}'] - q(\bar{c}, \bar{c}'') \right|$;
- $\text{val}_2 := \left| \Pr_{\bar{a}' \sim \bar{A}(\bar{c}, \bar{c}', \bar{c}'')} [\tilde{\gamma}' \sim \tilde{\alpha}'] - \mu_{\bar{c}'} \right|$;
- $\text{val}_3 := \Pr_{\substack{\bar{b} \sim \bar{B}(\bar{c}, \bar{c}') \\ \bar{a} \sim \bar{A}(\bar{b}) \\ \bar{a}' \sim \bar{A}(\bar{b}, \bar{c}')}} [\tilde{\alpha} \not\sim \tilde{\alpha}' \mid \tilde{\alpha} \sim \tilde{\gamma}' \sim \tilde{\alpha}'] + \Pr_{\substack{\bar{b}'' \sim \bar{B}(\bar{c}', \bar{c}'') \\ \bar{a}'' \sim \bar{A}(\bar{b}'') \\ \bar{a}' \sim \bar{A}(\bar{b}'', \bar{c})}} [\tilde{\alpha}' \not\sim \tilde{\alpha}'' \mid \tilde{\alpha}' \sim \tilde{\gamma}' \sim \tilde{\alpha}'']$.

We show that each val_i is small with very high probability over $(\bar{c}, \bar{c}', \bar{c}'')$ drawn as follows: $\bar{c}' \sim \bar{C}'_0$ such that $|N(\bar{c}')| \geq \varepsilon^3 |\bar{C}|$, $\bar{c}, \bar{c}'' \sim N(\bar{c}')$. These bounds will be used in the computation which follows. We have

$$\Pr_{(\bar{c}, \bar{c}', \bar{c}'')} [\text{val}_1 > \delta] \leq \varepsilon^{-3} \cdot \max_{\bar{c}, \bar{c}'' \in \bar{C}} \left\{ \Pr_{\bar{c}' \sim \bar{C}'} \left[\left| \mathbb{E}_{\bar{a}' \sim \bar{A}(\bar{c}, \bar{c}', \bar{c}'')} [f_1(\bar{a}')] - \mathbb{E}_{\bar{a}' \sim \bar{A}(\bar{c}, \bar{c}'')} [f_1(\bar{a}')] \right| > \delta \right] \right\},$$

where $f_1(\bar{a}') = 1$ if $\tilde{\gamma} \sim \tilde{\alpha}' \sim \tilde{\gamma}''$, 0 otherwise. Thus $\Pr_{(\bar{c}, \bar{c}', \bar{c}'')} [\text{val}_1 > \delta] \leq \delta/\varepsilon^3$, by the sampling of $\bar{A}(\bar{c}, \bar{c}'')/\bar{C}$ for all $\bar{c}, \bar{c}'' \in \bar{C}$. Likewise, $\Pr_{(\bar{c}, \bar{c}', \bar{c}'')} [\text{val}_2 > \delta] \leq \delta/\varepsilon^6$ follows from the same reasoning using the sampling of $\bar{A}(\bar{c}')/\bar{C}^2$ and the function $f_2(\bar{a}') = 1$ iff $\tilde{\gamma}' \sim \tilde{\alpha}'$. Finally,

$$\begin{aligned} \Pr_{(\bar{c}, \bar{c}', \bar{c}'')} [\text{val}_3 > 2\varepsilon^5] &\leq \varepsilon^{-6} \cdot \max_{\bar{c}' \in \bar{C}'_0} \left\{ \Pr_{\bar{c}, \bar{c}'' \sim \bar{c}} [\text{val}_3 > 2\varepsilon^5] \right\} \leq \frac{\varepsilon^{-11}}{2} \cdot \max_{\bar{c}' \in \bar{C}'_0} \left\{ \mathbb{E}_{\bar{c}, \bar{c}'' \sim \bar{c}} [\text{val}_3] \right\} \\ &= \frac{\varepsilon^{-11}}{2} \cdot \max_{\bar{c}' \in \bar{C}'_0} \left\{ 2 \cdot (1 - p(\bar{c}')) \right\} \leq \eta/\varepsilon^{11}. \end{aligned}$$

Now we show how these values figure into deriving the bound we need. The key point is that they let us bound $q(\bar{c}, \bar{c}'')$ in terms of $q(\bar{c}, \bar{c}') \cdot q(\bar{c}', \bar{c}'') \cdot \mu_{\bar{c}'}$, which is large when $\bar{c}, \bar{c}'' \in N(\bar{c}')$ and $\bar{c}' \in \bar{C}'_0$. We have:

$$\begin{aligned} q(\bar{c}, \bar{c}'') &= \Pr_{\bar{a}' \sim \bar{A}(\bar{c}, \bar{c}'')} [\tilde{\gamma} \sim \tilde{\alpha}' \sim \tilde{\gamma}'] \geq \Pr_{\bar{a}' \sim \bar{A}(\bar{c}, \bar{c}', \bar{c}'')} [\tilde{\gamma} \sim \tilde{\alpha}' \sim \tilde{\gamma}'] - \text{val}_1 \\ &\geq \Pr_{\substack{\bar{a}' \sim \bar{A}(\bar{c}, \bar{c}', \bar{c}'') \\ \bar{a}'' \sim \bar{A}(\bar{b}'')}} [\tilde{\gamma} \sim \tilde{\alpha} \sim \tilde{\gamma}' \sim \tilde{\alpha}'' \sim \tilde{\gamma}'' \ \& \ \tilde{\alpha} \approx \tilde{\alpha}' \approx \tilde{\alpha}'' \ \& \ \tilde{\gamma}' \sim \tilde{\alpha}'] - \text{val}_1 \\ &\geq \Pr_{\substack{\bar{a} \sim \bar{A}(\bar{c}, \bar{c}') \\ \bar{a}' \sim \bar{A}(\bar{c}, \bar{c}', \bar{c}'') \\ \bar{a}'' \sim \bar{A}(\bar{c}', \bar{c}'')}} [\tilde{\gamma} \sim \tilde{\alpha} \sim \tilde{\gamma}' \ \& \ \tilde{\gamma}' \sim \tilde{\alpha}' \ \& \ \tilde{\gamma}' \sim \tilde{\alpha}'' \sim \tilde{\gamma}'] - \text{val}_1 - \text{val}_3 \\ &\geq q(\bar{c}, \bar{c}') \cdot q(\bar{c}', \bar{c}'') \cdot \mu_{\bar{c}'} - \text{val}_1 - \text{val}_2 - \text{val}_3 \geq 4\varepsilon^5 - \text{val}_1 - \text{val}_2 - \text{val}_3. \end{aligned}$$

In the probability subscript in the second line, \bar{b} and \bar{b}'' are the restrictions of \bar{a}' to the lines spanned by (c, c') and (c', c'') , respectively. The result follows:

$$\Pr_{\substack{\bar{c}' \sim \bar{C}'_0 \\ \bar{c}, \bar{c}'' \sim N(\bar{c}')}} \left[q(\bar{c}, \bar{c}'') \geq \varepsilon^5 \mid |N(\bar{c}')| > \varepsilon^3 |\bar{C}| \right] \geq \Pr_{(\bar{c}, \bar{c}', \bar{c}'')} [\text{val}_1 + \text{val}_2 + \text{val}_3 \leq 3\varepsilon^5] \geq 1 - \sigma.$$

□

Claim 5 (Restated). *We have*

$$\Pr_{\substack{\bar{c} \sim \bar{C} \\ \bar{b}_1 \sim \bar{B}(\bar{c}) \\ \bar{c}'_1 \sim \bar{C}'(\bar{b}_1)}} [\mathbf{h}(\bar{c}) \sim \tilde{\beta}_1] \geq 1 - \tau,$$

where $\tilde{\beta} = \mathbf{g}_{\bar{c}'}(\bar{b})$, and where $\tau := (\sigma + 2\varepsilon^{-5}(\eta + \delta) + 2\delta)$. Recall $\mathbf{h}(\bar{c})$ is the distribution on $\Gamma_{\bar{C}}$ which draws $\bar{b}'_2 \sim \bar{B}(\bar{c})$, $\bar{c}'_2 \sim \bar{C}'(\bar{b}_2)$ and outputs $\mathbf{g}_{\bar{c}'_2}(\bar{b}_2)|_{\bar{c}}$.

Proof. We show $\Pr_{(\bar{c}, \bar{c}'_1, \bar{c}'_2, \bar{b}_1, \bar{b}_2)} [\tilde{\beta}_1 \sim \tilde{\beta}_2] \geq 1 - (\sigma + 2\varepsilon^{-5}(\eta + \delta))$, where the probability is over $\bar{c} \sim \bar{C}$, $\bar{c}'_1, \bar{c}'_2 \sim \bar{C}'$, $\bar{b}_1 \sim \bar{B}(\bar{c}, \bar{c}'_1)$, $\bar{b}_2 \sim \bar{B}(\bar{c}, \bar{c}'_2)$ and where $\tilde{\beta}_1 \sim \tilde{\beta}_2$ means that $\mathbf{g}_{\bar{c}'_1}(\bar{b}_1)$ and $\mathbf{g}_{\bar{c}'_2}(\bar{b}_2)$ agree at \bar{c} . The result then follows by the sampling of $\bar{B}(\bar{c})/\bar{C}$ for all $\bar{c} \in \bar{C}$. We have

$$\Pr_{(\bar{c}, \bar{c}'_1, \bar{c}'_2, \bar{b}_1, \bar{b}_2)} [\tilde{\beta}_1 \sim \tilde{\beta}_2] \geq \mathbb{E}_{\bar{c}'_1, \bar{c}'_2 \sim \bar{C}'} \left[\Pr_{(\bar{c}, \bar{b}_1, \bar{b}_2)} [\exists \bar{a} \in \bar{A}(\bar{b}_1, \bar{b}_2) \text{ st } \tilde{\gamma}'_1 \sim \tilde{\alpha} \sim \tilde{\gamma}'_2 \ \& \ \tilde{\beta}_1 \sim \tilde{\alpha} \sim \tilde{\beta}_2] \right]$$

$$\geq \mathbb{E}_{\bar{c}'_1, \bar{c}'_2 \sim \bar{C}'} \left[\Pr_{\substack{(\bar{c}, \bar{b}_1, \bar{b}_2) \\ \bar{a} \sim \bar{A}(\bar{b}_1, \bar{b}_2)}} \left[\tilde{\beta}_1 \sim \tilde{\alpha} \sim \tilde{\beta}_2 \mid \tilde{\gamma}'_1 \sim \tilde{\alpha} \sim \tilde{\gamma}'_2 \right] \right].$$

Let $\text{val} := \Pr_{(\bar{c}, \bar{b}_1, \bar{b}_2, \bar{a})} [\tilde{\beta}_1 \sim \tilde{\alpha} \sim \tilde{\beta}_2 \mid \tilde{\gamma}'_1 \sim \tilde{\alpha} \sim \tilde{\gamma}'_2]$ be shorthand for the quantity inside the expectation. We have

$$\begin{aligned} \text{val} &\geq 1 - \left[\Pr_{(\bar{c}, \bar{b}_1, \bar{b}_2, \bar{a})} [\tilde{\beta}_1 \not\sim \tilde{\alpha} \mid \tilde{\gamma}'_1 \sim \tilde{\alpha} \sim \tilde{\gamma}'_2] + \Pr_{(\bar{c}, \bar{b}_1, \bar{b}_2, \bar{a})} [\tilde{\beta}_2 \not\sim \tilde{\alpha} \mid \tilde{\gamma}'_1 \sim \tilde{\alpha} \sim \tilde{\gamma}'_2] \right] \\ &\geq 1 - \frac{1}{q(\bar{c}'_1, \bar{c}'_2)} \cdot \left[\Pr_{\substack{\bar{b}_1 \sim \bar{B}(\bar{c}'_1) \\ \bar{a} \sim \bar{A}(\bar{b}_1, \bar{c}'_2)}} [\tilde{\beta}_1 \not\sim \tilde{\alpha} \mid \tilde{\gamma}'_1 \sim \tilde{\alpha}] + \Pr_{\substack{\bar{b}_2 \sim \bar{B}(\bar{c}'_2) \\ \bar{a} \sim \bar{A}(\bar{b}_2, \bar{c}'_1)}} [\tilde{\beta}_2 \not\sim \tilde{\alpha} \mid \tilde{\gamma}'_2 \sim \tilde{\alpha}] \right] \end{aligned}$$

By definition of \bar{C}' , we have $\Pr_{\bar{c}'_1, \bar{c}'_2 \sim \bar{C}'} [q(\bar{c}'_1, \bar{c}'_2) < \varepsilon^5] \leq \sigma$ and also

$$\begin{aligned} \mathbb{E}_{\bar{c}'_1, \bar{c}'_2 \sim \bar{C}'} \left[\Pr_{\substack{\bar{b}_1 \sim \bar{B}(\bar{c}'_1) \\ \bar{a} \sim \bar{A}(\bar{b}_1, \bar{c}'_2)}} [\tilde{\beta}_1 \not\sim \tilde{\alpha} \mid \tilde{\gamma}'_1 \sim \tilde{\alpha}] \right] &\leq \max_{\bar{c}'_1 \in \bar{C}'} \left\{ \Pr_{\substack{\bar{b}_1 \sim \bar{B}(\bar{c}'_1) \\ \bar{a} \sim \bar{A}(\bar{b}_1)}} [\tilde{\beta}_1 \not\sim \tilde{\alpha} \mid \tilde{\gamma}'_1 \sim \tilde{\alpha}] + \delta \right\} \\ &= \max_{\bar{c}'_1 \in \bar{C}'} \{1 - p(\bar{c}'_1) + \delta\} \leq \eta + \delta. \end{aligned}$$

We have used the sampling of $\bar{A}(\bar{b})/\bar{C}$ for all $\bar{b} \in \bar{B}$, and that $p(\bar{c}'_1) \geq 1 - \eta$ since $\bar{c}'_1 \in \bar{C}'$. The result follows:

$$\mathbb{E}_{\bar{c}'_1, \bar{c}'_2 \sim \bar{C}'} [\text{val}] \geq (1 - \sigma) \cdot (1 - 2\varepsilon^{-5}(\eta + \delta)) \geq 1 - (\sigma + 2\varepsilon^{-5}(\eta + \delta)).$$

□

5 Affine Agreement

In this section we prove Lemma 2, restated in an expanded form below. We begin here by reducing Lemma 2 to Claims 7, 8 and 9, which we will prove in Section 5.2 after gathering some background on linearity/low-degree tests in Section 5.1. Recall that a function $T : \Gamma \rightarrow \Gamma$ is *affine* if there exists $u \in \mathbb{F}$ and $\Phi_0 \in \Gamma$ such that $T(\Phi) = u \cdot \Phi + \Phi_0$.

Lemma 2 (Restated). *Suppose $\{f_a\}_a \subset \{f : \Gamma_A \rightarrow \Gamma_A\}$, $h : \Gamma_C \rightarrow \Gamma_C$ and $G \subset A \times \Gamma$ are such that $|G| \geq 2\varepsilon \cdot |A \times \Gamma|$, and*

$$\Pr_{\substack{(a, \Phi) \sim G \\ \bar{c} \sim C(\bar{a})}} [\tilde{\gamma} \sim \tilde{\alpha}] \geq 1 - \zeta, \quad (6)$$

where (ε, ζ) are as in Lemma 1. Then there exists an affine map $T : \Gamma \rightarrow \Gamma$ such that

$$\Pr_{(a, \Phi) \sim G} [\tilde{\alpha} = T(\Phi)|_a] \geq 1/2.$$

Claim 7. *Let $(\varepsilon, \zeta, \{f_a\}, h, G)$ be as in the hypothesis of Lemma 2, so that (6) holds. Then there exist affine maps $\{T_c\}_{c \in C}$ with $T_c : \Gamma_C \rightarrow \Gamma_C$ such that $\Pr_{\bar{c} \sim \bar{C}} [\tilde{\gamma} = T_c(\gamma)] \geq 1 - \xi_7$ holds, where $\xi_7^2 := 32(d+1)(\zeta + \delta)$.*

Claim 8. Let $(\varepsilon, \zeta, \{f_a\}, h, G)$ be as in the hypothesis of Lemma 2, so that (6) holds, and let $\{\mathsf{T}_c\}$ be the family of affine maps promised by Claim 7. For each $c \in \mathsf{C}$, let $u_c, v_c \in \mathbb{F}$ be the scalars defining T_c , so $\mathsf{T}_c(\gamma) := u_c \cdot \gamma + v_c$. Then there exists $u \in \mathbb{F}$ such that $\Pr_{c \sim \mathsf{C}}[u_c = u] \geq 1 - \xi_8$, where $\xi_8 := (d+2)(\zeta + \delta) + 4\xi_7 + 2/|\mathbb{F}|$.

Claim 9. Let $(\varepsilon, \zeta, \{f_a\}, h, G)$ be as in the hypothesis of Lemma 2, so that (6) holds, and let $\{\mathsf{T}_c\}$ be the family of affine maps promised by Claim 7, with $\mathsf{T}_c(\gamma) := u_c \cdot \gamma + v_c$, as in Claim 8. Then there exists $\Phi_0 \in \Gamma$ such that $\Pr_{c \sim \mathsf{C}}[v_c = \Phi_0(c)] \geq 1 - \xi_9$, where $\xi_9^2 := 8(d+3)^2(\zeta + \xi_7 + \xi_8)$.

Proof of Lemma 2 Assuming Claims 7, 8 and 9. Let $(\varepsilon, \zeta, \{f_a\}, h, G)$ be as in the hypothesis of Lemma 2, so that (6) holds, and let $\{\mathsf{T}_c\}$ be the family of affine maps promised by Claim 7. Define the affine map $\mathsf{T} : \Gamma \rightarrow \Gamma$ by $\mathsf{T}(\Phi) := u \cdot \Phi + \Phi_0$, where $u \in \mathbb{F}$ and $\Phi_0 \in \Gamma$ are the quantities guaranteed by Claims 8 and 9, respectively. We have

$$\frac{3}{4} \leq \Pr_{\substack{(a, \Phi) \sim G \\ c \sim \mathsf{C}(a)}}[\tilde{\gamma} \sim \tilde{\alpha} \ \& \ \tilde{\gamma} = \mathsf{T}_c(\gamma) \ \& \ u_c = u \ \& \ v_c = \Phi_0(c)] \leq \Pr_{\substack{(a, \Phi) \sim G \\ c \sim \mathsf{C}(a)}}[\tilde{\alpha}|_c = \mathsf{T}(\Phi)|_c].$$

This follows from (6), Claims 7, 8, 9 and the sampling of $A \times \Gamma/\bar{\mathsf{C}}$. We have used the loose bound $1/4 \leq (\zeta + \xi_7 + \xi_8 + \xi_9 + \delta)$ where $\zeta > 0$ (resp. ξ_7, ξ_8, ξ_9) are the quantities from the statement of Lemma 2 (resp. Claims 7, 8, and 9), and $\delta > 0$ is the sampling parameter. It follows that $\Pr_{(a, \Phi) \sim G}[\tilde{\alpha} = \mathsf{T}(\Phi)|_a] \geq 1/2$, since whenever $\tilde{\alpha}$ and $\mathsf{T}(\Phi)|_a$ agree on half of the $c \in \mathsf{C}(a)$, they must be equal as they are both low degree. The lemma follows. \square

5.1 Linearity Testing Background

In this section we state three facts which we use in the next section to prove the claims. Throughout this section we use notations consistent with the rest of the paper. Additionally, in this section we use B as the set of lines in \mathbb{F}^m and Γ_{B} is the set of univariate polynomials over \mathbb{F} of degree at most d . Recall $\mathsf{T} : \Gamma_{\mathsf{C}} \rightarrow \Gamma_{\mathsf{C}}$ is *affine* if there exist coefficients $u, v \in \mathbb{F}$ such that $\mathsf{T}(x) = u \cdot x + v$ for all $x \in \Gamma_{\mathsf{C}}$. The first fact is standard and can be proved using linear algebraic methods.

Fact 4 (Linear Dependence of Polynomial Evaluations). Suppose $|\mathbb{F}| \geq d+2$. For any $b \in \mathsf{B}$ and distinct $c_0, \dots, c_{d+1} \in \mathsf{B}(c)$, there exist non-zero coefficients $r_0, r_1, \dots, r_{d+1} \in \mathbb{F}$ such that for all $\beta \in \Gamma_{\mathsf{B}}$,

$$\sum_{i=0}^{d+1} r_i \cdot \beta|_{c_i} = 0.$$

The second and third facts are proved in [RS96]. The second fact gives a sufficient condition for a function $f : \mathbb{F}^m \rightarrow \mathbb{F}$ being close to a multivariate low-degree polynomial.

Fact 5 (Robust Characterization of Low-Degree Functions). Fix $\kappa > 0$ such that $\kappa \leq \frac{1}{2(d+2)^2}$. If $f : \mathsf{C} \rightarrow \mathbb{F}$ is such that

$$\Pr_{b \sim \mathsf{B}}[\exists \beta \in \Gamma_{\mathsf{B}} \text{ st } \Pr_{c \sim \mathsf{C}(b)}[f(c) = \beta|_c] \geq 1 - \kappa] \geq 1 - \kappa,$$

then there exists $\Phi \in \Gamma$ such that $\Pr_{c \sim \mathsf{C}}[f(c) = \Phi(c)] \geq 1 - 2(d+3)\kappa$.

Fact 6 (Testing Affine Maps over Large Fields in High Soundness Regime). Fix $\kappa > 0$ such that $\kappa \leq \frac{1}{18}$. If $f : \Gamma_C \rightarrow \Gamma_C$ is such that

$$\Pr_{x,y,z \sim \Gamma_C} [f(x) + f(y+z) = f(x+y) + f(z)] \geq 1 - \kappa,$$

then there exists an affine $T : \Gamma_C \rightarrow \Gamma_C$ such that $\Pr_{x \sim \Gamma_C} [f(x) = T(x)] \geq 1 - 2\kappa$.

5.2 Proving the Claims

In this section we restate and prove the claims used to prove Lemma 2.

Notation. Throughout this section, we assume $\{f_a\}_a \subset \{f : \Gamma_A \rightarrow \Gamma_A\}$, $h : \Gamma_C \rightarrow \Gamma_C$ and $G \subset A \times \Gamma$ with $|G| \geq 2\varepsilon \cdot |A \times \Gamma|$ are such that (6) holds. Namely, we assume that the hypotheses of Lemma 2. We also use $\tilde{\gamma} = h(\bar{c})$ throughout.

Claim 7 (Restated). *There exist affine maps $\{T_c\}_{c \in C}$ such that $\Pr_{\bar{c} \sim \bar{C}} [\tilde{\gamma} = T_c(\gamma)] \geq 1 - \xi_7$.*

Proof. Consider the following distribution, \mathcal{D} on $C \times \Gamma_C^3$. Ultimately, the output of \mathcal{D} is just uniform, however the internal choices of \mathcal{D} help in our analysis. \mathcal{D} works as follows:

1. draw $b \sim B$ and distinct $c_0, c_1, \dots, c_{d+1} \sim C(b)$; let $r_0, \dots, r_{d+1} \in \mathbb{F}$ be the coefficients guaranteed by Fact 4;
2. draw $\gamma_0^0, \gamma_0^1, \dots, \gamma_d^0, \gamma_d^1 \sim \Gamma_C$; let $\bar{c}_{i,k} = (c_k, \gamma_k^i)$, and $\tilde{\gamma}_k^i = h(\bar{c}_{i,k})$ for $i = 0, 1$ and $k = 0, \dots, d$;
3. for $i, j \in \{0, 1\}$, let $\beta^{i,j} \in \Gamma_B$ be the unique polynomial that agrees with γ_0^i at c_0 and γ_k^j at c_k for all $k = 1, \dots, d$; let $\bar{b}_{i,j} = (b, \beta^{i,j})$;
4. for $i, j \in \{0, 1\}$, draw $(a_{i,j}, \Phi^{i,j}) \sim G(\bar{b}_{i,j})$ and set $\tilde{\alpha}^{i,j} = f_{a_{i,j}}(\Phi^{i,j}|_{a_{i,j}})$ and $\tilde{\beta}^{i,j} = \tilde{\alpha}^{i,j}|_b$;
5. let $(\tilde{\gamma}, \tilde{\gamma}', \tilde{\gamma}'', \tilde{\gamma}''') = (h(c_{d+1}, \gamma), h(c_{d+1}, \gamma'), h(c_{d+1}, \gamma''), h(c_{d+1}, \gamma'''))$, where

$$(\gamma, \gamma', \gamma'', \gamma''') = \left(\beta^{0,0}|_{c_{d+1}}, \beta^{1,0}|_{c_{d+1}}, \beta^{0,1}|_{c_{d+1}}, \beta^{1,1}|_{c_{d+1}} \right);$$

here $\beta|_c$ denotes the evaluation of the polynomial β at the point c ;

6. output $(c, x, y, z) = (c_{d+1}, \gamma, \gamma' - \gamma, \gamma'')$.

Note that the output of \mathcal{D} is uniform on $C \times \Gamma_C^3$. Indeed, c_{d+1} drawn in Step 1 is uniform since B/C is biregular. Moreover, given any fixed $\gamma_1^1, \dots, \gamma_k^1$, γ'' varies uniformly as γ_0^0 does. Then, given any fixing of $(\gamma_0^0, \gamma_1^1, \dots, \gamma_k^1)$, γ varies uniformly as $(\gamma_1^0, \dots, \gamma_k^0)$ does. Finally, given any fixing of γ_0^0 and $(\gamma_1^0, \gamma_1^1, \dots, \gamma_k^0, \gamma_k^1)$, γ' varies uniformly as γ_0^1 does.

Now, let E be the event: $\tilde{\gamma}_0^i \sim \tilde{\beta}^{i,j} \sim \tilde{\gamma}_k^j \forall (i, j, k) \in \{0, 1\}^2 \times \{1, \dots, d\}$, where the $\tilde{\gamma}_0^i$, $\tilde{\beta}^{i,j}$, and $\tilde{\gamma}_k^j$ are the internal values drawn during steps 2 and 4. By the assumptions of Lemma 2 and the

sampling of $A \times \Gamma/\bar{B}$, we have $\Pr_{\bar{b}, \bar{c}, (a, \Phi)}[\tilde{\gamma} \sim \tilde{\alpha}] \geq 1 - \zeta - \delta$, where the probability is over $\bar{b} \sim \bar{B}$, $\bar{c} \sim \bar{C}(\bar{b})$, $(a, \Phi) \sim G(\bar{b})$. It follows from the union bound that $\Pr_{\mathcal{D}}[\mathbf{E}] \geq 1 - \xi_7^2/8$ (substituting $\xi_7^2 = 32(d+1)(\zeta + \delta)$), since each $(\bar{b}_{i,j}, \bar{c}_{i,0}, \mathbf{a}_{i,j}, \Phi^{i,j})$ and $(\bar{b}_{i,j}, \bar{c}_{j,k}, \mathbf{a}_{i,j}, \Phi^{i,j})$ are, individually, drawn in this way for each $(i, j, k) \in \{0, 1\}^2 \times \{0, \dots, d\}$.

We complete the proof by showing that whenever the sampling of $(c, x, y, z) \sim \mathcal{D}$ is such that \mathbf{E} occurs, it holds that $h(c, x) + h(c, y + z) = h(c, x + y) + h(c, z)$. Together with Fact 6, this implies that there is a family of affine maps $\{\mathbb{T}_c\}_{c \in \mathcal{C}}$ such that

$$\Pr_{c \sim \mathcal{C}} \left[\Pr_{\gamma \sim \Gamma_c} [\tilde{\gamma} = \mathbb{T}_c(\gamma)] \geq 1 - \frac{\xi_7}{2} \right] \geq 1 - \frac{\xi_7}{2},$$

which implies the claim.

So it suffices to show that

$$\gamma - \gamma' = \gamma'' - \gamma''' \text{ and } \tilde{\gamma} - \tilde{\gamma}' = \tilde{\gamma}'' - \tilde{\gamma}'''$$

both hold whenever \mathbf{E} occurs (the first equality always holds, the second holds whenever \mathbf{E} occurs). This follows from Fact 4. The first equality holds since the $\beta^{i,j}$ are low-degree and for all (i, j, k) , γ_0^i and γ_k^j are the evaluations of $\beta^{i,j}$ at c_0 and c_k , respectively. Thus Fact 4 gives

$$\begin{aligned} r_0 \cdot \gamma_0^0 + \left(\sum_{k=1}^d r_k \cdot \gamma_k^0 \right) + r_{d+1} \cdot \gamma &= 0; & r_0 \cdot \gamma_0^1 + \left(\sum_{k=1}^d r_k \cdot \gamma_k^0 \right) + r_{d+1} \cdot \gamma' &= 0; \\ r_0 \cdot \gamma_0^0 + \left(\sum_{k=1}^d r_k \cdot \gamma_k^1 \right) + r_{d+1} \cdot \gamma'' &= 0; & r_0 \cdot \gamma_0^1 + \left(\sum_{k=1}^d r_k \cdot \gamma_k^1 \right) + r_{d+1} \cdot \gamma''' &= 0, \end{aligned}$$

which simplifies to $\gamma - \gamma' = \gamma'' - \gamma'''$ since $r_{d+1} \neq 0$. Likewise, for the second equality, the $\tilde{\beta}^{i,j}$ are low degree and when \mathbf{E} occurs, the $\tilde{\gamma}_0^i$ and $\tilde{\gamma}_k^j$ are the evaluations of $\tilde{\beta}^{i,j}$ at c_0 and c_k . As above, this implies $\tilde{\gamma} - \tilde{\gamma}' = \tilde{\gamma}'' - \tilde{\gamma}'''$. \square

Claim 8 (Restated). *Let $\{\mathbb{T}_c\}$ be the family of affine maps promised by Claim 7; for each $c \in \mathcal{C}$, let $\mathbb{T}_c(\gamma) := u_c \cdot \gamma + v_c$ for $u_c, v_c \in \mathbb{F}$. Then there exists $u \in \mathbb{F}$ such that $\Pr_{c \sim \mathcal{C}}[u_c = u] \geq 1 - \xi_8$, where $\xi_8 = (d+2)(\zeta + \delta) + 4\xi_7 + 2/|\mathbb{F}|$.*

Proof. We prove that $\Pr_{c, c' \sim \mathcal{C}}[u_c = u_{c'}] \geq 1 - \xi_8$ which suffices since

$$\Pr_{c, c' \sim \mathcal{C}}[u_c = u_{c'}] = \sum_{u \in \mathbb{F}} p_u^2 \leq \max \{p_u : u \in \mathbb{F}\},$$

where $p_u := \Pr_{c \sim \mathcal{C}}[u_c = u]$ is shorthand. As in the previous proof, we describe a distribution \mathcal{D}' on \mathcal{C}^2 :

1. draw $\mathbf{b} \sim \mathbf{B}$ and distinct $c_0, c_1, \dots, c_{d+1} \sim \mathcal{C}(\mathbf{b})$; let $r_0, \dots, r_{d+1} \in \mathbb{F}$ be the coefficients guaranteed by Fact 4; let $u_0, u_{d+1} \in \mathbb{F}$ denote the linear terms of \mathbb{T}_{c_0} and $\mathbb{T}_{c_{d+1}}$, respectively;
2. draw $\gamma_0^0, \gamma_0^1, \gamma_k \sim \Gamma_c$ for $k = 1, \dots, d$; let $\bar{c}_{i,0} = (c_0, \gamma_0^i)$ for $i = 0, 1$ and $\bar{c}_k = (c_k, \gamma_k)$ for $k = 1, \dots, d$; let $\tilde{\gamma}_0^i = h(\bar{c}_{i,0})$ and $\tilde{\gamma}_k = h(\bar{c}_k)$;

3. for $i \in \{0, 1\}$, let $\beta^i \in \Gamma_B$ be the unique polynomial that agrees with γ_0^i at c_0 and γ_k at c_k for all $k = 1, \dots, d$; let $\bar{b}_i = (b, \beta^i)$;
4. for $i \in \{0, 1\}$, draw $(a_i, \Phi^i) \sim G(\bar{b}_i)$ and set $\tilde{\alpha}^i = f_{a_i}(\Phi^i|_{a_i})$ and $\tilde{\beta}^i = \tilde{\alpha}^i|_b$;
5. let $(\tilde{\gamma}, \tilde{\gamma}') = (h(c_{d+1}, \gamma), h(c_{d+1}, \gamma'))$, where $(\gamma, \gamma') = (\beta^0|_{c_{d+1}}, \beta^1|_{c_{d+1}})$;
6. output $(c, c') = (c_0, c_{d+1})$.

Note that \mathcal{D}' outputs two random points on a random line, which is within statistical distance $2/|\mathbb{F}|$ of uniform on \mathbb{C}^2 . Let \mathbf{E}' be the event:

1. $\tilde{\gamma}_0^i \sim \tilde{\beta}^i \sim \tilde{\gamma}_k \forall (i, k) \in \{0, 1\} \times \{1, \dots, d\}$; and
2. $(\tilde{\gamma}_0^0, \tilde{\gamma}_0^1, \tilde{\gamma}, \tilde{\gamma}') = (\mathbb{T}_{c_0}(\gamma_0^0), \mathbb{T}_{c_0}(\gamma_0^1), \mathbb{T}_{c_{d+1}}(\gamma), \mathbb{T}_{c_{d+1}}(\gamma'))$

The first condition occurs with probability at least $1 - (d+2)(\zeta + \delta)$; as in the proof of Claim 7, this follows from (6), the sampling of $A \times \Gamma/\bar{B}$, and a union bound. The second condition occurs with probability at least $1 - 4\xi_7$, by Claim 7. Upon substituting $\xi_8 = (d+2)(\zeta + \delta) + 4\xi_7 + 2/|\mathbb{F}|$, we get $\Pr_{(c, c') \sim \mathbb{C}^2}[\mathbf{E}'] \geq \Pr_{\mathcal{D}'}[\mathbf{E}'] - 2/|\mathbb{F}| \geq 1 - \xi_8$. As in the proof of Claim 7, Fact 4 gives

$$r_0 \cdot (\gamma_0^0 - \gamma_0^1) + r_{d+1} \cdot (\gamma - \gamma') = 0 = r_0 \cdot (\tilde{\gamma}_0^0 - \tilde{\gamma}_0^1) + r_{d+1} \cdot (\tilde{\gamma} - \tilde{\gamma}').$$

Substituting $(\tilde{\gamma}_0^0 - \tilde{\gamma}_0^1) = u_0 \cdot (\gamma_0^0 - \gamma_0^1)$ and $(\tilde{\gamma} - \tilde{\gamma}') = u_{d+1} \cdot (\gamma - \gamma')$ gives $r_{d+1}(u_{d+1} - u_0)(\gamma - \gamma') = 0$ which means $u_{d+1} = u_0$ since $r_{d+1} \neq 0$ and $\gamma \neq \gamma'$. Thus, $\Pr_{c, c' \sim \mathbb{C}}[u_c = u_{c'}] \geq 1 - \xi_8$. \square

Claim 9 (Restated). *Let $\{\mathbb{T}_c\}$ be the family of affine maps promised by Claim 7. Then there exists $\Phi_0 \in \Gamma$ with $\Pr_{c \sim \mathbb{C}}[\mathbb{T}_c(\mathbf{0}) = \Phi_0(c)] \geq 1 - \xi_9$, where $\xi_9^2 = 8(d+3)^2(\zeta + \xi_7 + \xi_8)$.*

Proof. Let $v : \mathbb{C} \rightarrow \mathbb{F}$ as a function mapping $c \mapsto v_c = \mathbb{T}_c(\mathbf{0})$. Let $\xi := \frac{\xi_9}{2(d+3)}$. We will show that

$$\Pr_{b \sim B} \left[\exists \tilde{\beta}' \in \Gamma_B \text{ st } \Pr_{c \sim \mathbb{C}(b)} [v_c = \tilde{\beta}'|_c] \geq 1 - \xi \right] \geq 1 - \xi. \quad (7)$$

The claim then follows from Fact 5. Towards establishing (7), note that

$$\Pr_{\substack{(a, \Phi) \sim G \\ b \sim B(a) \\ c \sim \mathbb{C}(b)}} [v_c = \tilde{\beta}|_c - u \cdot \beta|_c] \geq 1 - (\zeta + \xi_7 + \xi_8) \geq 1 - \xi(\xi - \delta),$$

where $\beta = \Phi|_b$ and $\tilde{\beta} = \tilde{\alpha}|_b$; we have used $\xi(\xi - \delta) \geq \xi^2/2 = \zeta + \xi_7 + \xi_8$. This follows immediately from (6) and Claims 7 and 8. By an averaging argument,

$$\Pr_{\substack{(a, \Phi) \sim G \\ b \sim B(a)}} \left[\Pr_{c \sim \mathbb{C}(b)} [v_c = \tilde{\beta}'|_c] \geq 1 - \xi \right] \geq 1 - \xi + \delta,$$

where $\tilde{\beta}' = \tilde{\beta} - u \cdot \beta$. The bound (7) now follows from the sampling of $A \times \Gamma/B$. \square

6 A Locally Testable, Non-Malleable Code

In this section, we give a construction of a locally testable non-malleable code against coordinate wise tampering. To build our code, we take the LTMN reduction, $(E_{\text{LTMN}}, D_{\text{LTMN}}, T_{\text{LTMN}})$, from coordinate-wise tampering to affine tampering, from section 1.4 and compose it with a new non-malleable code, $(E_{\text{aff}}, D_{\text{aff}})$, against affine tampering.

6.1 A Simple Non-malleable Code against Affine Tampering

We begin with a new constant rate, non-malleable code against affine tampering. This result is not new, several prior works [ADL14, CZ14, Li16, CL17] give such codes, however, our construction is considerably simpler than those prior.

Notations. Let \mathbb{F} be a finite field and \mathbb{K}/\mathbb{F} a degree 3 extension, so $\mathbb{K} = \mathbb{F}[x]/(p(x))$ for an irreducible cubic polynomial $p(x) = x^3 - e_2x^2 - e_1x - e_0$. Thus \mathbb{K} is a 3-dimensional \mathbb{F} -vector space with basis $\{1, \sigma, \sigma^2\}$, where $\sigma \in \mathbb{K}$ is a root of $p(x)$. The ‘multiplication by σ ’ map $\mathbb{F}^3 \rightarrow \mathbb{F}^3$ is linear, specified over this basis by the matrix

$$\Sigma = \begin{bmatrix} 0 & 0 & e_0 \\ 1 & 0 & e_1 \\ 0 & 1 & e_2 \end{bmatrix} \in \mathbb{F}^{3 \times 3}.$$

Our code makes use of an ε -high entropy encoding, (E, D) , with codeword space \mathbb{F} , such that for all m, c^* , $\Pr_{c \sim E(m)}[c = c^*] \leq \varepsilon$. Such codes can be trivially constructed by appending a message with a random string of length $\log(1/\varepsilon)$.

Construction. Let (E, D) be an ε -high entropy code with message space \mathcal{M} and codeword space \mathbb{F} , and let $m \in \mathcal{M}$.

- $E_{\text{aff}}(m)$: Draw $r \sim \mathbb{F}$; $w \sim E(m)$ and output $w + r \cdot \sigma + wr \cdot \sigma^2 \in \mathbb{K}$.
- $D_{\text{aff}}(c)$: Parse $c = c_0 + c_1 \cdot \sigma + c_2 \cdot \sigma^2$; if $c_0 \cdot c_1 = c_2$, output $m = D(c_0)$; if not, output \perp .

Theorem 4. Fix $\varepsilon > 0$, and let (E, D) be an ε -high entropy code with message space \mathcal{M} and codeword space \mathbb{F} . Then $(E_{\text{aff}}, D_{\text{aff}})$ is a $(2\varepsilon + 2/|\mathbb{F}|)$ -non-malleable code against affine tampering functions.

Proof. Fix an affine map f given by $f(x) = sx + t$ where $s, t, x \in \mathbb{K}$ and fix any message $m \in \mathcal{M}$. Parse $s = s_0 + s_1 \cdot \sigma + s_2 \cdot \sigma^2$ and $t = t_0 + t_1 \cdot \sigma + t_2 \cdot \sigma^2$. To prove the theorem, we exhibit a trivial tampering function g_f (i.e., either constant or the identity) such that the tampering distribution $(D_{\text{aff}} \circ f \circ E_{\text{aff}})(m)$ outputs $g_f(m)$ with probability at least $1 - 2\varepsilon - 2/|\mathbb{F}|$. The trivial function g_f is f if f is either the identity or a constant function mapping to a valid codeword, and is the constant \perp function otherwise. Specifically, if $(s, t) = (1, 0)$, g_f is the identity; if $s = 0$ and $t_0 \cdot t_1 = t_2$, g_f is the constant function mapping everything to t ; otherwise g_f is the constant \perp

function. The key point, is that for all $m \in \mathcal{M}$, the distribution $f(\mathbb{E}_{\text{aff}}(m))$ draws $w \sim \mathbb{E}(m)$, $r \sim \mathbb{F}$ and outputs

$$S \begin{bmatrix} w \\ r \\ wr \end{bmatrix} + \begin{bmatrix} t_0 \\ t_1 \\ t_2 \end{bmatrix} = \begin{bmatrix} t_0 + s_0w + e_0s_2r + (e_0s_1 + e_0e_2s_2)wr \\ t_1 + s_1w + (s_0 + e_1s_2)r + (e_1s_1 + s_2e_0 + s_2e_1e_2)wr \\ t_2 + s_2w + (s_1 + e_2s_2)r + (s_0 + e_2s_1 + s_2e_2^2 + s_2e_1)wr \end{bmatrix} =: \begin{bmatrix} C_0(w, r) \\ C_1(w, r) \\ C_2(w, r) \end{bmatrix},$$

where $S \in \mathbb{F}^{3 \times 3}$ is the ‘multiplication by s ’ matrix: $S = s_0 \cdot \mathbb{1} + s_1 \cdot \Sigma + s_2 \cdot \Sigma^2$. In the above, we have defined bilinear (i.e., of the form $a + bx + cy + dxy$) polynomials $C_0, C_1, C_2 \in \mathbb{F}[x, y]$. Note that if $C_0(x, y) \cdot C_1(x, y) \not\equiv C_2(x, y)$ as polynomials, then $C_0(w, r) \cdot C_1(w, r) = C_2(w, r)$ holds with probability at most $2\varepsilon + 2/|\mathbb{F}|$, in which case $(D_{\text{aff}} \circ f \circ \mathbb{E}_{\text{aff}})(m) = \perp$ with high probability. This follows immediately from Schwartz-Zippel and the low entropy property of (\mathbb{E}, D) . Therefore, in order to prove the theorem, it suffices to show that if $C_0(x, y) \cdot C_1(x, y) \equiv C_2(x, y)$ holds, then either $s = 0$ or $(s, t) = (1, 0)$. We assume $C_0(x, y) \cdot C_1(x, y) \equiv C_2(x, y)$ holds, and we prove the following three items:

1. either $s_1 = 0$ or $s_2 = 0$;
2. $s_1 = 0 \Leftrightarrow s_2 = 0$;
3. if $s_1 = s_2 = 0$ then either $s_0 = 0$ or $s_0 = 1$ and $t_0 = t_1 = t_2 = 0$.

The third point is easiest: if $C_0(x, y) \cdot C_1(x, y) \equiv C_2(x, y)$ and $s_1 = s_2 = 0$ then plugging gives

$$(t_0 + s_0x) \cdot (t_1 + s_0y) = t_2 + s_0xy,$$

from which it follows that either $s_0 = 0$ or $s_0 = 1$ and $t_i = 0$ for all $i = 0, 1, 2$. To prove the first point, note that if $C_0(x, y) \cdot C_1(x, y) \equiv C_2(x, y)$, then $s_0 \cdot s_1 = 0$ (since the x^2 coefficient in C_2 is zero). If $s_1 = 0$ we are done; if $s_0 = 0$ then $e_0e_1s_2^2 = 0$ (since y^2 coefficient in C_2 is zero), which implies $e_1s_2 = 0$ since $e_0 \neq 0$ (else $p(x)$ is reducible). If $s_2 = 0$ we are done; if $e_1 = 0$ then $e_0^2s_2^2 = 0$ (since xy^2 coefficient in C_2 is zero). Again, $e_0 \neq 0$ so $s_2 = 0$ so the first point follows.

Finally, for the second point, assume $s_1 = 0$. Then $s_0s_2 \cdot (e_0 + e_1e_2) = 0$ since the coefficient of $x^2y = 0$ in C_2 . Note $e_0 \neq -e_1e_2$ since otherwise $p(x)$ is reducible: $p(x) = (x - e_2)(x^2 - e_1)$. However, if $s_0 = 0$ then, as shown in the proof of the first point, $s_2 = 0$; therefore $s_1 = 0$ implies $s_2 = 0$. Conversely, if $s_2 = 0$ then $e_0s_0s_1 = 0$ (coefficient of xy^2 in C_2 is zero), so $s_0s_1 = 0$. If $s_0 = 0$ then $e_0s_1^2 = 0$ (coefficient of x^2y in C_2 is zero). Thus $s_2 = 0$ implies $s_1 = 0$, and we are done. \square

Remark. In our LTNM code in the next section, we will use $(\mathbb{E}_{\text{aff}}, D_{\text{aff}})$ to encode a random $w \in \mathbb{F}$ and so the high entropy encoding is not necessary. The precise claim we use is stated below. The proof is the same as above since if $C_0(x, y) \cdot C_1(x, y) \not\equiv C_2(x, y)$ as polynomials, then $C_0(w, r) \cdot C_1(w, r) = C_2(w, r)$ holds with probability at most $4/|\mathbb{F}|$ over $w, r \sim \mathbb{F}$.

Claim 10. *Let $f : \mathbb{K} \rightarrow \mathbb{K}$ be affine of the form $f(x) = sx + t$ for $s, t \in \mathbb{K}$ such that $s \neq 0$ and $(s, t) \neq (1, 0)$. Then $\Pr_{w, r \sim \mathbb{F}} [D_{\text{aff}}(f(w + r \cdot \sigma + wr \cdot \sigma^2)) \neq \perp] \leq 4/|\mathbb{F}|$.*

6.2 A LTNM Code via Composition

Composition Overview. The local test of our main construction from Section 1.4 passes whenever codewords are tampered by a coordinate-wise affine function. Thus, in order to use our main construction to build a fully LTNM code, we must modify the test in such a way so that it fails whenever a non-trivial affine tampering function is used. We do this in two steps. First, we modify the local tester so that it locally decodes a specified polynomial evaluation. Second, the tester checks that the evaluation recovered is a valid codeword of $(E_{\text{aff}}, D_{\text{aff}})$, if not it outputs \perp . Essentially, the reason this works is that the local decoder will output \perp unless the codeword is tampered with an affine function, in which case the evaluation recovered is an affine function of the original evaluation. If the original evaluation is a random valid codeword of $(E_{\text{aff}}, D_{\text{aff}})$ then by Claim 10, the recovered evaluation is a valid codeword only if the affine tampering function is trivial.

Notations. As in the previous section, let \mathbb{K}/\mathbb{F} be a degree 3 extension with \mathbb{F} -basis $\{1, \sigma, \sigma^2\}$. Let $k \geq 5$ and $d \geq 2$. As in the rest of the paper, let A be the set of 3-planes in \mathbb{F}^k and $C = \mathbb{F}^k$. In this section, we use B and \bar{A} to denote the set of lines and 4-planes respectively (note, the second usage is different from rest of the paper where we used \bar{A} to denote $A \times \Gamma_A$). Let $p = (1, 0, \dots, 0) \in \mathbb{F}^k$.

Construction. Let $E_{\text{aff}}()$ denote the procedure which draws $w, r \sim \mathbb{F}$, and outputs the value $w + r \cdot \sigma + wr \cdot \sigma^2 \in \mathbb{K}$; let D_{aff} be the decoding algorithm from the previous section. Let $m \in \mathbb{K}$ be a message.

- **Enc(m):** Draw $v \sim E_{\text{aff}}()$; and $\Phi \sim \Gamma$ such that $\Phi(0) = m$ and $\Phi(p) = v$; output $\{(a, \Phi_a)\}_{a \in A}$.
- **Dec($\{(a, \alpha)\}_{a \in A}$):** Find $\Phi \in \Gamma$ such that $(a, \alpha) = (a, \Phi|_a)$ for all $a \in A$. If such Φ exists, and if $D_{\text{aff}}(\Phi(p)) \neq \perp$, output $m = \Phi(0)$, otherwise output \perp .
- **Test($\{(a, \alpha)\}_{a \in A}$):** Draw $b \sim B(p)$, $c_1, c_2, c_3 \sim C(b)$, $c, c' \sim C$, $a_1 \sim A(c, c_1)$, $a_2 \sim A(c, c', c_2)$, $a_3 \sim A(c', c_3)$. Read (a_1, α_1) , (a_2, α_2) , (a_3, α_3) and do the following.
 - 1) Check that $\alpha_1|_c = \alpha_2|_c$ and $\alpha_2|_{c'} = \alpha_3|_{c'}$; if not output 0; if so use interpolation to recover $\beta \in \Gamma_B$, the unique degree 2 polynomial such that $\beta|_{c_i} = \alpha_i|_{c_i}$ for $i = 1, 2, 3$; let $v = \beta|_p$.
 - 2) If $D_{\text{aff}}(v) \neq \perp$, output 1; otherwise output 0.

Theorem 5. Let ℓ, ε as in theorem 1. Then the code (Enc, Dec, Test) above is a (ℓ, ε') -locally testable, non-malleable code against \mathcal{F} , the family of coordinate-wise tampering functions where $\varepsilon' = \mathcal{O}(\varepsilon^{1/2})$.

Proof. Fix a tampering function $f = \{f_a\}_a \in \mathcal{F}$. We prove that (Enc, Dec, Test) is non-malleable using the sufficient conditions of Claim 1. The first three conditions are trivial.

1. The coordinate-wise \perp function is in \mathcal{H} ; for all other tampering functions in \mathcal{H} , the test passes when a valid codeword is tampered with $h \in \mathcal{H}$.

2. For all distinct $\{h_a\}_a, \{h'_a\}_a \in \mathcal{H}$, $\Pr_{\Phi, a}[h_a(\Phi|_a) = h'_a(\Phi|_a)] = \mathcal{O}(|\mathbb{F}|^{-1})$, as before.
3. for all $\{h_a\}_a \in \mathcal{H}$, the distribution which draws (v, Φ) like $\text{Enc}(m)$ and

$$\text{rand} := (b, c_1, c_2, c_3, c, c', a_1, a_2, a_3)$$

like Test , and outputs $((a_1, h_{a_1}(\Phi|_{a_1})), (a_2, h_{a_2}(\Phi|_{a_2})), (a_3, h_{a_3}(\Phi|_{a_3})))$ is within statistical distance $\mathcal{O}(|\mathbb{F}|^{-1})$ of the distribution which draws $\Phi \sim \Gamma$ uniformly, rand as in Test and gives the same output. This is because when $d = 2$ and $k \geq 5$, the statistical distance between these two distributions is upper bounded by the probability that at least one of the a_i intersects the line through 0 and p (which is $\mathcal{O}(|\mathbb{F}|^{-1})$ when $k \geq 5$), since the number of degrees of freedom in Φ with $\Phi(0)$ and $\Phi(p)$ fixed is more than the number of linear constraints imposed by the $h_{a_i}(\Phi|_{a_i})$.

Therefore, it remains to exhibit a list $L_f \subset \mathcal{H}$ of size at most $|L_f| \leq \ell$ such that $\text{val} \leq \mathcal{O}(\varepsilon^{1/2})$ where

$$\text{val} := \Pr_{\Phi, \text{rand}} \left[\text{Test passes} \ \& \ (\tilde{\alpha}_1, \tilde{\alpha}_2, \tilde{\alpha}_3) \notin \left\{ (h_{a_1}(\Phi|_{a_1}), h_{a_2}(\Phi|_{a_2}), h_{a_3}(\Phi|_{a_3})) : \{h_a\}_a \in L_f \right\} \right],$$

where $\tilde{\alpha}_i = f_{a_i}(\Phi|_{a_i})$. In the course of the proof of Theorem 1 from Section 2, a similar list $L'_f \subset \mathcal{G}$ of size at most $|L'_f| \leq \ell$ was constructed such that

$$\Pr_{\Phi, (c, a_1, a_2)} \left[\tilde{\alpha}_1|_c = \tilde{\alpha}_2|_c \ \& \ (\tilde{\alpha}_1, \tilde{\alpha}_2) \notin \left\{ (g_{a_1}(\Phi|_{a_1}), g_{a_2}(\Phi|_{a_2})) : \{g_a\}_a \in L'_f \right\} \right] \leq \varepsilon,$$

where this probability is over $\Phi \sim \Gamma$ and $c \sim C$, $a_1, a_2 \sim A(c)$. Our list $L_f \subset \mathcal{H}$ is the set of trivial (*i.e.*, constant or affine) $\{g_a\}_a \in L'_f$. The quantity val can now be bounded

$$\text{val} \leq \Pr_{\Phi, \text{rand}} [\mathbf{E}_1 \vee \mathbf{E}'_1 \vee \mathbf{E}_2 \vee \mathbf{E}_3]$$

for the following events:

$$\mathbf{E}_1: \tilde{\alpha}_1|_c = \tilde{\alpha}_2|_c \ \& \ (\tilde{\alpha}_1, \tilde{\alpha}_2) \notin \left\{ (g_{a_1}(\Phi|_{a_1}), g_{a_2}(\Phi|_{a_2})) : \{g_a\}_a \in L'_f \right\};$$

$$\mathbf{E}'_1: \tilde{\alpha}_2|_{c'} = \tilde{\alpha}_3|_{c'} \ \& \ (\tilde{\alpha}_2, \tilde{\alpha}_3) \notin \left\{ (g'_{a_2}(\Phi|_{a_2}), g'_{a_3}(\Phi|_{a_3})) : \{g'_a\}_a \in L'_f \right\};$$

$$\mathbf{E}_2: \text{the } \{g_a\}_a, \{g'_a\}_a \in \mathcal{G} \text{ which agree with } f \text{ from } \mathbf{E}_1 \text{ and } \mathbf{E}'_1 \text{ are distinct and such that } g_{a_2}(\Phi|_{a_2}) = g'_{a_2}(\Phi|_{a_2});$$

$$\mathbf{E}_3: \text{the same } \{g_a\}_a \in \mathcal{G} \text{ results from } \mathbf{E}_1 \text{ and } \mathbf{E}'_1; \text{ this } \{g_a\}_a \in \mathcal{G} \text{ is non-trivial, but the affine check passes: } D_{\text{aff}}(\tilde{v}) \neq \perp.$$

The marginal distribution on a_2 from rand is uniform, so $\Pr_{\Phi, \text{rand}}[\mathbf{E}_2] = \mathcal{O}(|\mathbb{F}|^{-1})$. By Claim 10, $\Pr_{\Phi, \text{rand}}[\mathbf{E}_3] \leq 4/|\mathbb{F}|$. We prove $\Pr_{\Phi, \text{rand}}[\mathbf{E}_1] \leq \varepsilon^{1/2} + \mathcal{O}(|\mathbb{F}|^{-1})$. The same holds for \mathbf{E}'_1 , and the result follows. Towards bounding $\Pr_{\Phi, \text{rand}}[\mathbf{E}_1]$, note that drawing $\Phi \sim \Gamma$ uniformly, rather than uniformly subject to $\Phi(0) = m$ and $\Phi(p) = v$ changes the probability by at most $\mathcal{O}(|\mathbb{F}|^{-1})$. Therefore, in the calculation below, we assume $\Phi \sim \Gamma$. We have

$$\Pr_{\Phi, \text{rand}} [\mathbf{E}_1]^2 = \mathbb{E}_{\Phi, c \sim C, a_2 \sim A(c)} \left[\Pr_{a_1 \sim \text{rand}(c, a_2)} [\mathbf{E}_1] \right]^2 \leq \mathbb{E}_{\Phi, c, a_2} \left[\Pr_{a_1 \sim \text{rand}(c, a_2)} [\mathbf{E}_1]^2 \right]$$

$$\leq \mathbb{E}_{\Phi, c, a_2} \left[\Pr_{a_1, a_3 \sim \text{rand}(c, a_2)} [\tilde{\alpha}_1|c = \tilde{\alpha}_2|c = \tilde{\alpha}_3|c \ \& \ (\tilde{\alpha}_1, \tilde{\alpha}_2, \tilde{\alpha}_3) \notin L'_f] \right] + \mathcal{O}(|\mathbb{F}|^{-1}),$$

where “ $(\tilde{\alpha}_1, \tilde{\alpha}_2, \tilde{\alpha}_3) \notin L'_f$ ” is shorthand for

$$(\tilde{\alpha}_1, \tilde{\alpha}_2, \tilde{\alpha}_3) \notin \{(\mathbf{g}_{a_1}(\Phi|_{a_1}), \mathbf{g}_{a_2}(\Phi|_{a_2}), \mathbf{g}_{a_3}(\Phi|_{a_3})) : \{\mathbf{g}_a\}_a \in L'_f\}$$

and the $\mathcal{O}(|\mathbb{F}|^{-1})$ term in the second line accounts for the case when there are $\{\mathbf{g}_a\}_a, \{\mathbf{g}'_a\}_a \in L'_f$ such that $\mathbf{g}_{a_2}(\Phi|_{a_2}) = \mathbf{g}'_{a_2}(\Phi|_{a_2})$ holds. Note that if $\tilde{\alpha}_1 = \mathbf{g}_{a_1}(\Phi|_{a_1})$, and $\tilde{\alpha}_2 \neq \mathbf{g}_{a_2}(\Phi|_{a_2})$, then $\tilde{\alpha}_1|c = \tilde{\alpha}_2|c$ occurs with probability $\mathcal{O}(|\mathbb{F}|^{-1})$. It follows that

$$\Pr_{\Phi, \text{rand}} [\mathbf{E}_1]^2 \leq \Pr_{\substack{\Phi, c, a_2 \\ a_1, a_3 \sim \text{rand}(c, a_2)}} [\tilde{\alpha}_1|c = \tilde{\alpha}_3|c \ \& \ (\tilde{\alpha}_1, \tilde{\alpha}_3) \notin L'_f] + \mathcal{O}(|\mathbb{F}|^{-1}).$$

Therefore, it suffices to show that for all $c \in C$, the distribution which draws $a_2 \sim A(c)$, $a_1, a_3 \sim \text{rand}(c, a_2)$ and outputs (a_1, a_3) is within statistical distance $\mathcal{O}(|\mathbb{F}|^{-1})$ of uniform on $A(c)^2$. The distribution $\text{rand}(c, a_2)$ draws $c_2 \sim C(a_2)$, $c_1 \sim C(b)$, where b is the line through p and c_2 , and outputs $a_1 \sim A(c, c_1)$. This is equivalent to drawing $c_1 \sim C(\bar{a}_2)$ and outputting $a_1 \sim A(c, c_1)$, where \bar{a}_2 is the 4-plane containing a_2 and p . Thus the distribution which draws $a_2 \sim A(c)$ and then $a_1, a_3 \sim \text{rand}(c, a_2)$, outputting (a_1, a_2, a_3) can be equivalently described by drawing $a_1, a_3 \sim A(c)$, $c_i \sim C(a_i)$ for $i = 1, 3$, $\bar{a}_2 \sim \bar{A}(c, p, c_1, c_3)$ (i.e., a random 4-plane containing c, p, c_1, c_3), $a_2 \sim A(c, \bar{a}_2)$ and outputting (a_1, a_2, a_3) . In the previous calculation we have ignored error terms of size $\mathcal{O}(|\mathbb{F}|^{-1})$. Thus the marginal distribution on (a_1, a_3) is $\mathcal{O}(|\mathbb{F}|^{-1})$ -close to uniform on $A(c)$, and the result follows. \square

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A Sufficient Conditions for Non-Malleability in LTCs

In this section we prove Claim 1, restated below.

Claim 1 (Restated). *Let $\varepsilon > 0$ and $\ell \in \mathbb{N}$ be parameters. Let $(\text{Enc}, \text{Dec}, \text{Test})$ be a LTC with $\text{Enc} : \Gamma^k \rightarrow \Gamma^n$, and let*

$$\mathcal{F} \subset \{f : \Gamma^n \rightarrow \Gamma^n\} \text{ and } \mathcal{G} \subset \{g : \Gamma^n \rightarrow (\Gamma \cup \{\perp\})^n\}$$

be function families. Suppose the following four conditions hold.

1. \mathcal{G} contains the constant “all \perp ” function; for all other $g \in \mathcal{G}$, and all valid codewords $\mathbf{x} \in \Gamma^n$, $\text{Test}(g(\mathbf{x}); I) = 1$ occurs with probability 1;
2. For all distinct $g, g' \in \mathcal{G}$ and $m \in \Gamma^k$, $\Pr_{\mathbf{x} \sim \text{Enc}(m), I} [g(\mathbf{x})_I = g'(\mathbf{x})_I] \leq \varepsilon/\ell^2$.
3. For all $f \in \mathcal{F}$, and $m, m' \in \Gamma^k$: $\Delta\left(\{f(\mathbf{x})_I\}_{\mathbf{x} \sim \text{Enc}(m), I}, \{f(\mathbf{x})_I\}_{\mathbf{x} \sim \text{Enc}(m'), I}\right) \leq \varepsilon$.

4. For all $f \in \mathcal{F}$ there exists a list $L_f = \{g^{(1)}, \dots, g^{(\ell)}\} \subset \mathcal{G}$ of size $|L_f| = \ell$ such that for all $m \in \Gamma^k$,

$$\Pr_{\mathbf{x} \sim \text{Enc}(m), I} \left[\text{Test}(f(\mathbf{x}); I) = 1 \ \& \ f(\mathbf{x})_I \notin \{g^{(j)}(\mathbf{x})_I : g^{(j)} \in L_f\} \right] \leq \varepsilon.$$

Then $(\text{Enc}, \text{Dec}, \text{Test})$ is an $(\ell + 1, 3\varepsilon)$ -locally testable, non-malleable reduction from \mathcal{F} to \mathcal{G} .

Proof. Define $g : \Gamma^n \rightarrow (\Gamma \cup \{\perp\})^n$ by setting $g(\mathbf{x})_i = f(\mathbf{x})_i$ if there is a unique $g^{(j)} \in L_f$ such that $f(\mathbf{x})_i = g^{(j)}(\mathbf{x})_i$, and $g(\mathbf{x})_i = \perp$ otherwise. Note every coordinate of g is a convex combination of the corresponding coordinate functions in $L_f \cup \{\perp\}$, a subset of \mathcal{G} of size $\ell + 1$. We must exhibit a distribution SIM on Γ^n such that for all $m \in \Gamma^k$, $\Delta(\mathcal{D}_f(m), \mathcal{D}_g^{\text{SIM}}) \leq 3\varepsilon$, where $\mathcal{D}_f(m)$ (resp. $\mathcal{D}_g^{\text{SIM}}$) are the distributions which 1) draw I according to Test , and $\mathbf{x} \sim \text{Enc}(m)$ (resp. $\mathbf{x} \sim \text{SIM}$); 2) tamper to obtain $\tilde{\mathbf{x}} = f(\mathbf{x})$ (resp. $\tilde{\mathbf{x}} = g(\mathbf{x})$); 3) output $\tilde{\mathbf{x}}_I$ if $\text{Test}(\mathbf{x}; I) = 1$, \perp otherwise. Let SIM be the distribution which draws $m \sim \Gamma^k$ and outputs $\mathbf{x} \sim \text{Enc}(m)$. The third condition implies that for all $m \in \Gamma^k$, $\Delta(\mathcal{D}_f(m), \mathcal{D}_f^{\text{SIM}}) \leq \varepsilon$. Moreover, for all $m \in \Gamma^k$,

$$\begin{aligned} \Delta(\mathcal{D}_f(m), \mathcal{D}_g(m)) &\leq \Pr_{\mathbf{x} \sim \text{Enc}(m), I} \left[\text{Test}(f(\mathbf{x}); I) = 1 \ \& \ f(\mathbf{x})_I \neq g(\mathbf{x})_I \right] \\ &\leq \Pr_{\mathbf{x} \sim \text{Enc}(m), I} \left[\exists g^{(j)}, g^{(j')} \in L_f \text{ st } g^{(j)}(\mathbf{x})_I = g^{(j')}(\mathbf{x})_I \right] + \\ &\quad + \Pr_{\mathbf{x} \sim \text{Enc}(m), I} \left[\text{Test}(f(\mathbf{x}); I) = 1 \ \& \ f(\mathbf{x})_I \notin \{g^{(j)}(\mathbf{x})_I : g^{(j)} \in L_f\} \right] \\ &\leq \binom{\ell}{2} \cdot \frac{\varepsilon}{\ell^2} + \varepsilon \leq 2\varepsilon. \end{aligned}$$

We have used the second and fourth conditions. It follows that $\Delta(\mathcal{D}_f^{\text{SIM}}, \mathcal{D}_g^{\text{SIM}}) \leq 2\varepsilon$, and the result follows. \square

B Sampler Replacement

In the body we used the following fact with $(\varepsilon', \delta') = (\varepsilon, \delta)$ and $\rho = \zeta = \varepsilon$.

Fact 2 (Restated). Let $\varepsilon, \delta, \varepsilon', \delta', \varepsilon^*, \delta^*, \rho, \zeta > 0$ be such that $\delta^*(\varepsilon^* - \varepsilon - \varepsilon' - 2\rho - \zeta) \geq \delta'/\zeta + \delta/\rho$. Suppose $A/B/C$ is such that:

- A/C , B/C and $B(a)/C(a)$ are 0-biregular for all $a \in A$; and
- A/C is (ε, δ) -sampling and $A(c)/B(c)$ is (ε', δ') -sampling for all $c \in C$.

Then A/B is $(\varepsilon^*, \delta^*)$ -sampling.

Proof. Fix $\varepsilon, \delta, \varepsilon', \delta', \varepsilon^*, \rho, \zeta > 0$ and $A/B/C$ as in the statement. Let $B' \subset B$ be a set of size $|B'| = \lambda \cdot |B|$, and let $A' \subset A$ be the set of $a \in A$ such that $|\Pr_{b \sim B(a)}(b \in B') - \lambda| > \varepsilon^*$, let $\nu = |A'|/|A|$. We must show that $\nu \leq (\delta'/\zeta + \delta/\rho)/(\varepsilon^* - \varepsilon - \varepsilon' - 2\rho - \zeta)$. We have

$$\varepsilon^* < \mathbb{E}_{a \sim A'} \left[\left| \Pr_{b \sim B(a)}(b \in B') - \lambda \right| \right] \leq \mathbb{E}_{a \sim A'} \left[\left| \mathbb{E}_{c \sim C(a)} \left[\Pr_{b \sim B(a,c)}(b \in B') \right] - \lambda \right| \right]$$

$$\leq \mathbb{E}_{\substack{a \sim A' \\ c \sim C(a)}} \left[\left| \Pr_{b \sim B(a,c)}(b \in B') - \lambda(c) \right| \right] + \mathbb{E}_{a \sim A'} \left[\left| \mathbb{E}_{c \sim C(a)}[\lambda(c)] - \mathbb{E}_{c \sim C}[\lambda(c)] \right| \right],$$

where for $c \in C$, $\lambda(c) := \Pr_{b \sim B(c)}(b \in B')$. We have used the biregularity of $B(a)/C(a)$ for all $a \in A$ and that $\mathbb{E}_{c \sim C}[\lambda(c)] = \lambda$, which follows from biregularity of B/C . Let RHS_1 and RHS_2 be the two expectations on the right hand side of the equation above. We bound RHS_1 and RHS_2 separately. Note,

$$\text{RHS}_2 \leq \varepsilon + 2\rho + \nu^{-1} \cdot \Pr_{a \sim A} \left[\left| \mathbb{E}_{c \sim C(a)}[\lambda(c)] - \mathbb{E}_{c \sim C}[\lambda(c)] \right| > \varepsilon + 2\rho \right] \leq \varepsilon + 2\rho + \nu^{-1} \cdot \delta/\rho.$$

Thus, it suffices to show that $\text{RHS}_1 \leq \zeta + \varepsilon' + \nu^{-1} \cdot \delta'/\zeta$. Let $C' \subset C$ be the set of $c \in C$ such that $\Pr_{\substack{a \sim A' \\ c' \sim C(a)}}(c' = c) < \zeta/|C|$. Clearly, $\Pr_{\substack{a \sim A' \\ c \sim C(a)}}(c \in C') < \zeta$. Also, whenever $c \notin C'$, we have

$$\nu \cdot \zeta \leq \nu \cdot |C| \cdot \Pr_{\substack{a \sim A \\ c' \sim C(a)}}[c' = c | a \in A'] = \Pr_{\substack{c' \sim C \\ a \sim A(c')}}[a \in A' | c' = c] = \Pr_{a \sim A(c)}[a \in A'].$$

We have used the biregularity of A/C . This gives

$$\begin{aligned} \text{RHS}_1 &< \zeta + \varepsilon' + \max_{c \notin C'} \left\{ \Pr_{a \sim A(c)} \left[\left| \Pr_{b \sim B(a,c)}(b \in B') - \lambda(c) \right| > \varepsilon' \right] / \Pr_{a \sim A(c)}(a \in A') \right\} \\ &\leq \zeta + \varepsilon' + \nu^{-1} \cdot \delta'/\zeta, \end{aligned}$$

and the result follows. □